

# 8.5 Achievability of $R_{l}(D)$

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

• An i.i.d. source  $\{X_k : k \ge 1\}$  with generic random variable  $X \sim p(x)$  is given.

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

- An i.i.d. source  $\{X_k : k \ge 1\}$  with generic random variable  $X \sim p(x)$  is given.
- For every random variable  $\hat{X}$  taking values in  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ , prove that the rate-distortion pair  $(I(X; \hat{X}), D)$  is achievable by showing for large n the existence of a rate-distortion code such that

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

- An i.i.d. source  $\{X_k : k \ge 1\}$  with generic random variable  $X \sim p(x)$  is given.
- For every random variable  $\hat{X}$  taking values in  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ , prove that the rate-distortion pair  $(I(X; \hat{X}), D)$  is achievable by showing for large n the existence of a rate-distortion code such that

1. the rate of the code is not more than  $I(X; \hat{X}) + \epsilon$ ;

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

- An i.i.d. source  $\{X_k : k \ge 1\}$  with generic random variable  $X \sim p(x)$  is given.
- For every random variable  $\hat{X}$  taking values in  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ , prove that the rate-distortion pair  $(I(X; \hat{X}), D)$  is achievable by showing for large n the existence of a rate-distortion code such that
  - 1. the rate of the code is not more than  $I(X; \hat{X}) + \epsilon$ ;
  - 2.  $d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + \epsilon$  with probability almost 1.

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

- An i.i.d. source  $\{X_k : k \ge 1\}$  with generic random variable  $X \sim p(x)$  is given.
- For every random variable  $\hat{X}$  taking values in  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ , prove that the rate-distortion pair  $(I(X; \hat{X}), D)$  is achievable by showing for large n the existence of a rate-distortion code such that

1. the rate of the code is not more than  $I(X; \hat{X}) + \epsilon$ ;

2.  $d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + \epsilon$  with probability almost 1.

• Minimize  $I(X; \hat{X})$  over all such  $\hat{X}$  to conclude that  $(R_I(D), D)$  is achievable.

**Theorem 8.17 (The Rate-Distortion Theorem)**  $R(D) = R_I(D)$ .

- An i.i.d. source  $\{X_k : k \ge 1\}$  with generic random variable  $X \sim p(x)$  is given.
- For every random variable  $\hat{X}$  taking values in  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ , prove that the rate-distortion pair  $(I(X; \hat{X}), D)$  is achievable by showing for large n the existence of a rate-distortion code such that
  - 1. the rate of the code is not more than  $I(X; \hat{X}) + \epsilon$ ;

2.  $d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + \epsilon$  with probability almost 1.

- Minimize  $I(X; \hat{X})$  over all such  $\hat{X}$  to conclude that  $(R_I(D), D)$  is achievable.
- This implies that  $R_I(D) \ge R(D)$ .

Parameter Settings

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### Parameter Settings

- 1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.
- 2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \leq \frac{1}{n} \log M \leq I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

The Random Coding Scheme

#### Parameter Settings

- 1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.
- 2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \dots, \hat{\mathbf{X}}(M)$ .

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \dots, \hat{\mathbf{X}}(M)$ .



#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \dots, \hat{\mathbf{X}}(M)$ .



#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \dots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence  $\mathbf{X}$  into an index K in the set  $\mathcal{I} = \{1, 2, \cdots, M\}$ . The index K takes the value i if

(a) 
$$(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta},$$

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

- (a)  $(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta},$
- (b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  $i' \leq i$ ;

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a) 
$$(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$$
,  
(b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  
 $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a) 
$$(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$$
,  
(b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  
 $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \dots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \cdots, M\}$ . The index K takes the value i if

(a) 
$$(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$$
,  
(b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  
 $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value i if

(a) 
$$(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$$
,  
(b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  
 $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value i if

(a) 
$$(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$$
,  
(b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  
 $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a)  $(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$ , (b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

Remarks

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a)  $(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$ , (b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

#### Remarks

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a)  $(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$ , (b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

#### Remarks

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

•  $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a)  $(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$ , (b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

#### Remarks

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and  $\hat{X}$  with  $Ed(X, \hat{X}) \leq D$ , where  $0 \leq D \leq D_{max}$ . Let  $\delta$  be specified later.

2. Let M be an integer satisfying

$$I(X; \hat{X}) + \frac{\epsilon}{2} \le \frac{1}{n} \log M \le I(X; \hat{X}) + \epsilon,$$

where n is sufficiently large.

