

# 7.4 Achievability

• Consider a DMC p(y|x).

- Consider a DMC p(y|x).
- For every input distribution p(x), prove that the rate I(X;Y) is achievable by showing for large n the existence of a channel code such that
  - 1. the rate of the code is arbitrarily close to I(X;Y);
  - 2. the maximal probability of error  $\lambda_{max}$  is arbitrarily small.

- Consider a DMC p(y|x).
- For every input distribution p(x), prove that the rate I(X;Y) is achievable by showing for large n the existence of a channel code such that
  - 1. the rate of the code is arbitrarily close to I(X;Y);
  - 2. the maximal probability of error  $\lambda_{max}$  is arbitrarily small.
- Choose the input distribution p(x) to be one that achieves the channel capacity, i.e., I(X;Y) = C.

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\boldsymbol{\delta}}\} \le 2^{-n(I(X;Y)-\boldsymbol{\tau})},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

 $\mathbf{Proof}$ 

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

 $\mathbf{Proof}$ 

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ . 5. Then we have

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y})$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ . 5. Then we have

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y})$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \leq 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

 $p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$ 

 $\operatorname{and}$ 

$$p(\mathbf{y}) \leq 2^{-n(H(Y)-\zeta)}$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

 $p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$ 

 $\operatorname{and}$ 

$$p(\mathbf{y}) \leq 2^{-n(H(Y)-\zeta)}$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

 $p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$ 

 $\operatorname{and}$ 

$$p(\mathbf{y}) \leq 2^{-n(H(Y)-\zeta)}$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T^n_{[XY]\delta}| \le 2^{n(H(X,Y)+\xi)}$$

where  $\xi \to 0$  as  $\delta \to 0$ . 5. Then we have

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

 $p(\mathbf{x}) \leq 2^{-n(H(X)-\eta)}$ 

 $\operatorname{and}$ 

$$p(\mathbf{y}) \leq 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T^n_{[XY]\delta}| \le 2^{n(H(X,Y)+\xi)}$$

where  $\xi \to 0$  as  $\delta \to 0$ . 5. Then we have

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### Proof

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{x}) p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$= 2^{-n(H(X) + H(Y) - H(X, Y) - \xi - \eta - \zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T_{[XY]\delta}^n\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T_{[XY]\delta}^n} p(\mathbf{x}) p(\mathbf{y})$$

$$\leq 2^{n(H(X,Y)+\xi)} \cdot 2^{-n(H(X)-\eta)} \cdot 2^{-n(H(Y)-\zeta)}$$

$$= 2^{-n(\underline{H}(X)+H(Y)-H(X,Y)} - \xi - \eta - \zeta)$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y})$$

$$\leq 2^{n(H(X,Y)+\xi)} \cdot 2^{-n(H(X)-\eta)} \cdot 2^{-n(H(Y)-\zeta)}$$

$$= 2^{-n(H(X)+H(Y)-H(X,Y)-\xi-\eta-\zeta)}$$

$$= 2^{-n(I(X;Y)-\xi-\eta-\zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\}$$

$$= \sum_{\substack{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}}} p(\mathbf{x})p(\mathbf{y})$$

$$\leq 2^{n(H(X,Y)+\xi)} \cdot 2^{-n(H(X)-\eta)} \cdot 2^{-n(H(Y)-\zeta)}$$

$$= 2^{-n(H(X)+H(Y)-H(X,Y)-\xi-\eta-\zeta)}$$

$$= 2^{-n(I(X;Y)-\xi-\eta-\zeta)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### Proof

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$= 2^{-n(H(X) + H(Y) - H(X, Y) - \xi - \eta - \zeta)}$$

$$= 2^{-n(I(X; Y) - \xi - \eta - \zeta)}$$

$$= 2^{-n(I(X; Y) - \tau)},$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)},$$

where  $\tau \to 0$  as  $\delta \to 0$ .

