

Answer to a Question Posed in class March 10:

Claim: Let y be recurrent, then $\lim_{n \rightarrow \infty} \frac{N_n(y)}{n} \stackrel{\text{prob.}}{=} \frac{1_{\{T_y < \infty\}}}{m_y}$.

Proof. $\{T_y < \infty\} = \{1 \leq T_y < \infty\} = \bigcup_{k=1}^{\infty} \{T_y = k\}$

Consider the limit in two cases:

$T_y = k$ for $k \in \{1, 2, \dots\}$ and $T_y = \infty$

Case 1. $T_y = \infty$, to show: $\lim_{n \rightarrow \infty} \frac{N_n(y)}{n} \stackrel{\text{prob.}}{=} 0$. $\left(\because 1_{\{T_y < \infty\}} = 0 \right.$
 $\left. \text{if } T_y = \infty \right)$

Indeed, recall " $T_y = \infty$ " means

$$X_m \neq y, \quad \forall m \geq 1$$

$$\therefore N_n(y) = \sum_{m=1}^n 1_y(X_m) = 0, \quad \forall n \geq 1$$

$$\therefore \lim_{n \rightarrow \infty} \frac{N_n(\gamma)}{n} = \lim_{n \rightarrow \infty} \frac{0}{n} = \lim_{n \rightarrow \infty} 0 = 0.$$

Case 2. $T_\gamma = k$ for a given finite $k \in \{1, 2, \dots\}$.

to show: $\lim_{n \rightarrow \infty} \frac{N_n(\gamma)}{n} = \frac{1}{m_\gamma}$. ($\because \mathbb{1}_{\{T_\gamma < \infty\}} = 1$
if $T_\gamma = k < \infty$)

Indeed, recall

$$k = T_\gamma = \min \{ m \geq 1 : X_m = \gamma \}$$

implying

$$X_1 \neq \gamma, \dots, X_{k-1} \neq \gamma, X_k = \gamma \quad (*)$$

We then write: for any $n > k$,

$$\begin{aligned}
\frac{N_n(y)}{n} &= \frac{\sum_{m=1}^n 1_y(X_m)}{n} \\
&= \frac{\sum_{m=1}^k 1_y(X_m) + \sum_{m=k+1}^n 1_y(X_m)}{n} \\
&= \left(\frac{1}{n}\right) + \frac{\sum_{m=k+1}^n 1_y(X_m)}{n} \quad (***) \\
&\quad (\because \sum_{m=1}^k 1_y(X_m) = 1 \text{ by } (*))
\end{aligned}$$

Note :

$$\lim_{n \rightarrow \infty} \frac{1}{n} = 0$$

and

$$\lim_{n \rightarrow \infty} \frac{\sum_{m=k+1}^n 1_Y(X_m)}{n} = \lim_{n \rightarrow \infty} \frac{\sum_{m=k+1}^n 1_Y(X_m)}{n-k} \cdot \frac{n-k}{n}$$

$$= \lim_{n \rightarrow \infty} \frac{\sum_{m=k+1}^n 1_Y(X_m)}{n-k}$$

Now, given $X_k = y \in S_R$ at time k , we may restrict our chain over some C_i that is an irreducible closed set of recurrent states in the state decomposition

$$S_R = \bigcup_{i=1}^r C_i$$

(disjoint union)

As such, we have: $X_m \in C_i, \forall m \geq k$

Due to the Markovian property, we may regard the chain starting at time k , then by the thm in the **COURSE NOTE 10**,

$$\lim_{n \rightarrow \infty} \frac{\sum_{m=k+1}^n \mathbb{1}_y(X_m)}{n-k} = \lim_{n \rightarrow \infty} \frac{\sum_{m=1}^{n-k} \mathbb{1}_y(\tilde{X}_m)}{n-k}$$

$$\left(\tilde{X}_m \stackrel{\text{def.}}{=} X_{m+k} \right)_{m=0, 1, \dots}$$

$$\stackrel{\text{prob.}}{=} \frac{1}{m_y}$$

Plugging those back to (**),

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{N_n(y)}{n} &\stackrel{\text{prob.}}{=} \lim_{n \rightarrow \infty} \frac{1}{n} + \lim_{n \rightarrow \infty} \frac{\sum_{m=k+1}^n 1_y(X_m)}{n} \\ &\stackrel{\text{prob.}}{=} 0 + \frac{1}{m_y} \\ &= \frac{1}{m_y}. \end{aligned}$$

Combining two cases above gives the proof of the desired result $\lim_{n \rightarrow \infty} \frac{N_n(y)}{n} = \frac{1_{\{T_y < \infty\}}}{m_y}$, $\forall y \in S_R$ for a general Markov chain. ###