#### The Random Coding Scheme

1. Construct a codebook  $\mathcal{C}$  of an (n, M) code by randomly generating M codewords in  $\hat{\mathcal{X}}^n$  independently and identically according to  $p(\hat{x})^n$ . Denote these codewords by  $\hat{\mathbf{X}}(1), \hat{\mathbf{X}}(2), \cdots, \hat{\mathbf{X}}(M)$ .



2. Reveal the codebook  $\mathcal{C}$  to both the encoder and the decoder.

3. The source sequence **X** is generated according to  $p(x)^n$ .

4. The encoder encodes the source sequence **X** into an index K in the set  $\mathcal{I} = \{1, 2, \dots, M\}$ . The index K takes the value *i* if

(a)  $(\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta}$ , (b) for all  $i' \in \mathcal{I}$ , if  $(\mathbf{X}, \hat{\mathbf{X}}(i')) \in T^n_{[X\hat{X}]\delta}$ , then  $i' \leq i$ ;

i.e., if there exists more than one i satisfying (a), let K be the largest one. Otherwise, K takes the constant value 1.

5. The index K is delivered to the decoder.

6. The decoder outputs  $\hat{\mathbf{X}}(K)$  as the reproduction sequence  $\hat{\mathbf{X}}$ .

#### Remarks

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.
- 2. If  $K \neq 1$ , then  $\hat{\mathbf{X}}(K)$  is jointly typical with  $\mathbf{X}$ .

**Performance Analysis** 

**Performance Analysis**
1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

•  $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in  $\mathcal{C}$  is jointly typical with  $\mathbf{X}$ .

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

 $\Pr\{K=1|\mathbf{X}=\mathbf{x}\}$ 

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ \leq \quad \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{x}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

$$\begin{aligned} \Pr\{K &= 1 | \mathbf{X} = \mathbf{x} \} \\ &\leq & \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x} \} \\ &= & \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x} \} \\ &= & (\Pr\{E_1^c | \mathbf{X} = \mathbf{x} \})^{M-1} \\ &= & (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x} \})^{M-1}. \end{aligned}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\begin{aligned} \Pr\{K = 1 | \mathbf{X} = \mathbf{x} \} \\ &\leq \quad \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x} \} \\ &= \quad \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x} \} \\ &= \quad \left(\Pr\{E_1^c | \mathbf{X} = \mathbf{x} \}\right)^{M-1} \\ &= \quad \left(1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x} \}\right)^{M-1}. \end{aligned}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\mathbf{X} \in S^n_{[X]\delta}} \approx 1$  for large n.

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

for n sufficiently large.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large n.

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\mathbf{X} \in S^n_{[X]\delta}} \approx 1$  for large n.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\mathbf{X} \in S_{[X]\delta}^n} \approx 1$  for large n.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_1 | \mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^n\right\}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

•

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

$$\Pr\{E_1 | \mathbf{X} = \mathbf{x}\} = \Pr\left\{ (\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^n \right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^n} p(\hat{\mathbf{x}})$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in  $\mathcal{C}$  is jointly typical with  $\mathbf{X}$ .

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \dots \cap E_M^c$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large n. 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_1 | \mathbf{X} = \mathbf{x}\} = \Pr\left\{ (\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^n \right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^n} p(\hat{\mathbf{x}})$$

Theorem 6.7 (Consistency) If  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ , then  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$

Theorem 6.10 (Conditional Strong AEP) If  $|T_{[Y|X]\delta}^{n}(\mathbf{x})| \geq 1$ , then  $2^{n(H(Y|X)-\nu)} \leq |T_{[Y|X]\delta}^{n}(\mathbf{x})| \leq 2^{n(H(Y|X)+\nu)}$ , where  $\nu \to 0$  as  $n \to \infty$  and  $\delta \to 0$ .

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\frac{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})}{[\hat{X}|X]\delta}$$

Theorem 6.10 (Conditional Strong AEP) If  $|T_{[Y|X]\delta}^{n}(\mathbf{x})| \geq 1$ , then  $2^{n(H(Y|X)-\nu)} \leq |T_{[Y|X]\delta}^{n}(\mathbf{x})| \leq 2^{n(H(Y|X)+\nu)}$ , where  $\nu \to 0$  as  $n \to \infty$  and  $\delta \to 0$ .

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large n. 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq \frac{2^{n(H(\hat{X}|X)-\xi)}2^{-n(H(\hat{X})+\eta)}}{2^{n(H(\hat{X}|X)-\xi)}2^{-n(H(\hat{X})+\eta)}}$$

Theorem 6.10 (Conditional Strong AEP) If  $|T_{[Y|X]\delta}^{n}(\mathbf{x})| \geq 1$ , then  $2^{n(H(Y|X)-\nu)} \leq |T_{[Y|X]\delta}^{n}(\mathbf{x})| \leq 2^{n(H(Y|X)+\nu)}$ , where  $\nu \to 0$  as  $n \to \infty$  and  $\delta \to 0$ .

