

MATH4210: Financial Mathematics Tutorial 4

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Review on Normal r.v.

Question

Assume a sequence of i.i.d. r.v.s $\{X_i\}_{i=1\dots n}$, $X_1 \sim N(\mu, \sigma^2)$. Denote by $Y := \frac{1}{n} \sum_{i=1}^n X_i$. Find $\mathbb{E}[Y]$ and $\text{Var}(Y)$

$$\mathbb{E}[Y] = \mathbb{E}\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n} \sum_{i=1}^n \mathbb{E}[X_i] = \frac{1}{n} \sum_{i=1}^n \mu = \mu$$

$$\begin{aligned}\text{var}(Y) &= \text{var}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \text{var}\left(\sum_{i=1}^n X_i\right) + \underbrace{\sum_{i \neq j} \frac{1}{n^2} \text{cov}(X_i, X_j)}_0 = \\ &= \frac{1}{n^2} n \cdot \sigma^2 = \frac{\sigma^2}{n}\end{aligned}$$

Review on Normal r.v.

If $Y \sim N(\mu_Y, \sigma_Y^2) \Leftrightarrow E[e^{i\theta Y}] = e^{i\theta \mu_Y - \frac{1}{2}\theta^2 \sigma_Y^2}$

Fix $\theta \in \mathbb{R}$

$$E[e^{i\theta Y}] = E[e^{i\theta \sum_{j=1}^n X_j}] = E\left[\prod_{j=1}^n e^{i\theta X_j}\right]$$

Question

Assume a sequence of independent r.v.s $\{X_i\}_{i=1\dots n}$, for any $i \in [|1, n|]$, $X_i \sim N(\mu_i, \sigma_i^2)$. Denote by $Y := \sum_{i=1}^n X_i$. Show that Y is gaussian. Find the parameters of Y .

Since X_j 's are independent .

$$\begin{aligned} E[e^{i\theta Y}] &= \prod_{j=1}^n E[e^{i\theta X_j}] = \prod_{j=1}^n e^{i\theta \mu_j - \frac{1}{2}\theta^2 \sigma_j^2} \\ &= e^{i\theta \sum_{j=1}^n \mu_j - \frac{1}{2}\theta^2 \sum_{j=1}^n \sigma_j^2} \end{aligned}$$

$$Y \sim N\left(\sum_{j=1}^n \mu_j, \sum_{j=1}^n \sigma_j^2\right)$$

Review on Normal r.v.

$$f_X(x) = \frac{d}{dx} P_X(X \leq x) = \frac{d}{dx} P_X(\ln X \leq \ln x) \sim N(\mu, \sigma^2)$$

Φ : cdf of $N(0,1)$

$$= \frac{d}{dx} \Phi\left(\frac{\ln x - \mu}{\sigma}\right)$$

φ : pdf of $N(0,1)$

$$= \varphi\left(\frac{\ln x - \mu}{\sigma}\right) \frac{d}{dx} \left(\frac{\ln x - \mu}{\sigma}\right)$$

Question

Assume X follows log-normal distribution with parameters μ, σ^2 . Find the probability density function of X .

Recall that $X \sim LN(\mu, \sigma^2)$ if

$$f_X(x) = \frac{1}{x\sqrt{2\pi}} e^{-\frac{x-\mu}{\sigma^2}}$$

$$\ln(X) \sim N(\mu, \sigma^2)$$

$$= \varphi\left(\frac{\ln x - \mu}{\sigma}\right) \cdot \frac{1}{\sigma x}$$

$$= \frac{1}{x\sqrt{2\pi}} \cdot e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}}$$

Review on Normal r.v.

fix $n \in \mathbb{N}^*$,

$$E[X^n] = \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} x^n e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} x^{n-1} \cdot x \cdot e^{-\frac{x^2}{2}} dx$$

$$v(x) = e^{-\frac{x^2}{2}}$$

$$u(x) = x^{n-1} = \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} -u(x) dV(x)$$

Question

Assume $X \sim N(0, 1)$. For $n \in \mathbb{N}^*$, find $E[X^n]$.

