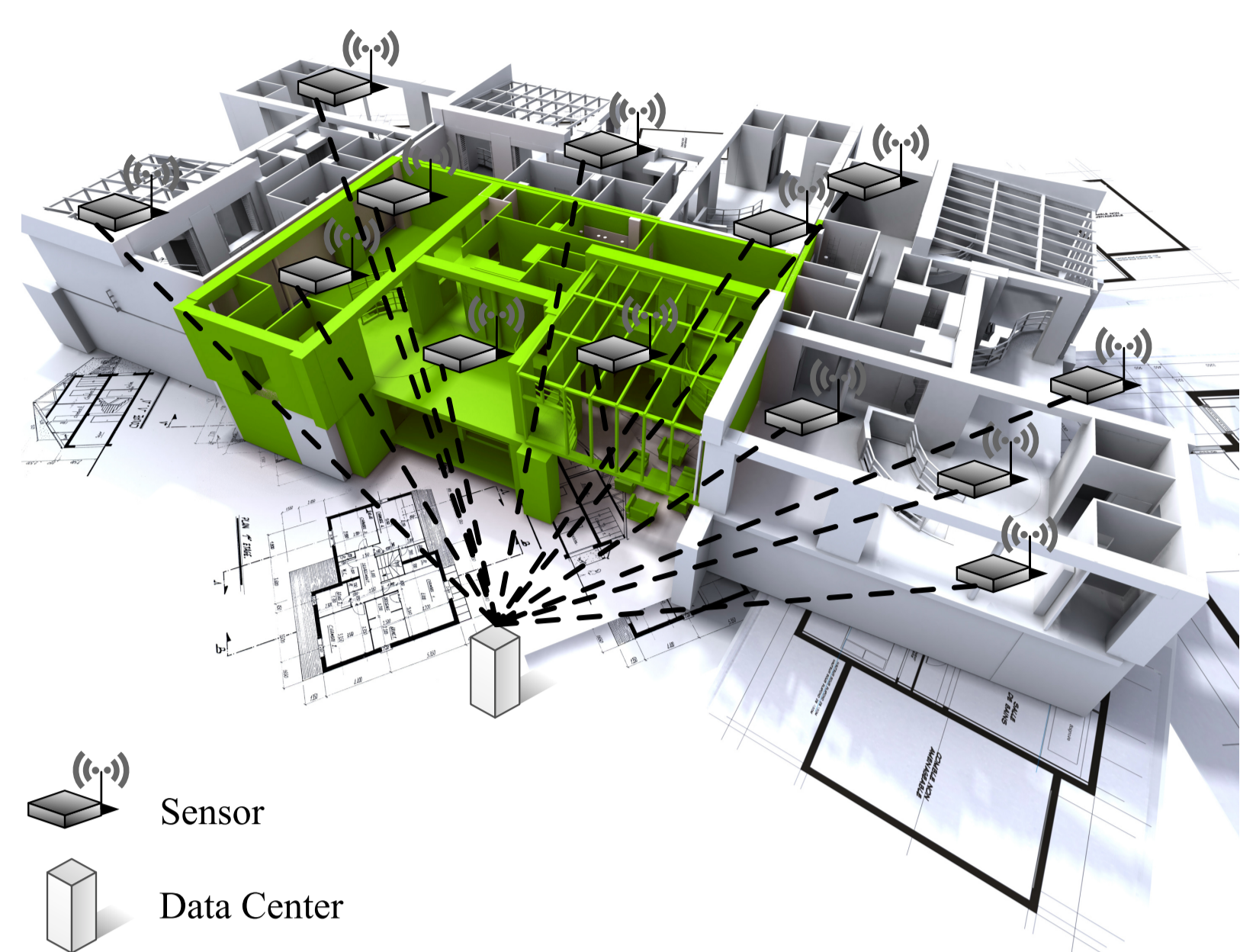


Introduction

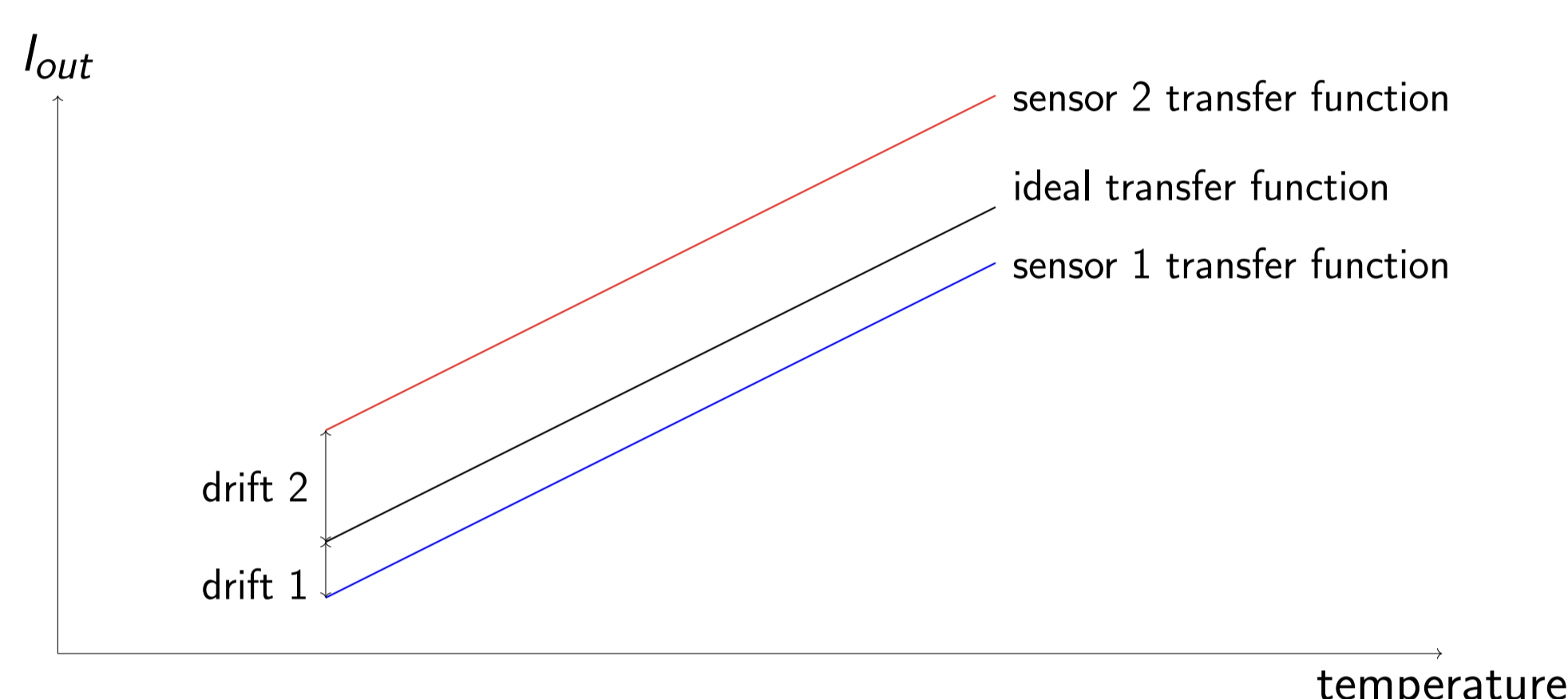


Part Number	Temp. Range	Accuracy	Price
SMT172	-45 ~ 130 °C	±0.25 °C	\$ 35.13
AD590JH	-50 ~ 150 °C	±0.5 °C	\$ 17.91
TMP100	-55 ~ 125 °C	±2.0 °C	\$ 1.79
MCP9509	-40 ~ 125 °C	±4.5 °C	\$ 0.88
LM335A	-40 ~ 100 °C	±5.0 °C	\$ 0.75