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{-n(H(\hat{X})-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(X;\hat{X})+\xi)},$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in  $\mathcal{C}$  is jointly typical with  $\mathbf{X}$ .

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in  $\mathcal{C}$  is jointly typical with  $\mathbf{X}$ .

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\xi)},$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\xi)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ .

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ .
1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\begin{aligned} &\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ &\leq &\Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\} \\ &= &\prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\} \\ &= &(\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1} \\ &= &(1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}. \end{aligned}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\begin{aligned} &\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ &\leq &\Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\} \\ &= &\prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\} \\ &= &(\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1} \\ &= &(1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}. \end{aligned}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\begin{aligned} &\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ &\leq &\Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\} \\ &= &\prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\} \\ &= &(\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1} \\ &= &(1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}. \end{aligned}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

Ρ

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^C \cap E_3^C \cap \dots \cap E_M^C | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^C | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^C | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large n. 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\mathbf{r}\{E_{\underline{1}}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(X;\hat{X})+\zeta)},$$

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

Ρ

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\mathbf{r}\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$

$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$

$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$

$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$

$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$

$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

1. The event  $\{K = 1\}$  occurs in one of the following two scenarios:

- $\hat{X}(1)$  is the only codeword in  $\mathcal{C}$  which is jointly typical with **X**.
- No codeword in C is jointly typical with **X**.

In other words, if K = 1, then **X** is jointly typical with none of the codewords  $\hat{X}(2), \hat{X}(3), \dots, \hat{X}(M)$ . We will show that  $\Pr\{K = 1\}$  can be made arbitrarily small.

2. Define the event

$$E_i = \left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \right\}.$$

3. Then

$$\{K=1\} \subset E_2^c \cap E_3^c \cap \cdots \cap E_M^c.$$

4. Since the codewords are generated i.i.d., conditioning on  $\{\mathbf{X} = \mathbf{x}\}$  for any  $\mathbf{x} \in \mathcal{X}^n$ , the events  $E_i$  are mutually independent and have the same probability.

5. Then for any  $\mathbf{x} \in \mathcal{X}^n$ ,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq \Pr\{E_2^c \cap E_3^c \cap \dots \cap E_M^c | \mathbf{X} = \mathbf{x}\}$$

$$= \prod_{i=2}^M \Pr\{E_i^c | \mathbf{X} = \mathbf{x}\}$$

$$= (\Pr\{E_1^c | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$= (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr{\{E_1 | \mathbf{X} = \mathbf{x}\}}$ :

$$\Pr\{E_{1} | \mathbf{X} = \mathbf{x}\} = \Pr\left\{ (\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n} \right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

9. Now

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n\left(I(X;\hat{X}) + \frac{\epsilon}{2}\right)}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

9. Now

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})+\xi)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M-1) \ln \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} < 0$$
  
$$\leq (M-1) \ln \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

.

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} < 0$$

$$\leq (\underline{M} - 1) \ln \left[ 1 - 2^{-n(I(X; \hat{X}) + \zeta)} \right]$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

•

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} < 0$$

$$\leq (\underline{M} - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( \frac{2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}}{1 - 1} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

.

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M-1) \ln \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]$$

$$\leq \left(2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1\right) \ln \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})+\xi)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M-1) \ln \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]$$

$$\leq \left(2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1\right) \ln \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ \leq \qquad (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ \leq \qquad (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad - \left( \frac{2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1}{2^{-n(I(X;\hat{X}) + \zeta)}} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ \leq \qquad (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad - \left( \frac{2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1}{2^{-n(I(X;\hat{X}) + \zeta)}} - 1 \right) \frac{2^{-n(I(X;\hat{X}) + \zeta)}}{2^{-n(I(X;\hat{X}) + \zeta)}}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ \leq \qquad (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \qquad - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \\ \leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right] \\ \leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X,\hat{X}) + \zeta)}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( \frac{2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1}{2^{-n(I(X;\hat{X}) + \zeta)}} - 1 \right) \frac{2^{-n(I(X;\hat{X}) + \zeta)}}{2^{-n(I(X;\hat{X}) + \zeta)}}$$

$$= - \left[ \frac{2^{n(\frac{\epsilon}{2} - \zeta)}}{2^{-n(I(X;\hat{X}) + \zeta)}} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}.$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

.