#### $\mathbf{Proof}$

1. Consider

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^n_{[XY]\delta}\} = \sum_{(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y}).$$

2. By the consistency of strong typicality, for  $(\mathbf{x}, \mathbf{y}) \in T^n_{[XY]\delta}$ ,  $\mathbf{x} \in T^n_{[X]\delta}$  and  $\mathbf{y} \in T^n_{[Y]\delta}$ .

3. By the strong AEP, all the  $p(\mathbf{x})$  and  $p(\mathbf{y})$  in the above summation satisfy

$$p(\mathbf{x}) \le 2^{-n(H(X) - \eta)}$$

and

$$p(\mathbf{y}) \le 2^{-n(H(Y)-\zeta)},$$

where  $\eta, \zeta \to 0$  as  $\delta \to 0$ .

4. By the strong JAEP,

$$|T_{[XY]\delta}^n| \le 2^{n(H(X,Y)+\xi)},$$

where  $\xi \to 0$  as  $\delta \to 0$ .

5. Then we have

$$\Pr\{(\mathbf{X}', \mathbf{Y}') \in T^{n}_{[XY]\delta}\}$$

$$= \sum_{(\mathbf{x}, \mathbf{y}) \in T^{n}_{[XY]\delta}} p(\mathbf{x})p(\mathbf{y})$$

$$\leq 2^{n(H(X, Y) + \xi)} \cdot 2^{-n(H(X) - \eta)} \cdot 2^{-n(H(Y) - \zeta)}$$

$$= 2^{-n(H(X) + H(Y) - H(X, Y) - \xi - \eta - \zeta)}$$

$$= 2^{-n(I(X; Y) - \xi - \eta - \zeta)}$$

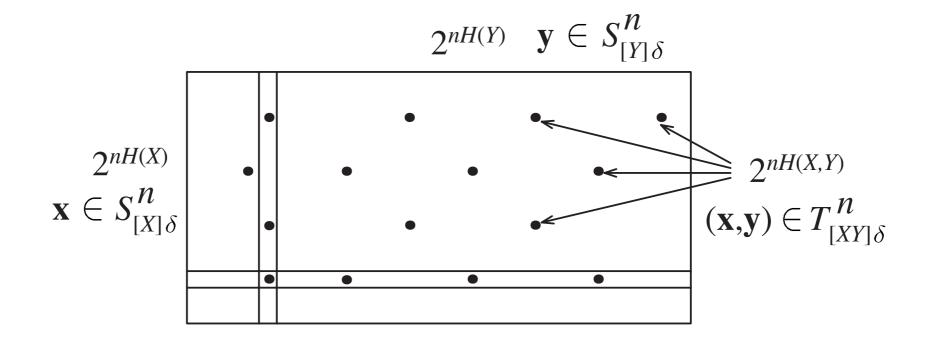
$$= 2^{-n(I(X; Y) - \tau)},$$

where

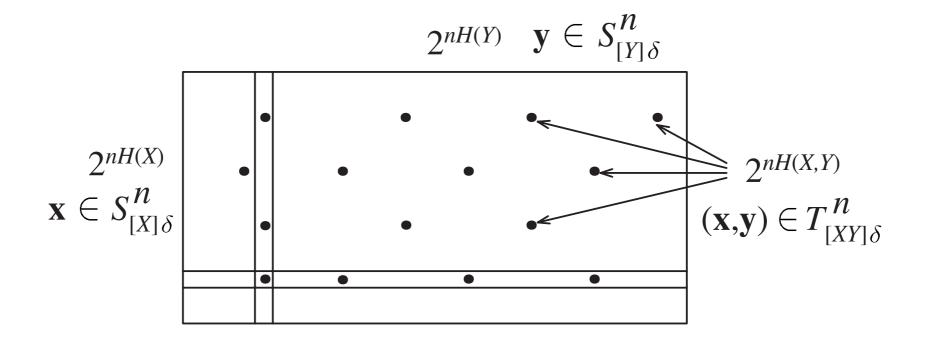
$$\tau = \xi + \eta + \zeta \to 0$$

as  $\delta \to 0$ . The lemma is proved.

