THE CHINESE UNIVERSITY OF HONG KONG Department of Mathematics MMAT 5340 Probability and Stochastic Analysis Suggested Solution of Homework 1

1.18 $\binom{7}{3}$ $\binom{7}{3}\binom{8}{3} = 1960$

- 1.19 $\binom{7}{3}$ $\binom{7}{3}\binom{8}{3}(6) = 11760$
- 1.20 $\binom{7}{3}$ $_{3}^{7}\left(\substack{8\\2\end{array}\right) \left(\substack{6\\1\end{math}\right) = 5880$
- 2.45 Note that

$$
P(A) = \frac{3 \times 3 + 3 \times 3}{6^2} = \frac{1}{2}
$$
, $P(B) = \frac{3}{6} = \frac{1}{2}$ and $P(A \cap B) = \frac{3 \times 3}{6^2} = \frac{1}{4}$.

Hence

$$
P(A | B) = \frac{P(A \cap B)}{P(B)} = \frac{1/4}{1/2} = \frac{1}{2},
$$

and

$$
P(B \mid A) = \frac{P(A \cap B)}{P(A)} = \frac{1/4}{1/2} = \frac{1}{2}.
$$

3.23 The distribution of X is given by

$$
P(X = k) = \frac{(1)(10 - k)}{\binom{10}{2}} = \frac{10 - k}{45}, \qquad 1 \le k \le 10.
$$

3.24

$$
\mathbf{E}X = \sum_{k=1}^{10} kP(X=k) = \sum_{k=1}^{10} \frac{k(10-k)}{45} = \frac{11}{3}
$$

.

.

$$
\mathbf{E}X^2 = \sum_{k=1}^{10} k^2 P(X=k) = \sum_{k=1}^{10} \frac{k^2 (10-k)}{45} = \frac{55}{3}.
$$

Hence

$$
Var X = EX^{2} - (EX)^{2} = \frac{55}{3} - \left(\frac{11}{3}\right)^{2} = \frac{44}{9}.
$$

3.25

$$
P(X \le 3) = P(X = 1) + P(X = 2) + P(X = 3) = \frac{8}{15}
$$

4.21 Since each try is independent to another,

 $P(\text{Exactly } 10 \text{ tries}) = P(1 \text{ success in first } 9 \text{ tries}) \cdot P(\text{success in } 10 \text{th try})$ \sim 1

$$
= {9 \choose 1} \left(\frac{2}{6}\right)^1 \left(\frac{4}{6}\right)^8 \cdot \left(\frac{2}{6}\right)
$$

$$
= \frac{256}{6561}.
$$

4.22 Let X and Y be the number of accidents this year from good and bad drivers, respectively. Then $X \sim Bin(5000, 0.0001)$ and $Y \sim Bin(3000, 0.0003)$. Note that $\lambda_1 = N_1 p_1 = 0.5$ and $\lambda_2 = N_2 p_2 = 0.9$. By Poisson approximation, X and Y are approximately distributed as $Poi(\lambda_1)$ and $Poi(\lambda_2)$, respectively. Since X and Y are independent, $X + Y$ is approximately distributed as Poi(λ), where $\lambda = \lambda_1 + \lambda_2 =$ $0.5 + 0.9 = 1.4$. Hence

$$
P(\text{at least two accidents}) = P(X + Y \ge 2)
$$

= 1 - P(X + Y = 0) - P(X + Y = 1)

$$
\approx 1 - \frac{1.4^0 e^{-1.4}}{0!} - \frac{1.4^1 e^{-1.4}}{1!}
$$

= 1 - 2.4e^{-1.4}

$$
\approx 0.4082.
$$

4.23 Let X_1 be the number of trials needed to get the first successful trial, and for $2 \leq i \leq 4$, let X_i be the number of trials needed after the $(i-1)$ -th successful trial to get the next successful trial. Then X_1, X_2, X_3, X_4 are independent with

$$
X_i \sim \text{Geo}(0.4)
$$
, $1 \le i \le 4$ and $X_1 + X_2 + X_3 + X_4 = X$.