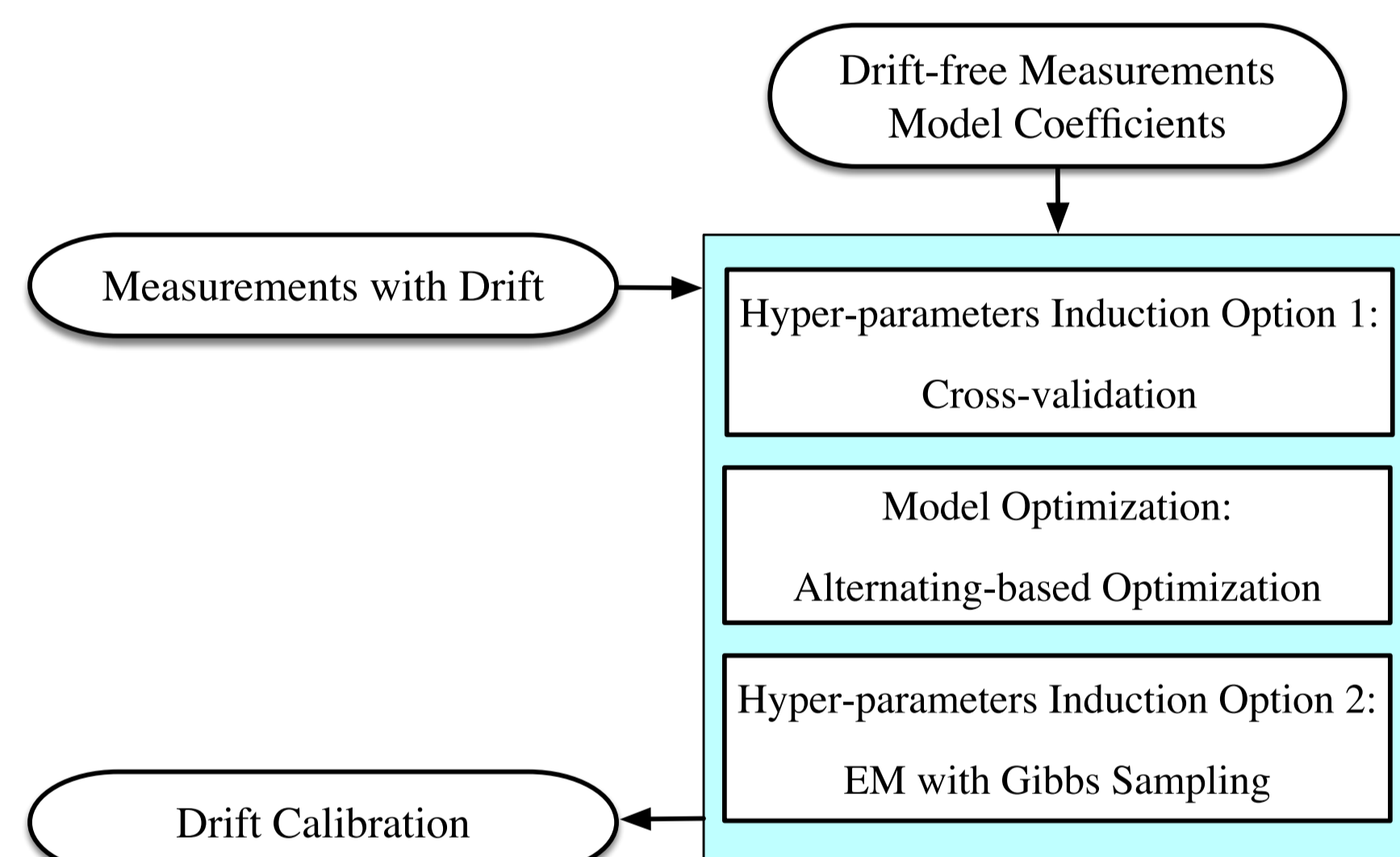
- ▶ Measurement from sensor without calibration is not accurate enough
- ▶ Looking for ways to do calibration

Spatial Correlation Model and Formulation



Sensor Drift Calibration

Given the measurement values sensed by all sensors during several time-instants, drifts will be accurately estimated and calibrated.