### An Interpretation of Lemma 7.17

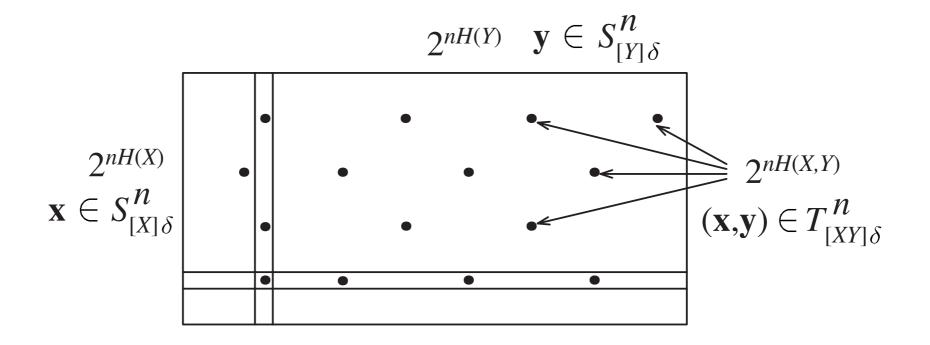


### An Interpretation of Lemma 7.17



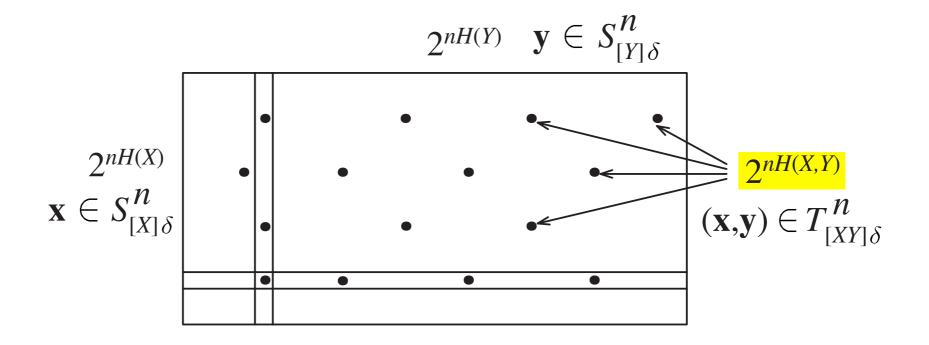
• Randomly choose a row with uniform distribution and randomly choose a column with uniform distribution.

### An Interpretation of Lemma 7.17



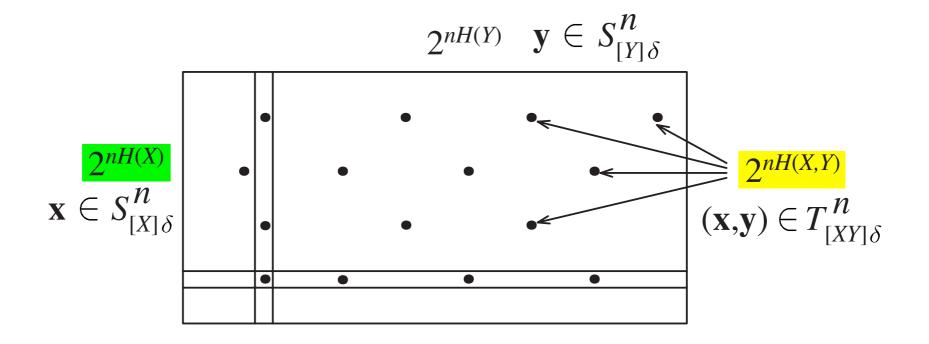
- Randomly choose a row with uniform distribution and randomly choose a column with uniform distribution.
- Then

 $\Pr\{\text{obtaining a jointly typical pair}\} \approx \frac{2^{nH(X,Y)}}{2^{nH(X)}2^{nH(Y)}} = 2^{-nI(X;Y)}$ 