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S^{n}_{[X]\delta} = \{ \mathbf{x} \in T^{n}_{[X]\delta} : |T^{n}_{[\hat{X}|X]\delta}(\mathbf{x})| \ge 1 \},\$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}.$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^n = \{ \mathbf{x} \in T_{[X]\delta}^n : |T_{[\hat{X}|X]\delta}^n(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X; \hat{X}) + \zeta)}\right]^{M-1}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \rightarrow -\infty$$

$$\leq (M-1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^n = \{ \mathbf{x} \in T_{[X]\delta}^n : |T_{[\hat{X}|X]\delta}^n(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

.

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \rightarrow -\infty$$

$$\leq (M-1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\} \approx 1 \text{ for large } n.}$ 

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

.

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \rightarrow -\infty$$

$$\leq (M-1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

6. We will focus on  $\mathbf{x} \in S^n_{[X]\delta}$  where

$$S_{[X]\delta}^{n} = \{ \mathbf{x} \in T_{[X]\delta}^{n} : |T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})| \ge 1 \},$$

because  $\Pr{\{\mathbf{X} \in S_{[X]\delta}^n\}} \approx 1$  for large *n*.

7. For  $\mathbf{x} \in S_{[X]\delta}^n$ , obtain the following lower bound on  $\Pr\{E_1 | \mathbf{X} = \mathbf{x}\}$ :

$$\Pr\{E_{1}|\mathbf{X} = \mathbf{x}\} = \Pr\left\{(\mathbf{x}, \hat{\mathbf{X}}(1)) \in T_{[X\hat{X}]\delta}^{n}\right\}$$
$$= \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} p(\hat{\mathbf{x}})$$
$$\geq \sum_{\hat{\mathbf{x}} \in T_{[\hat{X}|X]\delta}^{n}(\mathbf{x})} 2^{-n(H(\hat{X})+\eta)}$$
$$\geq 2^{n(H(\hat{X}|X)-\xi)} 2^{-n(H(\hat{X})+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(H(\hat{X})-H(\hat{X}|X)+\xi+\eta)}$$
$$= 2^{-n(I(X;\hat{X})+\zeta)},$$

where  $\zeta = \xi + \eta \to 0$  as  $n \to \infty$  and  $\delta \to 0$ . 8. Therefore,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (1 - \Pr\{E_1 | \mathbf{X} = \mathbf{x}\})^{M-1}$$

$$\leq \left[1 - 2^{-n(I(X;\hat{X}) + \zeta)}\right]^{M-1}$$

•

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let *n* be sufficiently large and  $\delta$  be sufficiently small so that  $\epsilon$ 

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$
9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

9. Now

12. It follows that

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that  $\epsilon$ 

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$

9. Now

12. It follows that

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}. \qquad \Pr\{K = 1\}$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$

9. Now

12. It follows that

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$
  
= 
$$\sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

12. It follows that

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$
  
= 
$$\sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \underline{\mathbf{X}} = \underline{\mathbf{x}}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

12. It follows that

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$
  
= 
$$\sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \underline{\mathbf{X}} = \underline{\mathbf{x}}\} \Pr\{\underline{\mathbf{X}} = \underline{\mathbf{x}}\}$$

9. Now

12. It follows that

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

$$\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$
  
= 
$$\sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \underline{\mathbf{X}} = \underline{\mathbf{x}}\} \Pr\{\underline{\mathbf{X}} = \underline{\mathbf{x}}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}}{\Pr\{\mathbf{X} = \mathbf{x}\}} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \le \frac{\epsilon}{2}$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}}{\Pr\{X = \mathbf{x}\}} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \frac{\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}}{\Pr\{\mathbf{X} = \mathbf{x}\}} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \frac{\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}}{\Pr\{\mathbf{X} = \mathbf{x}\}} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \frac{1}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \frac{\Pr\{\mathbf{X} = \mathbf{x}\}}{\Pr\{\mathbf{X} = \mathbf{x}\}} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \frac{\Pr\{\mathbf{X} = \mathbf{x}\}}{\sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}}$$

$$= \frac{\epsilon}{2} \cdot \frac{\Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\}}{\sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \frac{\Pr\{\mathbf{X} = \mathbf{x}\}}{\mathbf{x} \in S_{[X]\delta}^{n}}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \frac{\Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}}{\mathbf{x} \notin S_{[X]\delta}^{n}}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \frac{\Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\}}{\Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\}} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \frac{\Pr\{\mathbf{X} \in S_{[X]}^{n} \delta\}}{\Pr\{\mathbf{X} \in S_{[X]}^{n} \delta\}} + \Pr\{\mathbf{X} \notin S_{[X]}^{n} \delta\}$$

$$\leq \frac{\epsilon}{2} \cdot \frac{1}{2} + \Pr\{\mathbf{X} \notin S_{[X]}^{n} \delta\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]}^{n} \delta\} + \Pr\{\mathbf{X} \notin S_{[X]}^{n} \delta\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]}^{n} \delta\}$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]}^{n} \delta} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]}^{n} \delta} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$< \frac{\epsilon}{2} + \underline{\delta}.$$

