Recall

Independence of events

- E and F are independent if $P(EF) = P(E)P(F) \iff P(E|F) = P(E)$ if P(F) > 0.
- $\{E_1, \ldots, E_n\}$ are independent if for **every** $\{i_1, \ldots, i_r\} \subset \{1, \ldots, n\}$,

$$P(E_{i_1}\cdots E_{i_r}) = P(E_{i_1})\cdots P(E_{i_r}).$$

• An infinite family of events are *independent* if every finite subset of events from that family are independent.

Example 1 (Pairwise independence \implies independence). Roll a die twice. Consider the events

$$A = \{ \text{ sum of the two numbers is } 7 \},$$

 $B = \{ \text{ the first number is } 3 \},$

 $C = \{ \text{ the second number is 4 } \}.$

Then

$$P(A) = P(B) = P(C) = \frac{1}{6}$$
 and $P(A \cap B) = P(A \cap C) = P(B \cap C) = \frac{1}{36}$.

However,

$$P(A \cap B \cap C) = \frac{1}{36}$$
 while $P(A)P(B)P(C) = \frac{1}{216}$.

Hence the events $\{A, B, C\}$ are pairwise independent but NOT independent.

Remark. It follows from Example 1 that we should check all the 'product' equations in the definition to assure the independence of a finite family of events.

Discrete random variables

A random variable X is a function from the sample space S to the real numbers \mathbb{R} . The randomness comes from the probability $P(\cdot)$ on the sample space. If the range of X is a countable set in \mathbb{R} , then X is called *discrete* random variable. The *probability mass function* (PMF) of a discrete random variable is defined as

$$p(x) \coloneqq P(X = x).$$

Expectation / expected value / mean.

$$E[X] := \sum_{x: p(x) > 0} xp(x).$$

Let $g: \mathbb{R} \to \mathbb{R}$. Then it is proved that (law of the unconscious statistician)

$$E[g(X)] = \sum_{x: p(x)>0} g(x)p(x).$$

In particular, for any $a, b \in \mathbb{R}$,

$$E[aX + b] = aE[X] + b.$$

Variance.

$$Var(X) := E[(X - E[X])^2] = E[X^2] - (E[X])^2,$$

$$Var(aX + b) = a^2 Var(X).$$

Computing expectations

And for $a, b \in \mathbb{R}$,

Example 2 (discrete layer-cake). Let X be a random variable with non-negative integral values, i.e., $X: S \to \mathbb{Z}_{>0}$ where S denotes the sample space. Then

$$E[X] = \sum_{k=1}^{\infty} P(X \ge k).$$

Proof. By countable additivity and change of the order of summation,

$$\sum_{k=1}^{\infty} P(X \ge k) = \sum_{k=1}^{\infty} \sum_{i=k}^{\infty} P(X = i)$$

$$= \sum_{k=1}^{\infty} \sum_{i=1}^{\infty} \chi_{\{(x,y): x \ge y\}}(i,k) P(X = i) = \sum_{i=1}^{\infty} \sum_{k=1}^{\infty} \chi_{\{(x,y): y \le x\}}(i,k) P(X = i)$$

$$= \sum_{i=1}^{\infty} \sum_{k=1}^{i} P(X = i)$$

$$= \sum_{i=1}^{\infty} i P(X = i) = E[X],$$

where the blue line contains the formal details about the change of the order of summation. Recall the indicator function for a set $E \subset \mathbb{R}^2$ is defined as

$$\chi_E(i,k) := \begin{cases} 1 & \text{if } (i,k) \in E, \\ 0 & \text{if } (i,k) \notin E. \end{cases}$$

Remark. The blue line is not only a safe trick for us to change the order of summation or integration, but also a simple example showing the usefulness of indicator functions.

Example 3. Randomly choose 3 numbers from $\{1, \ldots, 10\}$. Let X be the smallest number among the 3 chosen numbers. Find E[X].

Solution. First determine the probability mass function

$$p(k) = P(X = k) = \begin{cases} \frac{\binom{10-k}{2}}{\binom{10}{3}} & k = 1, \dots, 8; \\ 0 & k = 9, 10, \end{cases}$$

where the first case is obtained by choosing the other 2 numbers that are greater than k, and the second case results from the observation that 9, 10 can never be the smallest numbers among the 3 chosen numbers. Then

$$E[X] = \sum_{k=1}^{10} kp(k) = \sum_{k=1}^{8} k \frac{\binom{10-k}{2}}{\binom{10}{3}} = \frac{11}{4}.$$

Example 4. In a game show, there are 3 doors:

$$[+\$30], [-\$10], [-\$10]$$

To start the game, you have to pay \$5 to randomly open a door from these 3 doors which look the same to you. You will get \$30 if opening $\boxed{+\$ 30}$ and lose \$10 if opening $\boxed{-\$ 10}$.

Assume you are 'reasonable': if you win \$30, then you quit the game; if you lose \$10, then you flip a coin to decide whether you should continue, i.e., if the coin shows tail then you quit the game; if the coin shows head then you pay another \$5 to randomly open a door from the remaining 2 doors. Assume you are only allowed to play at most 2 rounds.

Let X be the winnings when you leave. Find E[X] and Var(X).

Solution. First determine the sample space of the door results each of which is represented by a vector

$$S = \{(+30), (-10), (-10, +30), (-10, -10)\}.$$

Let G be the random variable labeling the above outcomes with 1, 2, 3, 4 from left to right, e.g. the event $\{G=2\}$ is the outcome (-10). Then the r.v. X is a function of the r.v. G, i.e. X=f(G) for some function $f:\mathbb{R}\to\mathbb{R}$ which assigning the winnings to outcomes. Hence

$$\begin{split} E[X] &= E[f(G)] \\ &= \sum_{k=1}^{4} P(G=k)f(k) \\ &= \frac{1}{3} \times (-5+30) + \frac{2}{3} \times \frac{1}{2} \times (-5-10) + \\ &\quad + \frac{2}{3} \times \frac{1}{2} \times \frac{1}{2} \times (-5-10-5+30) + \frac{2}{3} \times \frac{1}{2} \times \frac{1}{2} \times (-5-10-5-10) \\ &= 0, \end{split}$$

And since E[X] = 0,

$$Var(X) = E[X^{2}] - (E[X])^{2} = E[(f(G))^{2}] = \sum_{k=1}^{4} P(G = k) (f(k))^{2} = 450.$$