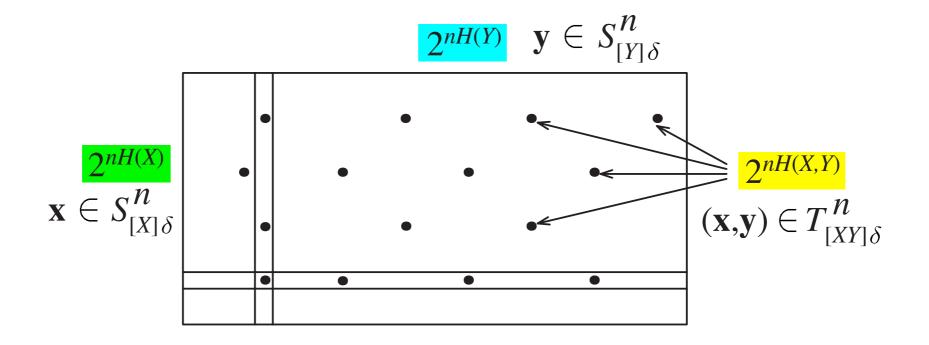
- Randomly choose a row with uniform distribution and randomly choose a column with uniform distribution.
- Then

 $\Pr\{\text{obtaining a jointly typical pair}\} \approx \frac{2^{nH(X,Y)}}{2^{nH(X)}2^{nH(Y)}} = 2^{-nI(X;Y)}$ 



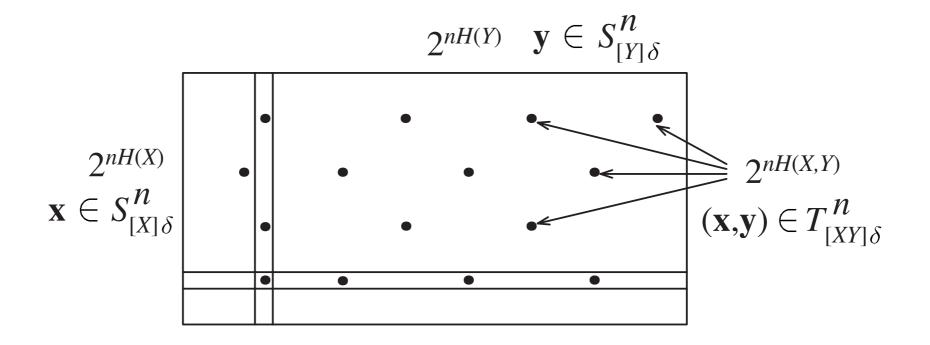
- Randomly choose a row with uniform distribution and randomly choose a column with uniform distribution.
- Then

 $\Pr\{\text{obtaining a jointly typical pair}\} \approx \frac{2^{nH(X,Y)}}{2^{nH(X)}2^{nH(Y)}} = 2^{-nI(X;Y)}$ 



- Randomly choose a row with uniform distribution and randomly choose a column with uniform distribution.
- Then

$$\Pr\{\text{obtaining a jointly typical pair}\} \approx \frac{2^{nH(X,Y)}}{2^{nH(X)}2^{nH(Y)}} = 2^{-nI(X;Y)}$$



- Randomly choose a row with uniform distribution and randomly choose a column with uniform distribution.
- Then

 $\Pr\{\text{obtaining a jointly typical pair}\} \approx \frac{2^{nH(X,Y)}}{2^{nH(X)}2^{nH(Y)}} = 2^{-nI(X;Y)}$ 

Parameter Settings

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

The Random Coding Scheme

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

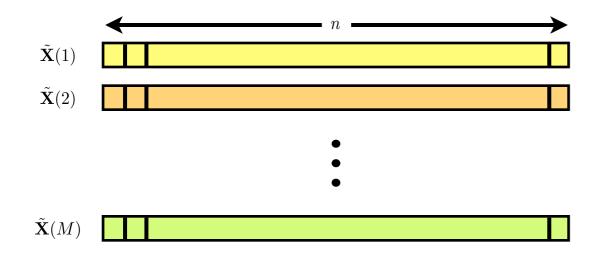
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