Input:

- ▶ $\hat{x}_i^{(k)}$: the measurement value sensed by i th sensor at k th time-instant.
- ▶ $a_{i,j}$: the drift-free model coefficient.

Output:

- ▶ ϵ_i : a time-invariant drift calibration.

Spatial Correlation Model:

- ▶ drift-free model:
 $x_i^{(k)} \approx \sum_{j=1, j \neq i}^m a_{i,j} x_j^{(k)} + a_{i,0}, k = 1, 2, \dots, m_0.$
- ▶ drift-with model:
 $\hat{x}_i^{(k)} + \epsilon_i \approx \sum_{j=1, j \neq i}^m \hat{a}_{i,j} (\hat{x}_j^{(k)} + \epsilon_j) + \hat{a}_{i,0}, k = 1, 2, \dots, m.$

Further Assumption:

- ▶ Likelihood:
 $\mathcal{P}(\hat{\mathbf{x}}|\hat{\mathbf{a}}, \epsilon) \propto \exp(-\frac{\delta_0}{2} \sum_{i=1}^n \sum_{k=1}^m [\hat{x}_i^{(k)} + \epsilon_i - \sum_{j=1, j \neq i}^m \hat{a}_{i,j} (\hat{x}_j^{(k)} + \epsilon_j) - \hat{a}_{i,0}]^2).$
- ▶ Prior distribution of model coefficients (Bayesian Model Fusion):
 $\mathcal{P}(\hat{\mathbf{a}}) \propto \exp(-\sum_{i=1}^n \sum_{j=0, j \neq i}^m \frac{\lambda}{2a_{i,j}^2} (\hat{a}_{i,j} - a_{i,j})^2).$
- ▶ Prior distribution of calibration:
 $\mathcal{P}(\epsilon) \propto \exp(-\frac{\delta_\epsilon}{2} \sum_{i=1}^n \epsilon_i^2).$

Maximum-a-posteriori:

$$\min_{\hat{\mathbf{a}}, \epsilon} \delta_0 \sum_{i=1}^n \sum_{k=1}^m [\hat{x}_i^{(k)} + \epsilon_i - \sum_{j=1, j \neq i}^m \hat{a}_{i,j} (\hat{x}_j^{(k)} + \epsilon_j) - \hat{a}_{i,0}]^2 + \lambda \sum_{i=1}^n \sum_{j=0, j \neq i}^m \frac{1}{2a_{i,j}^2} (\hat{a}_{i,j} - a_{i,j})^2 + \delta_\epsilon \sum_{i=1}^n \epsilon_i^2.$$

Challenges:

- ▶ How to handle this Formulation
- ▶ How to determine hyper-parameters

Model Optimization

- ▶ With fixed one variable, Formulation is a convex unconstrained QP problem w.r.t. another variable;
- ▶ With the first-order optimality condition, sub-Formulation can be transferred as linear equations.

Alternating-based Optimization

Require: Sensor measurements $\hat{\mathbf{x}}$, prior \mathbf{a} and hyper-parameters $\lambda, \delta_0, \delta_\epsilon$.

- 1: Initialize $\hat{\mathbf{a}} \leftarrow \mathbf{a}$ and $\epsilon \leftarrow \mathbf{0}$;
- 2: **repeat**
- 3: **for** $i \leftarrow 1$ to n **do**
- 4: Fix ϵ , solve linear equations (1) using Gaussian elimination to update \hat{a}_i ;
- 5: **end for**
- 6: Fix $\hat{\mathbf{a}}$, solve linear equations (2) using Gaussian elimination to update ϵ ;
- 7: **until** Convergence
- 8: **return** $\hat{\mathbf{a}}$ and ϵ .

$$\delta_0 \sum_{k=1}^m (\hat{x}_i^{(k)} + \epsilon_i) \left[\sum_{j=1}^n \hat{a}_{i,j} (\hat{x}_j^{(k)} + \epsilon_j) + \hat{a}_{i,0} \right] + \lambda \frac{(\hat{a}_{i,t} - a_{i,t})^2}{a_{i,t}^2} = 0, \quad (1)$$