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that  $\epsilon$ 

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} + \delta.$$

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{c}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} + \delta.$$

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

13. By letting n be sufficiently large and  $\delta$  be sufficiently small so that both (1) and  $\delta < \frac{\epsilon}{2}$  are satisfied, we obtain

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let *n* be sufficiently large and  $\delta$  be sufficiently small so that  $\frac{\epsilon}{2} - \zeta > 0.$  (1)

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

13. By letting n be sufficiently large and  $\delta$  be sufficiently small so that both (1) and  $\delta < \frac{\epsilon}{2}$  are satisfied, we obtain

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let *n* be sufficiently large and  $\delta$  be sufficiently small so that  $\frac{\epsilon}{2} - \zeta > 0.$  (1)

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for 
$$\mathbf{x} \in S^n_{[X]\delta}$$
, for sufficiently large  $n$ ,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \le \frac{\epsilon}{2}$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

13. By letting n be sufficiently large and  $\delta$  be sufficiently small so that both (1) and  $\delta < \frac{\epsilon}{2}$  are satisfied, we obtain

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

13. By letting n be sufficiently large and  $\delta$  be sufficiently small so that both (1) and  $\delta < \frac{\epsilon}{2}$  are satisfied, we obtain

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

9. Now

$$\frac{1}{n}\log M \ge I(X;\hat{X}) + \frac{\epsilon}{2} \iff M \ge 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})}.$$

Then

$$\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$$

$$\leq (M - 1) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) \ln \left[ 1 - 2^{-n(I(X;\hat{X}) + \zeta)} \right]$$

$$\leq - \left( 2^{n(I(X;\hat{X}) + \frac{\epsilon}{2})} - 1 \right) 2^{-n(I(X;\hat{X}) + \zeta)}$$

$$= - \left[ 2^{n(\frac{\epsilon}{2} - \zeta)} - 2^{-n(I(X;\hat{X}) + \zeta)} \right].$$

10. Let n be sufficiently large and  $\delta$  be sufficiently small so that

$$\frac{\epsilon}{2} - \zeta > 0. \tag{1}$$

Then the upper bound on  $\ln \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\}$  tends to  $-\infty$  as  $n \to \infty$ , i.e.,  $\Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \to 0$  as  $n \to \infty$ .

11. This implies for  $\mathbf{x} \in S^n_{[X]\delta}$ , for sufficiently large n,

$$\Pr\{K=1|\mathbf{X}=\mathbf{x}\} \leq \frac{\epsilon}{2}.$$

12. It follows that

$$\Pr\{K = 1\}$$

$$= \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$+ \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} \Pr\{K = 1 | \mathbf{X} = \mathbf{x}\} \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$\leq \sum_{\mathbf{x} \in S_{[X]\delta}^{n}} \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} = \mathbf{x}\} + \sum_{\mathbf{x} \notin S_{[X]\delta}^{n}} 1 \cdot \Pr\{\mathbf{X} = \mathbf{x}\}$$

$$= \frac{\epsilon}{2} \cdot \Pr\{\mathbf{X} \in S_{[X]\delta}^{n}\} + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

$$\leq \frac{\epsilon}{2} \cdot 1 + \Pr\{\mathbf{X} \notin S_{[X]\delta}^{n}\}$$

13. By letting n be sufficiently large and  $\delta$  be sufficiently small so that both (1) and  $\delta < \frac{\epsilon}{2}$  are satisfied, we obtain

$$\Pr\{K=1\} < \epsilon.$$

**Proposition 6.13** With respect to a joint distribution p(x, y) on  $\mathcal{X} \times \mathcal{Y}$ , for any  $\delta > 0$ ,

$$\Pr\{\mathbf{X} \in S^n_{[X]\delta}\} > 1 - \delta$$

• Randomly generate M codewords in  $\hat{\mathcal{X}}^n$  according to  $p(\hat{x})^n$ , where n is large.

- Randomly generate M codewords in  $\hat{\mathcal{X}}^n$  according to  $p(\hat{x})^n$ , where n is large.
- $\mathbf{X} \in S^n_{[X]\delta}$  with high probability.

