

Review.

- Normal r.v  $X$  with parameters  $\mu$  and  $\sigma^2$ :

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad x \in (-\infty, \infty)$$

$Z := \frac{X-\mu}{\sigma}$  is a standard normal r.v.

(with mean 0, Variance 1)

- Exponential r.v with parameter  $\lambda$  ( $\lambda > 0$ ),

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x > 0 \\ 0, & \text{otherwise.} \end{cases}$$

## § 5.7 The distribution of a function of a cts r.v.

Q: Let  $X$  be a cts r.v. with density  $f_X$   
Let  $g: \mathbb{R} \rightarrow \mathbb{R}$  be a function.

Let  $Y = g(X)$ .

How to find the distribution of  $Y$ ?

Exer 2. Let  $X$  be a cts r.v. with density  $f_X$ .

Let  $Y = X^2$ .

Find the P.d.f. of  $Y$ .

Solution: We first calculate the CDF of  $Y$ :

$$F_Y(y) = P\{Y \leq y\}$$

$$= P\{X^2 \leq y\}$$

$$= \begin{cases} 0 & \text{if } y < 0 \\ P\{-\sqrt{y} \leq X \leq \sqrt{y}\} & \text{if } y \geq 0 \end{cases}$$
$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

Taking derivative of  $F_Y$  with respect to  $y$  gives

$$f_Y(y) = \begin{cases} 0 & \text{if } y < 0 \\ f_X(\sqrt{y}) \cdot \frac{1}{2\sqrt{y}} + f_X(-\sqrt{y}) \cdot \frac{1}{2\sqrt{y}} & \text{if } y > 0 \end{cases}$$

Ex 3. Let  $X$  be an exponential r.v with parameter  $\lambda$ . Let  $Y = \frac{1}{X}$   
 Find the pdf of  $Y$ .

Solution: First notice that

$$P\{X \leq 0\} = 0$$

Now

$$\begin{aligned} F_Y(y) &= P\{Y \leq y\} \\ &= P\left\{\frac{1}{X} \leq y\right\} \\ &= \begin{cases} P\left\{X \geq \frac{1}{y}\right\} & \text{if } y > 0 \\ 0 & \text{if } y \leq 0 \end{cases} \end{aligned}$$

$$= \begin{cases} 1 - F_X\left(\frac{1}{y}\right) & \text{if } y > 0 \\ 0 & \text{otherwise} \end{cases}$$

Taking derivative of  $F_Y$  wrt  $y$  gives

$$f_Y(y) = \begin{cases} -f_X\left(\frac{1}{y}\right) \cdot \left(-\frac{1}{y^2}\right) & y > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} f_X\left(\frac{1}{y}\right) \cdot \frac{1}{y^2} & \text{if } y > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \lambda \cdot e^{-\lambda \frac{1}{y}} \cdot \frac{1}{y^2} & \text{if } y > 0 \\ 0 & \text{otherwise} \end{cases}$$

Prop. 4: Let  $X$  be a r.v. with density  $f_X$ . Let  $g: \mathbb{R} \rightarrow \mathbb{R}$  be a differentiable, strictly monotone function. Let  $Y = g(X)$ , then the p.d.f of  $Y$  is given by

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \cdot \left| \frac{d g^{-1}(y)}{dy} \right| & \text{if } y = g(x) \text{ for some } x \\ 0 & \text{otherwise.} \end{cases}$$

Pf. Suppose that  $g$  is strictly increasing.

Then

$$\begin{aligned}
 F_Y(y) &= P\{Y \leq y\} \\
 &= P\{g(X) \leq y\} \\
 &= \begin{cases} 1, & \text{if } y > \max_{x \in \mathbb{R}} g(x) \\ P\{X \leq g^{-1}(y)\}, & \text{if } y \in \text{Range}(g) \\ 0, & \text{if } y < \min_{x \in \mathbb{R}} g(x) \end{cases}
 \end{aligned}$$

$$\begin{aligned}
 P\{X \leq g^{-1}(y)\} \\
 = F_X(g^{-1}(y)).
 \end{aligned}$$

Taking derivative of  $F_Y$  gives

$$f_Y(y) = \begin{cases} 0 & \text{if } y \notin \text{Range}(g) \\ f_X(g^{-1}(y)) \cdot \frac{d g^{-1}(y)}{dy}, & \text{if } y = g(x) \text{ for some } x. \end{cases}$$

Example 5. Let  $Z$  be a standard normal r.v. Find the density of  $Z^3$ .

Solution: Let  $g(x) = x^3$ . Then  $g$  is monotone increasing on  $\mathbb{R}$  with  $\text{range}(g) = \mathbb{R}$ .

Hence

$$\begin{aligned} f_{Z^3}(z) &= f_Z(g^{-1}(z)) \cdot \frac{d g^{-1}(z)}{dz} \\ &= f_Z(z^{1/3}) \cdot \frac{1}{3} \cdot z^{-2/3} \\ &= \frac{1}{3\sqrt{2\pi}} \cdot z^{-2/3} \cdot e^{-\frac{z^2}{2}}. \end{aligned}$$