Hence

$$
\mathbf{E}X = \mathbf{E}X_1 + \mathbf{E}X_2 + \mathbf{E}X_3 + \mathbf{E}X_4 = \frac{1}{0.4} \times 4 = 10,
$$

and

$$
Var X = Var X_1 + Var X_2 + Var X_3 + Var X_4 = \frac{1 - 0.4}{0.4^2} \times 4 = 15.
$$

5.18

$$
P(X \le 3 \mid X \ge 2) = \frac{P(2 \le X \le 3)}{P(X \ge 2)}
$$

=
$$
\frac{\int_2^3 f(x) dx}{\int_2^\infty f(x) dx}
$$

=
$$
\frac{\int_2^3 2x^{-3} dx}{\int_2^\infty 2x^{-3} dx}
$$

=
$$
\frac{-x^{-2}|_2^3}{-x^{-2}|_2^\infty}
$$

=
$$
\frac{5}{9}.
$$

5.45 Since $X \sim \text{Exp}(1/2)$, X only takes non-negative values. For $0 \le a \le b$,

$$
P(a \le 3X^4 \le b) = P((a/3)^{1/4} \le X \le (b/3)^{1/4}) = \int_{(a/3)^{1/4}}^{(b/3)^{1/4}} \frac{1}{2} e^{-\frac{1}{2}x} dx.
$$

Now by the change of variable $x = (y/3)^{1/4}$, we have

$$
a \le y \le b
$$
 and $dx = (1/4)(y/3)^{-3/4}(1/3) = \frac{1}{12} \left(\frac{y}{3}\right)^{-3/4}.$

Then the integral takes the form

$$
P(a \le 3X^4 \le b) = \int_a^b \frac{1}{2} e^{-\frac{1}{2}(y/3)^{1/4}} \frac{1}{12} \left(\frac{y}{3}\right)^{-3/4} dy = \int_a^b p(y) dy,
$$

where

$$
p(y) := \begin{cases} \frac{1}{24} e^{-\frac{1}{2}(y/3)^{1/4}} \left(\frac{y}{3}\right)^{-3/4} & \text{if } y > 0, \\ 0 & \text{if } y \le 0, \end{cases}
$$

is the density of $3X^4$.

5.46 Since $X \sim \Gamma(3,2)$ and $Y \sim \Gamma(5,2)$, we have X, Y positive and

$$
p_X(x) = \frac{2^3}{2!}x^2e^{-2x}
$$
, and $p_Y(y) = \frac{2^5}{4!}y^4e^{-2y}$ for $x, y > 0$.

Then the density of $X + Y$ is given by

$$
p_{X+Y}(x) = \int_0^x p_X(y)p_Y(x-y)dy
$$

\n
$$
= \int_0^x \frac{2^3}{2!} y^2 e^{-2y} \frac{2^5}{4!} (x-y)^4 e^{-2(x-y)} dy
$$

\n
$$
= \frac{2^8 e^{-2x}}{2!4!} \int_0^x y^2 (x-y)^4 dy
$$

\n
$$
= \frac{2^8 e^{-2x}}{2!4!} \left(-\frac{1}{5} y^2 (y-x)^5 \Big|_0^x + \frac{2}{5} \int_0^x y (x-y)^5 dy \right)
$$

\n
$$
= \frac{2^8 e^{-2x}}{2!4!} \left(\frac{2}{5} \int_0^x y (x-y)^5 dy \right)
$$

\n
$$
= \frac{2^8 e^{-2x}}{2!4!} \left(\frac{2}{5} \cdot \frac{1}{6} \int_0^x (x-y)^6 dy \right)
$$

\n
$$
= \frac{2^8 e^{-2x}}{2!4!} \left(\frac{2}{5} \cdot \frac{1}{6} \cdot \frac{-1}{7} (x-y)^7 \Big|_0^x \right)
$$

\n
$$
= \frac{2^8 e^{-2x}}{2!4!} \left(\frac{2}{5} \cdot \frac{1}{6} \cdot \frac{1}{7} x^7 \right)
$$

\n
$$
= \frac{2^8}{7!} x^7 e^{-2x},
$$

for $x > 0$ and equals to zero elsewhere. Hence $X + Y \sim \Gamma(8, 2)$.

