Review

. The Central limit Thm

Let X1, ..., Xn, ..., be an i.i.d. sequence of r.u.'s, each having finite mean  $\mu$  and vaniance  $\sigma^2$ .

Then Yack,

$$P\left\{ \frac{X_1 + \dots + X_n - n \mu}{\sqrt{n} \sigma} \leq \alpha \right\} \longrightarrow \Phi(\alpha) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{X^2}{2}} dx,$$

as  $n \to \infty$ .

Example 1 If 10 fair dice are rolled, find the approximate prob. that the sum obtained is between 30 and 40.

Solution: Let Xi be the value obtained in the i-th roll, i=1, 2, ..., (0.

We need to calculate

Notice that 
$$\mu = E[X_i] = \frac{1}{6}(1+2+\cdots+6)$$

$$= \frac{7}{5}$$

$$E[X_{i}^{2}] = \frac{1}{6} (1^{2} + 2^{2} + \dots + 6^{2})$$

$$= \frac{1}{6} \cdot \frac{6 \times 7 \times 13}{6}$$

$$\sigma^{2} = Var(X_{i}) = E[X_{i}^{2}] - E[X_{i}^{2}]^{2}$$

$$= \frac{35}{12}$$

$$= p \left\{ \frac{29.5 - 10x_{2}^{2}}{\sqrt{350/12}} \leqslant \frac{\chi_{1} + \dots + \chi_{10} - 10 \times \frac{7}{2}}{\sqrt{10} \cdot \sqrt{\frac{35}{12}}} \leqslant \frac{40.5 - 10x_{2}^{2}}{\sqrt{350/12}} \right\}$$

$$= 2 \cdot \Phi(1.018) - 1$$

$$\approx 0.69^2$$

Thm2 (The strong law of large numbers).

Let  $X_1, \dots, X_n, \dots$ , be an i.i.d. sequence of r.u.'s with a finite mean  $\mu$ . Then with prob. 1,  $\frac{X_1 + \dots + X_n}{n} \longrightarrow \mu \qquad \text{as } n \to \infty.$ In other word,

$$P\left\{ \lim_{n\to\infty} \frac{\chi_1 + \dots + \chi_n}{n} = \mu \right\} = 1.$$

The proof of Thm 2 is based on two Lemmas.

Lem 3. Assume 
$$X$$
 is a non-negative r.v. which may take the value  $+\infty$ . Suppose  $E[X] < \infty$ .

Then  $P\{X < \infty\} = 1$ .

1 1 X X W J 1.

Pf. For any positive integer 
$$n$$
, by the Markov inequality, 
$$P\{X = \infty\} \leq P\{X \geq n\} \leq \frac{E[X]}{n}.$$

Letting  $n \to \infty$  gives  $P \{ X = \infty \} = 0$ .

Hence 
$$P\{X < \infty\} = 1 - P\{X = \infty\} = 1$$
.

Lem 4. Let 
$$X$$
 be a r.v. Then
$$E[X^4] > (E[X^2])$$

Since 
$$Var(X^2) \ge 0$$
, it follows that  $E[X^4] \ge (E[X^2])^2$ 

Write 
$$S_n = X_1 + \dots + X_n$$
.

WLOG, assume  $\mu = 0$ .

We will estimate
$$E[S_n^4] = E[(X_1 + \cdots + X_n)^4]$$

Notice that
$$Var(X^{2}) = E[X^{4}] - (E[X^{2}])^{2}.$$

$$E[X_{4}] \Rightarrow (E[X_{4}])$$

E[X:] = K < 6.

Expand 
$$(X_1 + \dots + X_n)^4$$
 in terms of  $X_i^4$ ,  $X_i^3 X_j$ ,  $X_i^2 X_j^2$ ,  $X_i^2 X_j^2 X_k$ ,  $X_i X_j X_k X_k$ 

with distinct 
$$\hat{z}, \hat{j}, k, l$$
.

Notice that  $E[X_i^3] = E[X_i^3] = E[X_j] = 0$ .

$$E[X_{i}^{2}X_{j}X_{k}] = E[X_{i}^{2}]E[X_{j}]E[X_{k}]$$

$$= 0.$$

$$=0.$$

$$E[X_iX_jX_kX_\ell]=0.$$

$$E[X_i X_j X_k X_{\ell}] = 0.$$
Hence
$$E[C^4] = E[(X_i + \dots + X_k)^4]$$

Hence
$$E[S_n^4] = E[(X_1 + \dots + X_n)^4]$$

$$= n E[X_i^4]$$

$$+ {\binom{n}{2}} {\binom{4}{2}} E[X_i^2 X_j^2]$$

$$= n E[X_i^4] + 6 {\binom{n}{2}} E[X_i^2] E[X_i^2]$$

$$\leq n \left[ \left[ X_{i}^{4} \right] + 6 {n \choose 2} \right] \left[ \left[ X_{i}^{4} \right] \right]$$

$$= \left( 3n^{2} - 2n \right) \left[ \left[ X_{i}^{4} \right] \right] \left[ \left[ \left[ X_{i}^{4} \right] \right] \right]$$

$$\leq 3n^2 k$$
. where  $k = E[X_i^4] < \infty$ .

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Hence
$$\sum_{n=1}^{\infty} E\left[\left(\frac{S_n}{n}\right)^4\right] \leqslant \sum_{n=1}^{\infty} \frac{3k}{n^2} < \infty.$$

Thus
$$\mathbb{E}\left[\sum_{n=1}^{\infty} \left(\frac{S_n}{n}\right)^4\right] = \sum_{n=1}^{\infty} \mathbb{E}\left[\left(\frac{S_n}{n}\right)^4\right] < \infty$$

Let 
$$X = \sum_{n=1}^{\infty} \left(\frac{S_n}{n}\right)^4$$
 Then X is a r.v, non-negative (may take the value  $\infty$ )

However 
$$E[X] < \infty$$
  
By Lem 3,  $P\{X < \infty\} = 1$ .

Hence 
$$P\left\{\sum_{n=1}^{\infty} \left(\frac{S_n}{n}\right)^4 < \infty\right\} = 1$$
However  $w \in C$ 

$$\sum_{n=1}^{\infty} \left( \frac{S_n}{n} \right)^{\frac{1}{4}} < \infty \implies \frac{S_n}{n} \to 0 \text{ as } n \to \infty.$$

Hence 
$$P \left\{ \lim_{n \to \infty} \frac{S_n}{n} = 0 \right\} = 1$$

Thus With Prob. 1, 
$$\frac{S_n}{n} = \frac{X_i + \dots + X_n}{n} \rightarrow 0.$$

If  $\mu \neq 0$ , then letting  $\widehat{X}_n = X_n - \mu$  applying the SLLN to  $(\widehat{X}_n)$  gives

$$\frac{\widehat{X}_1 + \dots + \widehat{X}_n}{n} \to 0 \qquad \text{almost sure}.$$

$$\iff \frac{\chi_1 + \dots + \chi_n}{n} \to \mu \qquad \text{almost sure}.$$