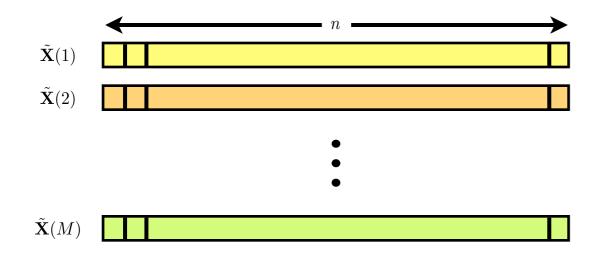
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

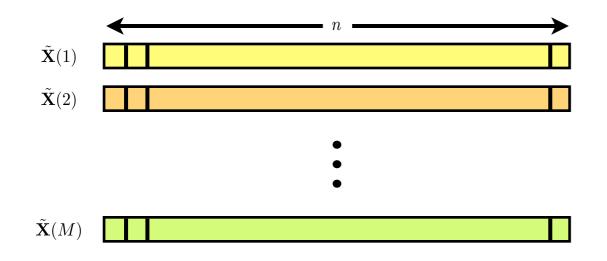
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

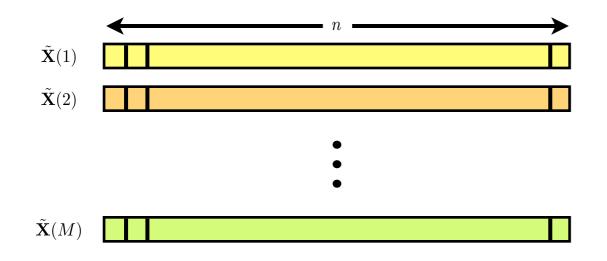
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

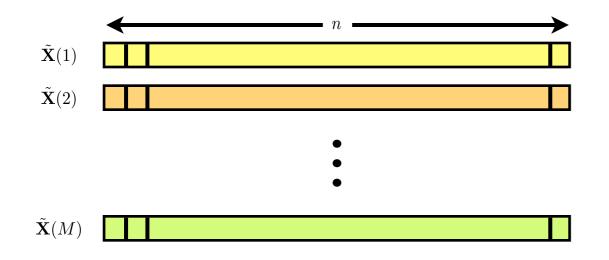
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

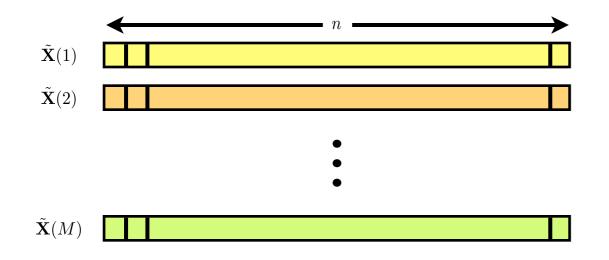
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

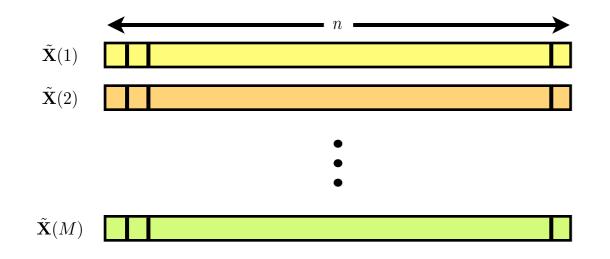
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

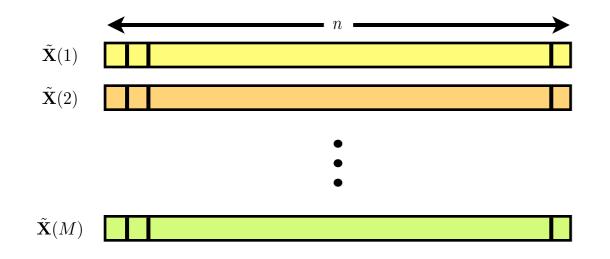
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

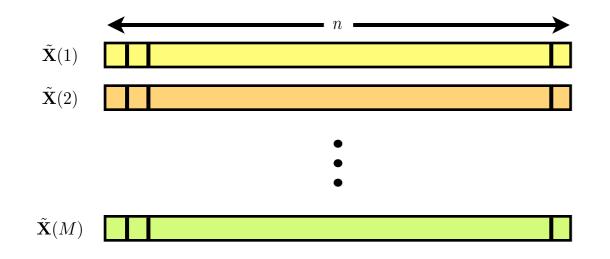
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