6.8 Let X be the number of votes Richard gets. Then $X \sim Bin(100, 0.49)$, so that $\mathbf{E}X = (100)(0.49) = 49$ and $\text{Var}X = (100)(0.49)(0.51) = 24.99$. Now

$$
P(\text{Richard wins}) = P(X > 50) = P\left(\frac{X - 49}{\sqrt{24.99}} > \frac{1}{\sqrt{24.99}}\right).
$$

By Normal approximation, the probability is approximately $P\left(Z>\frac{1}{\sqrt{24.99}}\right)$, where $Z \sim \mathcal{N}(0, 1)$. Hence

$$
P(\text{Richard wins}) \approx P\left(Z > \frac{1}{\sqrt{24.99}}\right) \approx 0.4207.
$$

8.11 Recall that the generating function of a random variable X is given by

$$
\varphi_X(s) := \mathbf{E} s^X = \sum_{k=0}^{\infty} P(X=k) s^k.
$$

Since X and Y are independent, we have

$$
F_{X+2Y+1}(t) = \mathbf{E} s^{(X+2Y+1)} = s \mathbf{E} (s^X s^{2Y}) = s \mathbf{E} s^X \mathbf{E} (s^2)^Y = s \varphi_X(t) \varphi_Y(s^2).
$$

Since $X \sim Bin(3, 0.4)$ and $Y \sim Geo(0.3)$, we have

$$
\varphi_X(s) = (0.4s + 0.6)^3,
$$

and

$$
\varphi_Y(s) = \frac{0.3s}{1 - 0.7s}.
$$

Hence,

$$
\varphi_{X+2Y+1}(s) = s(0.4s+0.6)^3 \frac{0.3s^2}{1-0.7s^2} = (0.4s+0.6)^3 \frac{3s^3}{10-7s^2}.
$$

8.21 Note that, for $X = -1$, $F_X(t) = e^{-t}$, and for $Y \sim \Gamma(2, 2)$, $F_Y(t) = \left(\frac{2}{2}\right)^t$ $2 - t$ \setminus^2 . Since X is a constant, X and Y must be independent. Hence

$$
F_{X+Y}(t) = F_X(t)F_Y(t) = e^{-t} \left(\frac{2}{2-t}\right)^2 = \frac{4e^{-t}}{(2-t)^2}.
$$

Thus the distribution corresponds to $X + Y$.

9.19 (The number 100 is changed to 50.) Let $S = X_1 + \cdots + X_{50} + Y_1 + \cdots + Y_{40}$. Then $\mu = \mathbf{E}S = 50p_X + 40p_Y = 64.$ Note that $50 = \mu(1-\delta)$, where $\delta = \frac{7}{26}$ 32 . Hence, by Chernov's inequality,

$$
P(S \le 50) = P(S \le \mu(1 - \delta))
$$

\n
$$
\le \exp\left(-\frac{\delta^2 \mu}{2}\right)
$$

\n
$$
= \exp(-\frac{49}{32})
$$

\n
$$
\approx 0.2163.
$$

10.13 Note that the triangle can also be described as

$$
\{(x,y)\in\mathbb{R}^2 \mid 0\leq x\leq 1, 0\leq y\leq x\} = \{(x,y)\in\mathbb{R}^2 \mid 0\leq y\leq 1, y\leq x\leq 1\}.
$$

So, the marginal density of Y is given by

$$
p_Y(y) = \int p(x, y) dx = \begin{cases} \int_y^1 2dx = 2(1 - y) & \text{if } 0 \le y \le 1, \\ 0 & \text{otherwise.} \end{cases}
$$

Thus,

$$
p_{X|Y}(x | y) = \frac{p(x, y)}{p_Y(y)} = \begin{cases} \frac{1}{1 - y} & \text{if } 0 \le y < 1, \ y \le x \le 1, \\ 0 & \text{otherwise.} \end{cases}
$$

Hence, for $0\leq y<1,$

$$
\mathbf{E}(X^2 | Y = y) = \int x^2 p_{X|Y}(x | y) dx
$$

=
$$
\int_y^1 \frac{x^2}{1 - y} dx
$$

=
$$
\frac{x^3}{3(1 - y)} \Big|_y^1
$$

=
$$
\frac{1}{3}(y^2 + y + 1).
$$