**ENGG 2430 / ESTR 2004:** Probability and Statistics Spring 2019

# **3. Independence and Random Variables**

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#### Independence of two events



Let  $E_1$  be "first coin comes up H"  $E_2$  be "second coin comes up H" Then  $\mathbf{P}(E \perp E) = \mathbf{P}(E)$ 

Then  $P(E_2 | E_1) = P(E_2)$  $P(E_2 \cap E_1) = P(E_2)P(E_1)$ 

Events *A* and *B* are independent if  $\mathbf{P}(A \cap B) = \mathbf{P}(A) \mathbf{P}(B)$ 

#### **Examples of (in)dependence**



Let  $E_1$  be "first die is a 4"  $S_6$  be "sum of dice is a 6"  $S_7$  be "sum of dice is a 7"

 $P(E, NS_{c}) \stackrel{?}{=} P(E, ) \cdot P(S_{c})$   $1/36 \neq 1/6 5/36$  NOT  $E_1$  and  $S_6$ ?  $P(E_1 \cap S_7) = P(E_1) \cdot P(S_7)$   $\frac{1}{36} \quad \frac{1}{16} \quad \frac{1}{16}$   $P(S_2 \cap S_7) \stackrel{!}{=} P(S_2) \cdot P(S_7)$   $\stackrel{!}{\otimes} \quad \frac{1}{4} \quad \frac{1}{50} \quad \frac{1}{50}$  $E_1$  and  $S_7$ ?  $S_6$  and  $S_7$ ?



If A and B are independent, then A and  $B^c$  are also independent.

**Proof:** 

$$P(A \cap B^{c}) = P(A) - P(A \cap B) = P(A) - P(A)P(B)$$
  
=  $P(A)(I - P(B))$   
=  $P(A)P(B^{c})$   
A \cap B^{c}



Events A, B, and C are independent if  $\mathbf{P}(A \cap B) = \mathbf{P}(A) \mathbf{P}(B)$   $\mathbf{P}(B \cap C) = \mathbf{P}(B) \mathbf{P}(C)$   $\mathbf{P}(A \cap C) = \mathbf{P}(B) \mathbf{P}(C)$ and  $\mathbf{P}(A \cap B \cap C) = \mathbf{P}(A) \mathbf{P}(B) \mathbf{P}(C)$ .

#### (In)dependence of three events



Let  $E_1$  be "first die is a 4"  $E_2$  be "second die is a 3"  $S_7$  be "sum of dice is a 7"



#### (In)dependence of three events



Let A be "first roll is 1, 2, or 3 "  $\frac{1}{2}$ B be "first roll is 3, 4, or 5"  $\frac{1}{2}$ C be "sum of rolls is 9"  $\frac{1}{2}$ 

A, B? 
$$\frac{1}{2} \times \frac{1}{2} \pm \frac{1}{6}$$
  
A, C?  $\frac{1}{2} \times \frac{1}{9} \pm \frac{1}{12}$   
B, C?  $\frac{1}{2} \times \frac{1}{9} \pm \frac{1}{12}$   
A, B, C?  $\frac{1}{2} \times \frac{1}{2} \times \frac{1}{9} = \frac{1}{36}$ 

Events  $A_1, A_2, \ldots$  are independent if for every subset of the events, the probability of the intersection is the product of their probabilities.

#### **Algebra of independent events**

Independence is preserved if we replace some event(s) by their complements, intersections, unions



$$P(ER) = 70\%$$
  
 $P(WR) = 75\%$   
 $P(KT) = 95\%$   
 $P(TW) = 85\%$ 

$$P(E) = P(ER \cap (WR \cup (KT \cap TW)))$$

$$= P(ER) \cdot P(WR \cup (KT \cap TW))$$

$$70\% \qquad 25\%$$

$$P(WR \cup (KT \cap TW)) = 1 - (1 - P(WR)) \cdot (1 - P(KT)) \cdot (1 - P(KT))$$

Alice wins 60% of her ping pong matches against Bob. They meet for a 3 match playoff. What are the chances that Alice will win the playoff?

#### **Probability model**

Let  $A_i$  be the event Alice wins match i

**Assume**  $P(A_1) = P(A_2) = P(A_3) = 0.6$ 

Also assume  $A_1, A_2, A_3$  are independent

outcome	probability
AAA	0.63
AAB	$0.6^{2} \cdot 0.4$
ABA	0.62.0.4
BAA	0.62.0.4

 $\mathbf{P}(A) = 0.6^3 + 3 \cdot 0.6^2 \cdot 0.4 = 0.648$ 

### $\it n$ trials, each succeeds independently with probability $\it p$

The probability at least *k* out of *n* succeed is

$$\binom{n}{k} p^{k} (1-p)^{n+k} + \binom{n}{k+1} p^{k+1} (1-p)^{n+k-1} + \binom{n}{n} p^{n}$$

#### Playoffs



The probability that Alice wins an *n* game tournament

### The Lakers and the Celtics meet for a 7-game playoff. They play until one team wins four games.



Suppose the Lakers win 60% of the time. What is the probability that all 7 games are played?