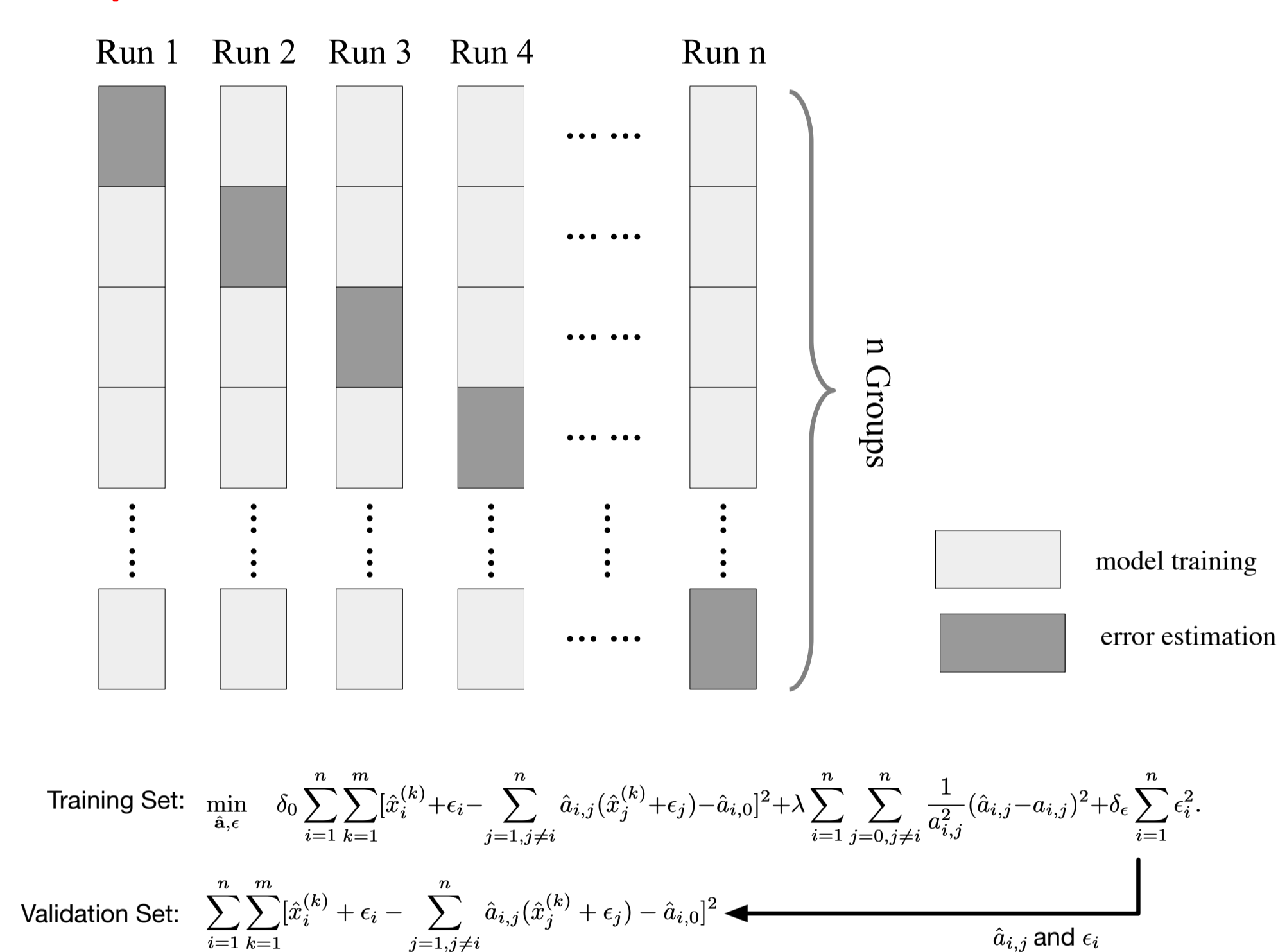
$$\delta_\epsilon \sum_{i=1}^n \sum_{k=1}^m \left[\hat{a}_{i,t} \left(\sum_{j=1}^n \hat{a}_{i,j} (\hat{x}_j^{(k)} + \epsilon_j) + \hat{a}_{i,0} \right) \right] + \delta_\epsilon \epsilon_t = 0, \quad (2)$$

Estimation of Hyper-parameters

Comparison of Estimation for Hyper-parameters

- ▶ **Unsupervised Cross-validation:**
simple, accurate but time-consuming.
- ▶ **Monte Carlo Expectation Maximization:**
fast, flexible but no-accurate.

Unsupervised Cross-validation



Maximum Likelihood Estimation:

$$\max_{\delta_\epsilon, \delta_0, \lambda} \mathcal{P}(\hat{\mathbf{x}}; \delta_0, \lambda, \delta_\epsilon).$$

E-Step

$$Q(\Omega|\Omega^{\text{old}}) = \int \int \mathcal{P}(\hat{\mathbf{a}}, \epsilon|\hat{\mathbf{x}}; \Omega^{\text{old}}) \ln \mathcal{P}(\hat{\mathbf{x}}, \hat{\mathbf{a}}, \epsilon; \Omega) d\hat{\mathbf{a}} d\epsilon$$

$$\approx \frac{1}{L} \sum_{l=1}^L \ln \mathcal{P}(\hat{\mathbf{x}}, \hat{\mathbf{a}}^{(l)}, \epsilon^{(l)}; \Omega)$$

M-Step

$$\max_{\Omega} \frac{1}{L} \sum_{l=1}^L \ln \mathcal{P}(\hat{\mathbf{x}}, \hat{\mathbf{a}}^{(l)}, \epsilon^{(l)}; \Omega).$$

Conditional distribution of each variable

$$\hat{a}_{p,q} \sim \mathcal{P}(\hat{a}_{p,q}|\epsilon, \hat{\mathbf{a}}_{\setminus p,q}, \hat{\mathbf{x}}; \delta_\epsilon, \delta, \lambda) = \mathcal{N}(\mu_{\hat{a}_{p,q}}, \sigma_{\hat{a}_{p,q}}^{-1}),$$

$$\epsilon_t \sim \mathcal{P}(\epsilon_t|\epsilon_{\setminus t}, \hat{\mathbf{a}}, \hat{\mathbf{x}}; \delta_\epsilon, \delta, \lambda) = \mathcal{N}(\mu_{\epsilon_t}, \sigma_{\epsilon_t}^{-1}),$$

Monte Carlo Expectation Maximization

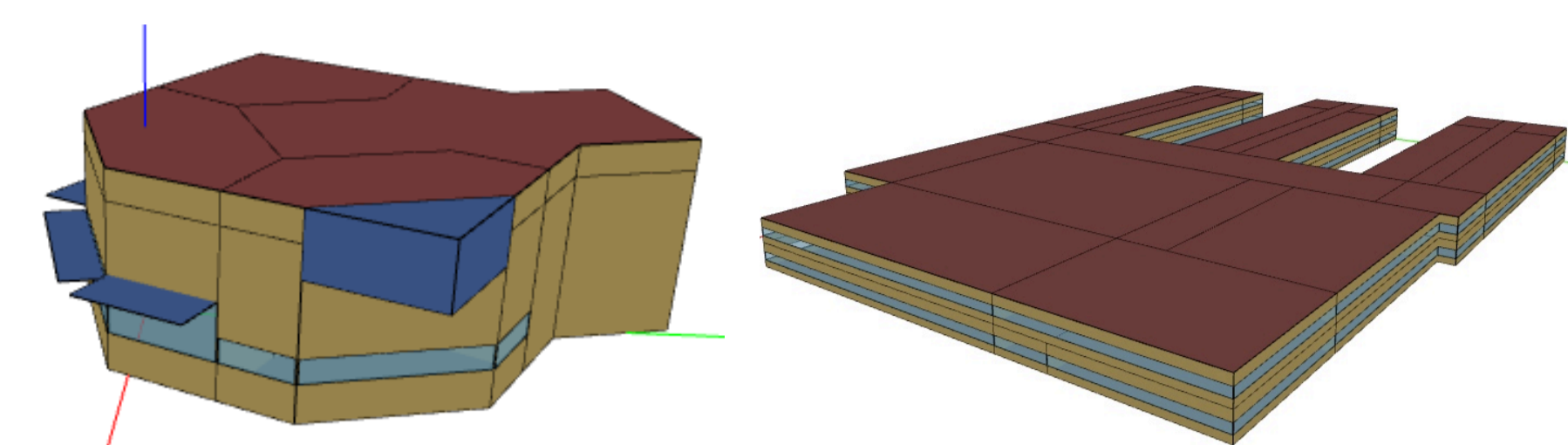
Require: Sensor measurements $\hat{\mathbf{x}}$, prior \mathbf{a} ;