- Randomly generate M codewords in  $\hat{\mathcal{X}}^n$  according to  $p(\hat{x})^n$ , where n is large.
- $\mathbf{X} \in S^n_{[X]\delta}$  with high probability.
- For  $\mathbf{x} \in S_{[X]\delta}^n$ , by conditional strong AEP,

$$\Pr\left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \,|\, \mathbf{X} = \mathbf{x} \right\} \approx 2^{-nI(X;\hat{X})}$$

- Randomly generate M codewords in  $\hat{\mathcal{X}}^n$  according to  $p(\hat{x})^n$ , where n is large.
- $\mathbf{X} \in S^n_{[X]\delta}$  with high probability.
- For  $\mathbf{x} \in S_{[X]\delta}^n$ , by conditional strong AEP,

$$\Pr\left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \,|\, \mathbf{X} = \mathbf{x} \right\} \approx 2^{-nI(X;\hat{X})}.$$

• If M grows with n at a rate higher than  $I(X; \hat{X})$ , then the probability that there exists at least one  $\hat{\mathbf{X}}(i)$  which is jointly typical with the source sequence  $\mathbf{X}$  with respect to  $p(x, \hat{x})$  is high.

- Randomly generate M codewords in  $\hat{\mathcal{X}}^n$  according to  $p(\hat{x})^n$ , where n is large.
- $\mathbf{X} \in S^n_{[X]\delta}$  with high probability.
- For  $\mathbf{x} \in S_{[X]\delta}^n$ , by conditional strong AEP,

$$\Pr\left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \,|\, \mathbf{X} = \mathbf{x} \right\} \approx 2^{-nI(X;\hat{X})}.$$

- If M grows with n at a rate higher than  $I(X; \hat{X})$ , then the probability that there exists at least one  $\hat{\mathbf{X}}(i)$  which is jointly typical with the source sequence  $\mathbf{X}$  with respect to  $p(x, \hat{x})$  is high.
- Such an  $\hat{\mathbf{X}}(i)$ , if exists, would have  $d(\mathbf{X}, \hat{\mathbf{X}}) \approx Ed(X, \hat{X}) \leq D$ , because the joint relative frequency of  $(\mathbf{x}, \hat{\mathbf{X}}(i)) \approx p(x, \hat{x})$ . See the next proposition.

- Randomly generate M codewords in  $\hat{\mathcal{X}}^n$  according to  $p(\hat{x})^n$ , where n is large.
- $\mathbf{X} \in S^n_{[X]\delta}$  with high probability.
- For  $\mathbf{x} \in S_{[X]\delta}^n$ , by conditional strong AEP,

$$\Pr\left\{ (\mathbf{X}, \hat{\mathbf{X}}(i)) \in T^n_{[X\hat{X}]\delta} \,|\, \mathbf{X} = \mathbf{x} \right\} \approx 2^{-nI(X;\hat{X})}.$$

- If M grows with n at a rate higher than  $I(X; \hat{X})$ , then the probability that there exists at least one  $\hat{\mathbf{X}}(i)$  which is jointly typical with the source sequence  $\mathbf{X}$  with respect to  $p(x, \hat{x})$  is high.
- Such an  $\hat{\mathbf{X}}(i)$ , if exists, would have  $d(\mathbf{X}, \hat{\mathbf{X}}) \approx Ed(X, \hat{X}) \leq D$ , because the joint relative frequency of  $(\mathbf{x}, \hat{\mathbf{X}}(i)) \approx p(x, \hat{x})$ . See the next proposition.
- Use this  $\hat{\mathbf{X}}(i)$  to represent  $\mathbf{X}$  to satisfy the distortion constraint.

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then

 $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$
$$d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) = \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k)$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) = \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k)$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{\underline{x}, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{x, \hat{x}} \underline{d(x, \hat{x})} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) = \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k)$$
$$= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}})$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{\max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{\max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (\underline{np(x, \hat{x})} + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} \frac{d(x, \hat{x})}{(np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x}))} \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{\max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_k, \hat{x}_k) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} \frac{d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} \frac{d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} \frac{d(x, \hat{x}) (\mathbf{x} p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} \frac{d(x, \hat{x}) (\mathbf{x} p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + \underline{N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}})} - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + \underline{N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}})} - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \underline{n p(x, \hat{x})}) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \underline{n p(x, \hat{x})}) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \underline{p(x, \hat{x})} \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \mathbf{p}(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \mathbf{p}(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \underline{p(x, \hat{x})}) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - \underline{p(x, \hat{x})} \right) \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \\ &= \underbrace{Ed(X, \hat{X})}_{x, \hat{x}} + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) = \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k})$$

$$= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}})$$

$$= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x}))$$

$$= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right)$$

$$= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right)$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) = \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k})$$