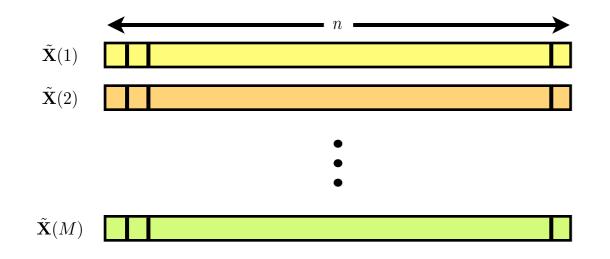
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

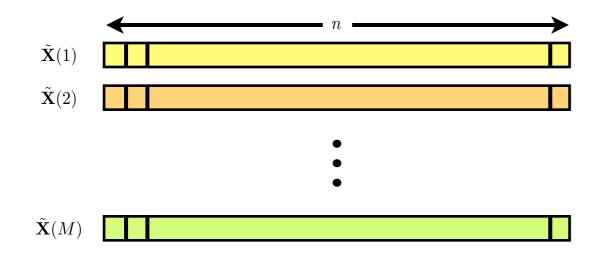
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

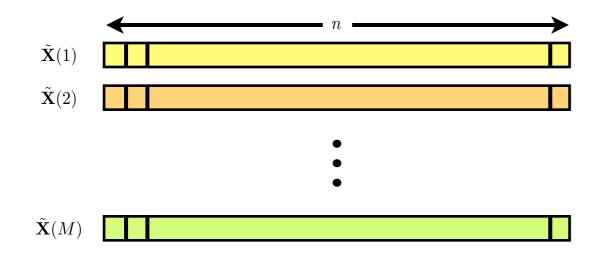
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

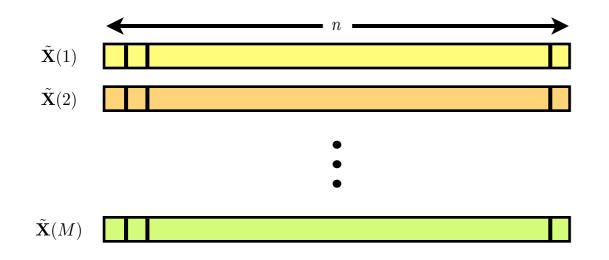
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

4. Transmit  $\mathbf{X} = \tilde{\mathbf{X}}(W)$  through the channel.

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

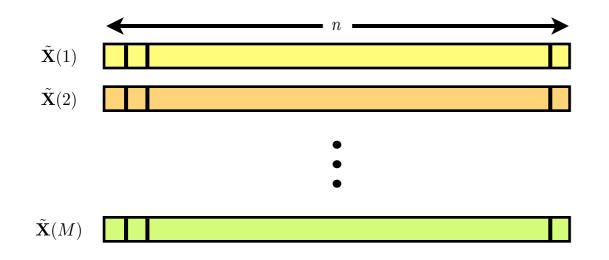
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

4. Transmit  $\mathbf{X} = \tilde{\mathbf{X}}(W)$  through the channel.

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

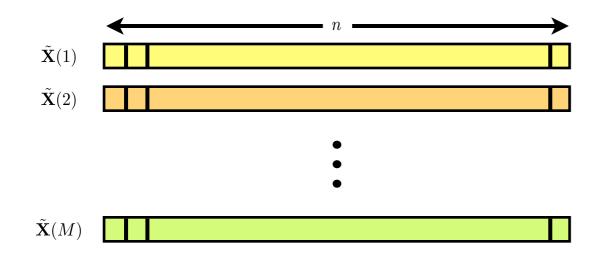
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

- 4. Transmit  $\mathbf{X} = \tilde{\mathbf{X}}(W)$  through the channel.
- 5. The channel outputs a sequence  $\mathbf{Y}$  according to

$$\Pr{\mathbf{Y} = \mathbf{y} | \tilde{\mathbf{X}}(W) = \mathbf{x}} = \prod_{i=1}^{n} p(y_i | x_i).$$