- 1: Initialize hyper-parameters Ω ;
- 2: **repeat**
- 3: Initialize samples $\Psi^{(0)}$ by Alternating-based Optimization;
- 4: **for** $l \leftarrow 1$ to L **do**
- 5: **for** $i \leftarrow 1$ to $n^2 + n$ **do**
- 6: Sample $\psi_i^{(l)}$ from the desired conditional distribution $\mathcal{N}(\mu_{\psi_i}, \sigma_{\psi_i})$ with $\psi_1^{(l)}, \dots, \psi_{i-1}^{(l)}, \psi_{i+1}^{(l-1)}, \dots, \psi_{n^2+n}^{(l-1)}$.
- 7: **end for**
- 8: **end for**
- 9: Update hyper-parameters Ω ;
- 10: **until** Convergence
- 11: **return** hyper-parameters Ω .

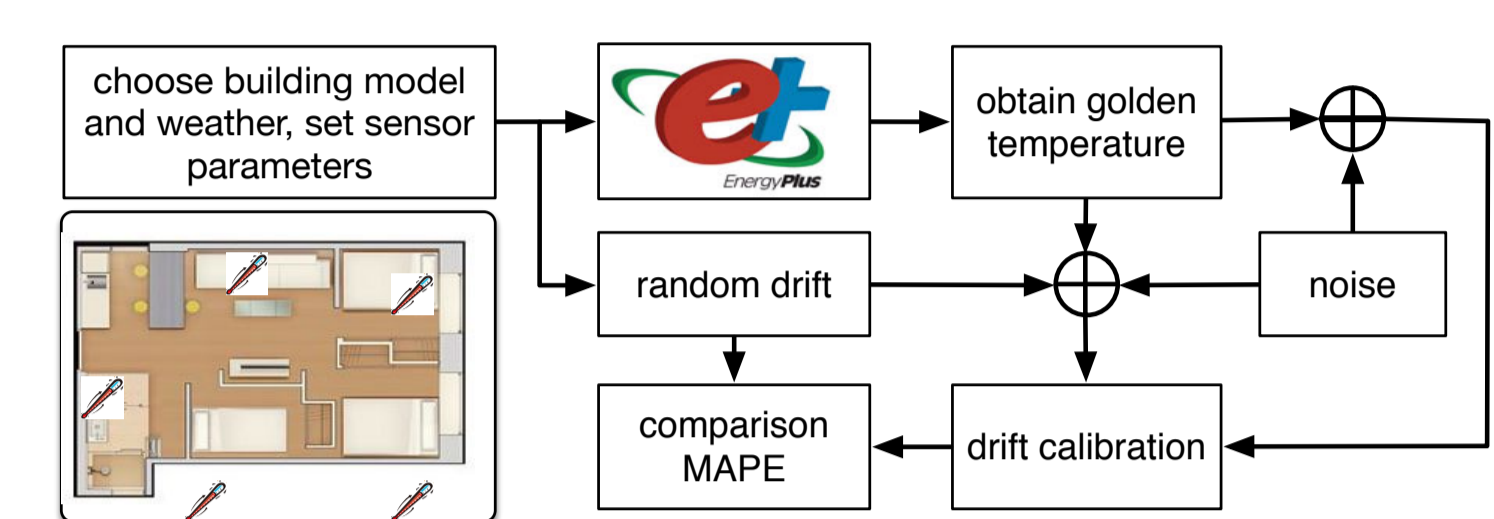
Experimental Results

Benchmark

- ▶ Hall
- ▶ Secondary School