If n is odd, $E[X^n] = 0$

If n is even, $E[X^n] = (n-1)(n-3)\dots 1$

Remark

If $Y \sim N(\mu, \sigma^2)$, we can write $E[Y^n]$ explicitly by considering

IDP

$$Y = \mu + \sigma Z, \quad Z \sim N(0, 1).$$

$$E[X^n] = \begin{cases} 0 & \text{if } n=2k+1 \\ \frac{(2k)!}{2^k \cdot k!} & \text{if } n=2k \end{cases}$$

$$\begin{aligned} E[X^n] &= \frac{1}{\sqrt{2\pi}} \left(-x^n e^{-\frac{x^2}{2}} \Big|_{-\infty}^{\infty} \right) - \left(\int_{\mathbb{R}} -e^{-\frac{x^2}{2}} (n-1) \cdot x^{n-2} dx \right) \\ &= \frac{1}{\sqrt{2\pi}} \left(0 + (n-1) \int_{\mathbb{R}} x^{n-2} e^{-\frac{x^2}{2}} dx \right) \Bigg| \begin{array}{l} E[X] = 0 \\ E[X^2] = 1 \end{array} \\ &= (n-1) E[X^{n-2}] \end{aligned}$$

Convergence of r.v.s

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space. X and $\{X_n\}$ are \mathbb{R} valued (sequence of) r.v.s.

Definition (Convergence almost surely)

Denote by $X_n \rightarrow X$ a.s. (almost surely) if

$$\mathbb{P}[\{\omega \in \Omega : \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)\}] = 1$$

Definition (Convergence in Probability)

Denote by $X_n \rightarrow X$ in probability if for any $\rho > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}[\{\omega \in \Omega : |X_n(\omega) - X(\omega)| \geq \rho\}] = 0$$

Convergence of r.v.s

since $X_n \rightarrow X$ a.s. $\Rightarrow P\{w \in \Omega : \lim_{n \rightarrow \infty} X_n(w) = X(w)\} = 1$

① fix $p > 0$, and $n \in \mathbb{N}$

$$A_n = \bigcup_{m \geq n} \{ |X_m - X| \geq p \} \quad \Rightarrow \quad P(\{ |X_n - X| \geq p \}) \leq P(A_n) \quad (\star)$$

② A_n is an decreasing sequence

$$P(A_1) \leq P(\Omega) = 1$$

$$P\left(\bigcap_{n=1}^{\infty} A_n\right) = \lim_{n \rightarrow \infty} P(A_n) \quad (\star\star)$$

$$\{ |X_n - X| \geq p \} \subseteq A_n$$

Question

Let X and $\{X_n\}$ be \mathbb{R} valued (sequence of) r.v.s. Assume $X_n \rightarrow X$ a.s.. Show that $X_n \rightarrow X$ in probability.

③ consider $B = \{ \lim_{n \rightarrow \infty} |X_n - X| = 0 \}$

$$P(B) = 1$$

Fix $w \in B$ since $X_n \rightarrow X$ a.s.

There exist $N > 0$

$$|X_n(w) - X(w)| < \varepsilon \text{ for all } n \geq N$$

But $\forall n \geq N \quad w \notin A_n \Rightarrow B \cap \left(\bigcap_{n=1}^{\infty} A_n\right) = \emptyset \Rightarrow X_n \rightarrow X \text{ in probability}$

$$P(B \cup \left(\bigcap_{n=1}^{\infty} A_n\right)) \leq P(\Omega) = 1$$

$$= P(B) + P\left(\bigcap_{n=1}^{\infty} A_n\right) \leq 1$$

$$\Rightarrow P\left(\bigcap_{n=1}^{\infty} A_n\right) = 0 \quad (\star\star)$$