$$= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}})$$

$$= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x}))$$

$$= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right)$$

$$= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right)$$

$$\leq Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left[\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right]$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \\ &= E d(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \\ &\leq E d(X, \hat{X}) + \sum_{x, \hat{x}} \frac{d(x, \hat{x})}{d(x, \hat{x})} \left| \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right| \end{aligned}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \\ &= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left( \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right) \\ &\leq Ed(X, \hat{X}) + \sum_{x, \hat{x}} \frac{d(x, \hat{x})}{x} \left| \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right| \\ &\leq Ed(X, \hat{X}) + \frac{d_{max}}{x, \hat{x}} \sum_{x, \hat{x}} \left| \frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x}) \right| \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &\leq Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \\ &\leq Ed(X, \hat{X}) + d_{max} \sum_{x, \hat{x}} \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

**Proof** For  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , consider

$$\begin{aligned} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &\leq Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \\ &\leq Ed(X, \hat{X}) + d_{max} \sum_{x, \hat{x}} \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \end{aligned}$$

 $\leq Ed(X, \hat{X}) + d_{max} \underline{\delta}$ 

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (np(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - np(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &\leq Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \\ &\leq Ed(X, \hat{X}) + d_{max} \sum_{x, \hat{x}} \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \\ &\leq Ed(X, \hat{X}) + d_{max} \delta \end{split}$$

$$d(\mathbf{x}, \hat{\mathbf{x}}) \le D + d_{max}\delta.$$

$$\begin{split} d(\mathbf{x}, \hat{\mathbf{x}}) &= \frac{1}{n} \sum_{k=1}^{n} d(x_{k}, \hat{x}_{k}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) \\ &= \frac{1}{n} \sum_{x, \hat{x}} d(x, \hat{x}) (n p(x, \hat{x}) + N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - n p(x, \hat{x})) \\ &= \sum_{x, \hat{x}} p(x, \hat{x}) d(x, \hat{x}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &= Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left(\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right) \\ &\leq Ed(X, \hat{X}) + \sum_{x, \hat{x}} d(x, \hat{x}) \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \\ &\leq Ed(X, \hat{X}) + d_{max} \sum_{x, \hat{x}} \left|\frac{1}{n} N(x, \hat{x} | \mathbf{x}, \hat{\mathbf{x}}) - p(x, \hat{x})\right| \\ &\leq Ed(X, \hat{X}) + d_{max} \delta \\ &\leq D + d_{max} \delta. \end{split}$$

1. For sufficiently large n, consider

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

=  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$
1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$ 

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$ 

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$ 

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$ 

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$ 

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \frac{\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\}}{\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}} \Pr\{K \neq 1\}$ 

1. For sufficiently large n, consider

- $= \frac{\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\}}{\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}} \Pr\{K \neq 1\}$
- $\leq \quad \underline{1} \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$

1. For sufficiently large n, consider

- $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$
- $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$

1. For sufficiently large n, consider

- $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$
- $\leq 1 \cdot \underline{\epsilon} + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$

1. For sufficiently large n, consider

- $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$
- $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$

1. For sufficiently large n, consider

- $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$
- $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot \underline{1}$

1. For sufficiently large n, consider

- $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$
- $\leq \quad 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$
- $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$

1. For sufficiently large n, consider

- $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$  $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$
- $\leq \quad 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$
- $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$   $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$   $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$   $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$ 

 $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

1. For sufficiently large n, consider

 $\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

1. For sufficiently large n, consider

 $\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

 $d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta.$ 

1. For sufficiently large n, consider

 $\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \underline{\delta}.$$

3. By taking  $\delta \leq \frac{\epsilon}{d_{max}}$ , we obtain

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \underline{\delta}.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \underline{\delta}.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \underline{\delta}.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}$ ,  $(\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}$ ,  $(\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}$ ,  $(\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}$ ,  $(\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$   $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$   $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$   $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$   $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$ 

2. Recall that conditioning on  $\{K \neq 1\}$ ,  $(\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}$ ,  $(\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$ 

 $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ \leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ = \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

1. For sufficiently large n, consider

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$$

$$= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$$

$$+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$$

$$\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$$

$$= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

1. For sufficiently large n, consider

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$$

$$= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$$

$$+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$$

$$\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$$

$$= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

1. For sufficiently large n, consider

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$$

$$= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$$

$$+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$$

$$\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$$

$$= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

and

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon.$$

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$   $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$   $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$   $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$   $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

 $\operatorname{and}$ 

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \le \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ ,  $(I(X; \hat{X}), D)$  is achievable.

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$   $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$   $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$   $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$   $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le \underline{I(X;\hat{X})} + \epsilon$$

and

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \le \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ ,  $(I(X; \hat{X}), D)$  is achievable.