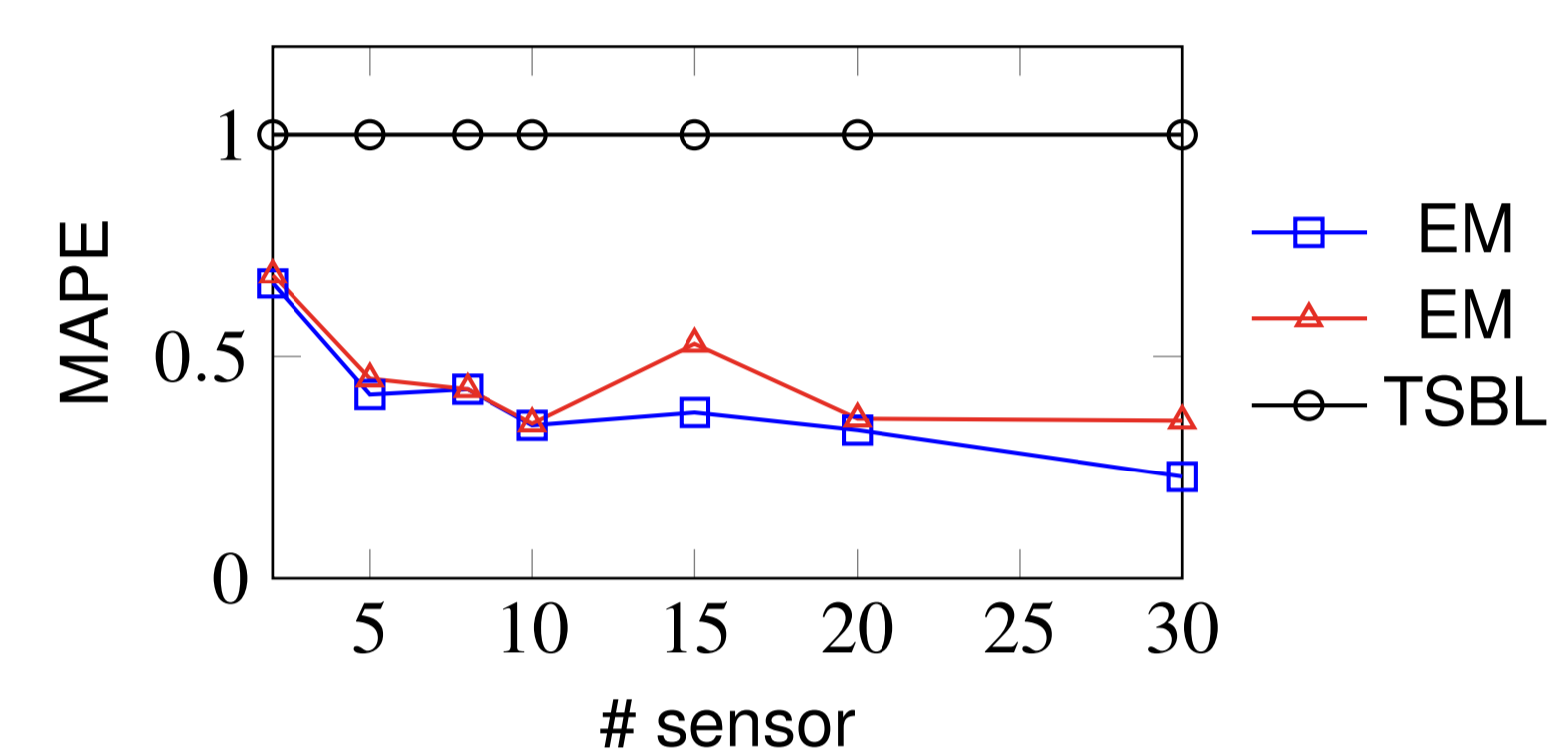


The generated simulation data

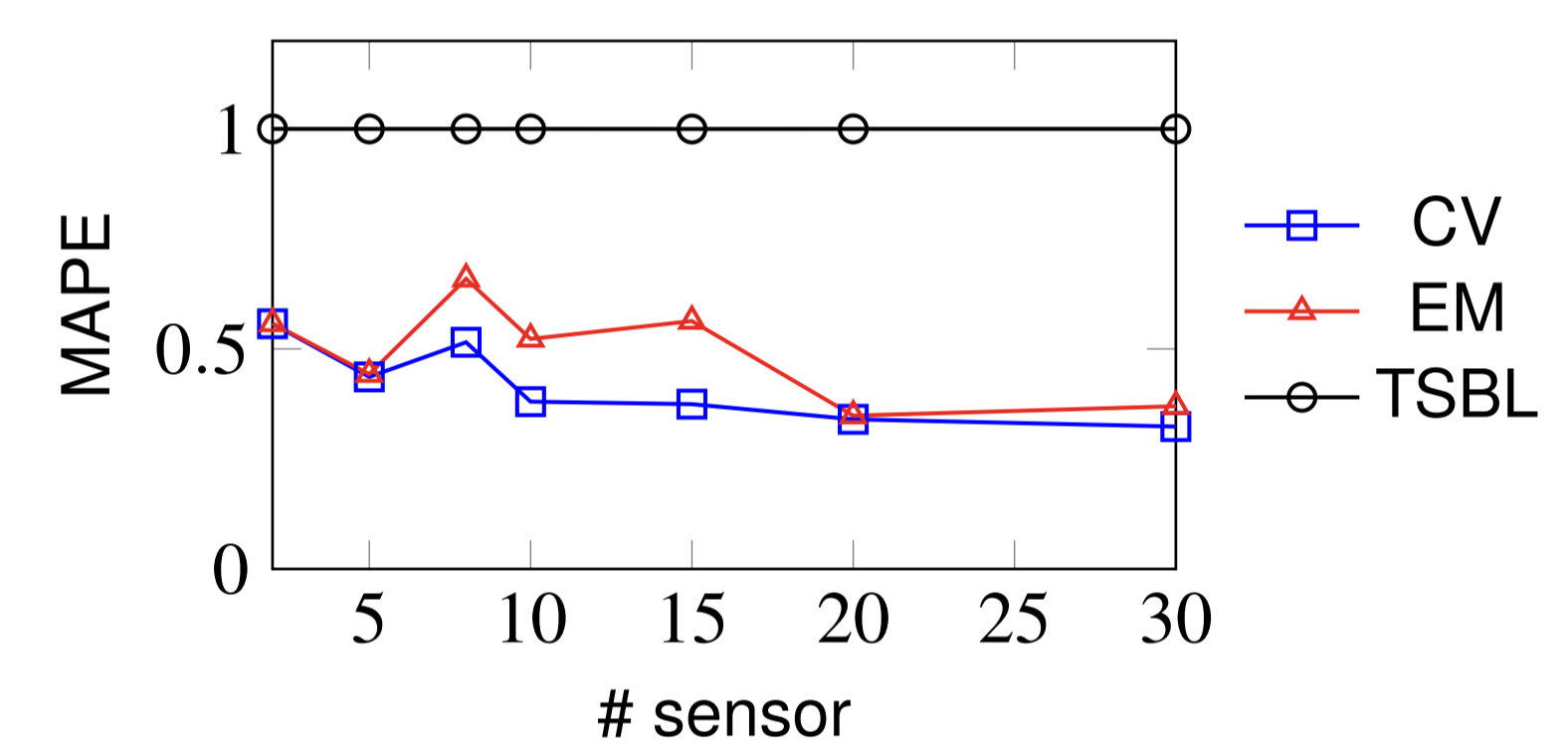


Accuracy

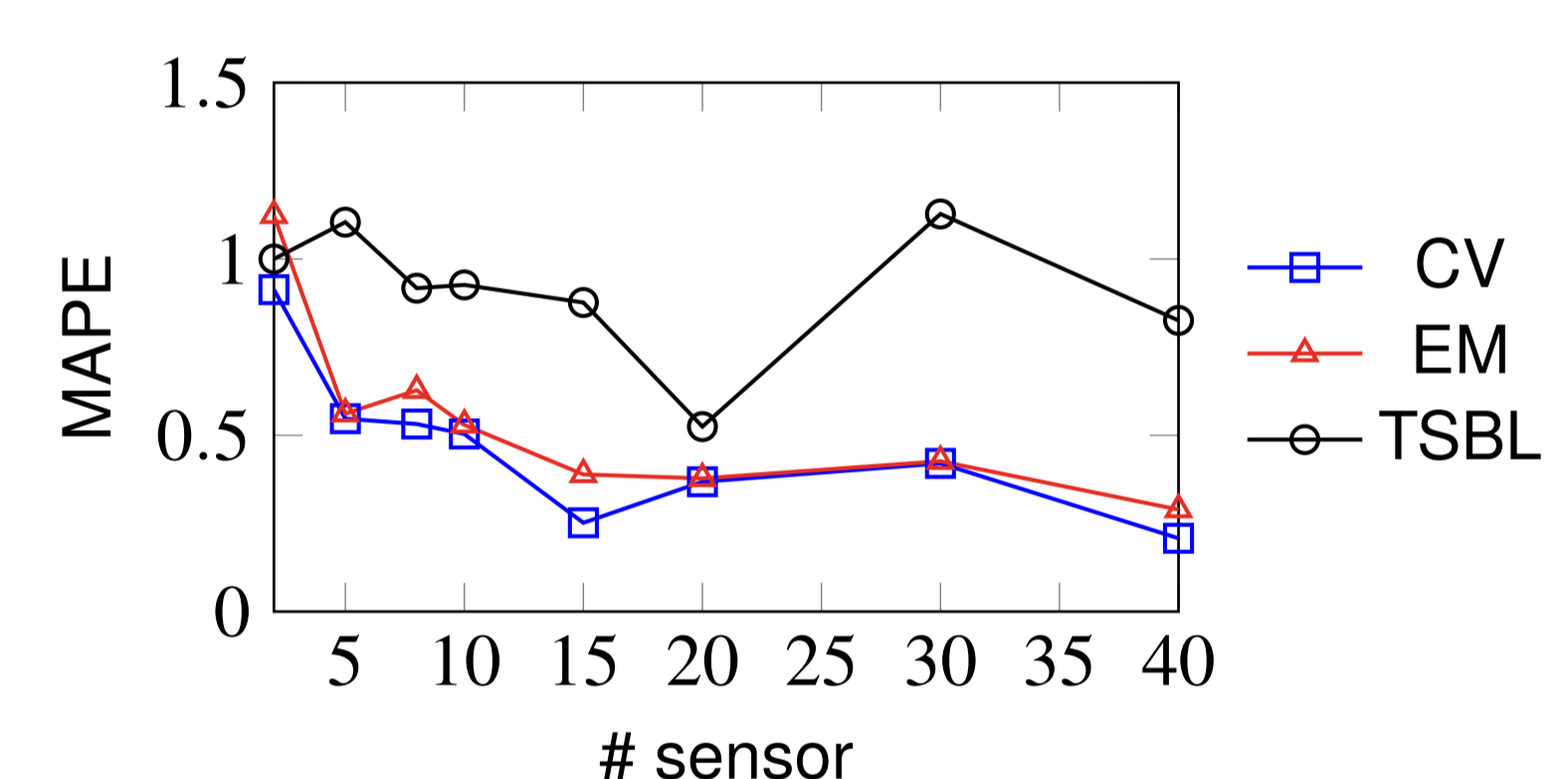
- ▶ Drift variance is set to 2.25; Benchmark: Hall.



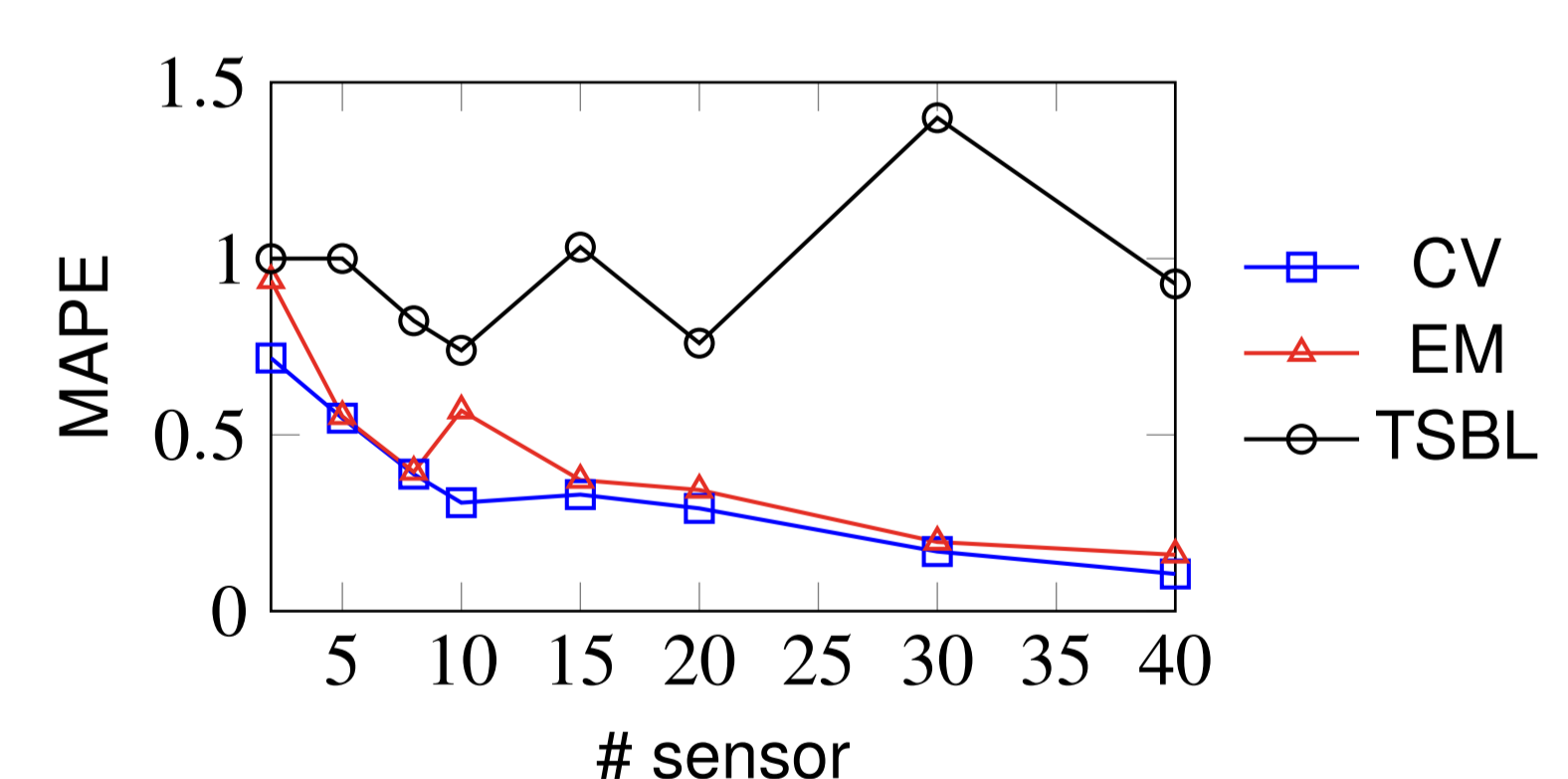
- ▶ Drift variance is set to 2.78; Benchmark: Hall.