ALL 7 PLAYED I FIRST 6 HAVE 3 LAKERS WINSH 3 LAKERS LOSSES (1) 3 13

 $P(E) = \binom{6}{3} 0.6^3 0.4^3$ 

#### A and B are independent conditioned on F if $\mathbf{P}(A \cap B \mid F) = \mathbf{P}(A \mid F) \mathbf{P}(B \mid F)$

#### **Alternative definition:**

$$\mathbf{P}(\mathcal{A} \mid B \cap F) = \mathbf{P}(\mathcal{A} \mid F)$$



It is  $\neq$  on Monday. Will it  $\bigcirc$  on Wednesday? P(W|MnT) = P(W|T)  $P(T) = P(T|M)P(M) + P(T|M^{c})P(M^{c}) = 0.8$   $= P(W|T)P(T) + P(W|T^{c})P(T^{c}) = 0.72$   $P(W) = P(W|T)P(T) + P(W|T^{c})P(T^{c}) = 0.72$ 

#### **Conditioning does not preserve independence**



Let  $E_1$  be "first die is a 4"  $E_2$  be "second die is a 3"  $S_7$  be "sum of dice is a 7"

 $E_{1_1}E_2 \text{ INDEPENDENT BUT}$   $P(E_1 \cap E_2 | S_7) \neq P(E_1 | S_7) P(E_2 | S_7)$   $\frac{1}{6} \frac{1}{6} \frac{1}{6} \frac{1}{6}$ 

The New York Times

### *Crazy Rich Asians' Has Soared, but It May Not Fly in China*



#### **Conditioning may destroy dependence**



A discrete random variable assigns a discrete value to every outcome in the sample space.



{ HH, HT, TH, TT }

N = number of Hs

The probability mass function (p.m.f.) of discrete random variable *X* is the function

$$p(x) = P(X = x)$$



 $\{ \text{HH, HT, TH, TT} \}$  $\frac{1}{4}$   $\frac{1}{4}$   $\frac{1}{4}$   $\frac{1}{4}$ 

N = number of Hs

 $p(0) = P(N = 0) = P({TT}) = 1/4$   $p(1) = P(N = 1) = P({HT, TH}) = 1/2$  $p(2) = P(N = 2) = P({HH}) = 1/4$  We can describe the p.m.f. by a table or by a chart.



Two six-sided dice are tossed. Calculate the p.m.f. of the difference *D* of the rolls.



#### What is the probability that D > 1? D is odd?

D=d 11 12 13 14 15 16 0 1/36 26 21 22 23 24 25 -5 2/36 32 33 34 35 31 36 3/36 -3 43 44 45 41 42 46 4/36 -2 52 53 54 55 56 51 5/36 - 1 61 62 63 64 65 66 6 /36 わ 5136 4/36 2 3/36 3 2/36 1/36

Binomial(*n*, *p*): Perform *n* independent trials, each of which succeeds with probability *p*.

X = number of successes

#### **Examples**

Toss *n* coins. "number of heads" is  $Binomial(n, \frac{1}{2})$ . Toss *n* dice. "Number of **•**, s" is Binomial(n, 1/6).

### Toss *n* coins. Let *C* be the number of consecutive changes (HT or TH).

Examples:	ω	$C(\omega)$
	ННННННН	0
	THHHHHT	2
	HTHHHHT	3

Then C is  $Binomial(n-1, \frac{1}{2})$ .

#### Draw a 10-card hand from a 52-card deck.

Let N = number of aces among the drawn cards

#### Is N a Binomial(10, 1/13) random variable?

## **NO!** Trial outcomes are not independent.

#### **Probability mass function**

**If** *X* **is** Binomial(*n*, *p*), **its p.m.f. is** 

$$p(k) = P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$







Let  $X_1, X_2, \ldots$  be independent trials with success *p*.

A Geometric(p) random variable N is the time of the first success among  $X_1, X_2, ...$ :

N = first (smallest) *n* such that  $X_n = 1$ .

So 
$$P(N = n) = P(X_1 = 0, ..., X_{n-1} = 0, X_n = 1)$$
  
=  $(1 - p)^{n-1}p$ .

This is the p.m.f. of N.





About 10% of the apples on your farm are rotten.

You sell 10 apples. How many are rotten?

#### **Probability model**

Number of rotten apples you sold is Binomial(n = 10, p = 1/10).





#### You improve productivity; now only 5% apples rot.

You can now sell 20 apples.

N is now Binomial(20, 1/20).





A  $Poisson(\lambda)$  random variable has this p.m.f.:

$$p(k) = e^{-\lambda} \lambda^k / k!$$
  $k = 0, 1, 2, 3, ...$ 

### Poisson random variables do not occur "naturally" in the sample spaces we have seen.

They approximate Binomial(n, p) random variables when  $\lambda = np$  is fixed and *n* is large (so *p* is small)

$$p_{\text{Poisson}(\lambda)}(k) = \lim_{n \to \infty} p_{\text{Binomial}(n, \lambda/n)}(k)$$

#### **Functions of random variables**



If *X* is a random variable with p.m.f.  $p_X$ , then Y = f(X) is a random variable with p.m.f.

$$p_Y(y) = \sum_{x: f(x) = y} p_X(x).$$