1. For sufficiently large n, consider

 $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$   $= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$   $+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$   $\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$   $= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$ 

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le \underline{I(X;\hat{X})} + \epsilon$$

 $\operatorname{and}$ 

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > \underline{D} + \epsilon\} \leq \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq \underline{D}$ ,  $(I(X; \hat{X}), \underline{D})$  is achievable.

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

 $\operatorname{and}$ 

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ ,  $(I(X; \hat{X}), D)$  is achievable.

6. Finally, minimize  $I(X; \hat{X})$  over all such  $\hat{X}$  to conclude that  $(R_I(D), D)$  is achievable, i.e.,

1. For sufficiently large n, consider

$$\begin{aligned} \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \\ &= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\} \\ &+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\} \\ &\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1 \\ &= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}. \end{aligned}$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

 $\operatorname{and}$ 

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ ,  $(I(X; \hat{X}), D)$  is achievable.

6. Finally, minimize  $I(X; \hat{X})$  over all such  $\hat{X}$  to conclude that  $(R_I(D), D)$  is achievable, i.e.,

1. For sufficiently large n, consider

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$$

$$= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$$

$$+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$$

$$\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$$

$$= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max}\delta$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

 $\operatorname{and}$ 

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ ,  $(I(X; \hat{X}), D)$  is achievable.

6. Finally, minimize  $I(X; \hat{X})$  over all such  $\hat{X}$  to conclude that  $(R_I(D), \underline{D})$  is achievable, i.e.,
## The Remaining Details

1. For sufficiently large n, consider

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\}$$

$$= \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K = 1\} \Pr\{K = 1\}$$

$$+ \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \Pr\{K \neq 1\}$$

$$\leq 1 \cdot \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} \cdot 1$$

$$= \epsilon + \Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\}.$$

2. Recall that conditioning on  $\{K \neq 1\}, (\mathbf{X}, \hat{\mathbf{X}}) \in T^n_{[X\hat{X}]\delta}$ . Then

$$d(\mathbf{X}, \hat{\mathbf{X}}) \le D + d_{max}\delta.$$

3. By taking 
$$\delta \leq \frac{\epsilon}{d_{max}}$$
, we obtain

$$d(\mathbf{X}, \hat{\mathbf{X}}) \leq D + d_{max} \left(\frac{\epsilon}{d_{max}}\right) = D + \epsilon.$$

**Proposition** For  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ , if  $(\mathbf{x}, \hat{\mathbf{x}}) \in T^n_{[X\hat{X}]\delta}$ , then  $d(\mathbf{x}, \hat{\mathbf{x}}) \leq D + d_{max}\delta.$  4. Therefore,  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon | K \neq 1\} = 0$ , which implies  $\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \leq \epsilon$ .

5. Thus we have shown the existence of a code such that

$$\frac{1}{n}\log M \le I(X;\hat{X}) + \epsilon$$

 $\operatorname{and}$ 

$$\Pr\{d(\mathbf{X}, \hat{\mathbf{X}}) > D + \epsilon\} \le \epsilon.$$

Hence, for any  $\hat{X}$  such that  $Ed(X, \hat{X}) \leq D$ ,  $(I(X; \hat{X}), D)$  is achievable.

6. Finally, minimize  $I(X; \hat{X})$  over all such  $\hat{X}$  to conclude that  $(R_I(D), \underline{D})$  is achievable, i.e.,

$$R_I(D) \ge R(D).$$



$$\frac{2^{nH(X)}}{2^{nH(X|\hat{X})}} \approx 2^{nI(X;\hat{X})} \ge 2^{nR_I(D)}$$



$$\frac{2^{nH(X)}}{2^{nH(X|\hat{X})}} \approx 2^{nI(X;\hat{X})} \ge 2^{nR_I(D)}$$



$$\frac{2^{nH(X)}}{2^{nH(X|\hat{X})}} \approx 2^{nI(X;\hat{X})} \ge 2^{nR_I(D)}.$$



$$\frac{2^{nH(X)}}{2^{nH(X|\hat{X})}} \approx 2^{nI(X;\hat{X})} \ge 2^{nR_I(D)}.$$



$$\frac{2^{nH(X)}}{2^{nH(X|\hat{X})}} \approx 2^{nI(X;\hat{X})} \ge 2^{nR_I(D)}.$$



$$\frac{2^{nH(X)}}{2^{nH(X|\hat{X})}} \approx \frac{2^{nI(X;\hat{X})}}{2^{nH(X|\hat{X})}} \ge 2^{nR_I(D)}.$$