- ▶ Drift variance is set to 2.25; Benchmark: Secondary School.

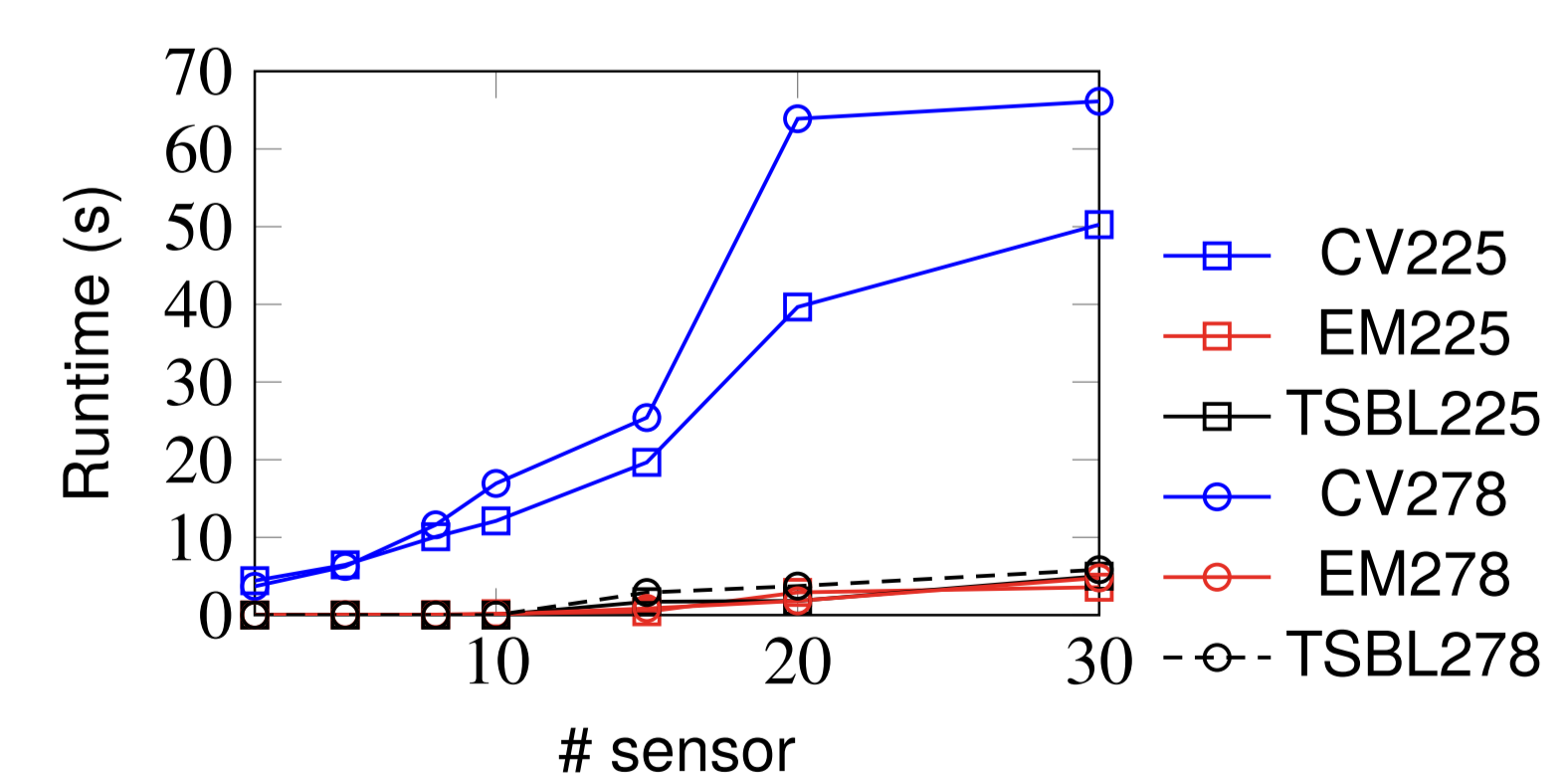


- ▶ Drift variance is set to 2.78; Benchmark: Secondary School.



Runtime

- ▶ Runtime vs. # sensor; Benchmark: Hall.



- ▶ Runtime vs. # sensor; Benchmark: Secondary School.