#### **Parameter Settings**

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

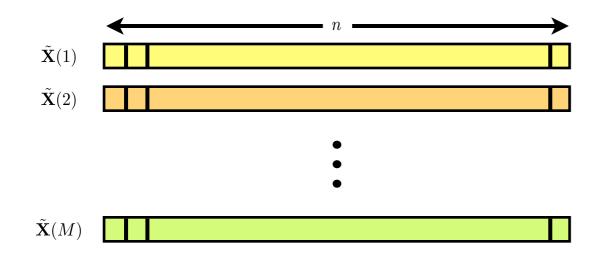
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

- 4. Transmit  $\mathbf{X} = \tilde{\mathbf{X}}(W)$  through the channel.
- 5. The channel outputs a sequence  $\mathbf{Y}$  according to

$$\Pr{\mathbf{Y} = \mathbf{y} | \tilde{\mathbf{X}}(W) = \mathbf{x}} = \prod_{i=1}^{n} p(y_i | x_i).$$

#### Parameter Settings

1. Fix  $\epsilon > 0$  and input distribution p(x). Let  $\delta$  to be specified later.

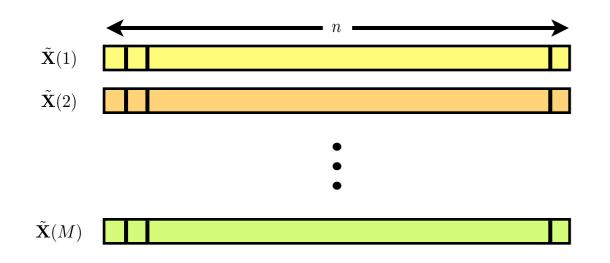
2. Let M be an even integer satisfying

$$I(X;Y) - \frac{\epsilon}{2} < \frac{1}{n} \log M < I(X;Y) - \frac{\epsilon}{4},$$

where n is sufficiently large, i.e.,  $M \approx 2^{nI(X;Y)}$ .

#### The Random Coding Scheme

1. Construct the codebook  $\mathcal{C}$  of an (n, M) code by generating M codewords in  $\mathcal{X}^n$  independently and identically according to  $p(x)^n$ . Denote these codewords by  $\tilde{\mathbf{X}}(1), \tilde{\mathbf{X}}(2), \cdots, \tilde{\mathbf{X}}(M)$ .



• Generate each component according to p(x).

 $\bullet$  There are a total of  $|\mathcal{X}|^{Mn}$  possible codebooks that can be constructed.

• Regard two codebooks whose sets of codewords are permutations of each other as two different codebooks.

2. Reveal the codebook  ${\mathcal C}$  to both the encoder and the decoder.

3. A message W is chosen from  $\mathcal{W}$  according to the uniform distribution.

- 4. Transmit  $\mathbf{X} = \tilde{\mathbf{X}}(W)$  through the channel.
- 5. The channel outputs a sequence  $\mathbf{Y}$  according to

$$\Pr{\mathbf{Y} = \mathbf{y} | \tilde{\mathbf{X}}(W) = \mathbf{x}} = \prod_{i=1}^{n} p(y_i | x_i).$$

6. The sequence  $\mathbf{Y}$  is decoded to the message w if

• 
$$(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta}$$
, and

• there does not exists  $w' \neq w$  such that  $(\tilde{\mathbf{X}}(w'), \mathbf{Y}) \in T^n_{[XY]\delta}$ .

Otherwise,  $\mathbf{Y}$  is decoded to a constant message in  $\mathcal{W}$ . Denote by  $\hat{W}$  the message to which  $\mathbf{Y}$  is decoded.

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

$$\Pr\{Err\} = \sum_{w=1}^{M} \frac{\Pr\{Err|W=w\}}{\Pr\{W=w\}}$$
$$= \frac{\Pr\{Err|W=1\}}{\sum_{w=1}^{M}} \Pr\{W=w\}$$

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

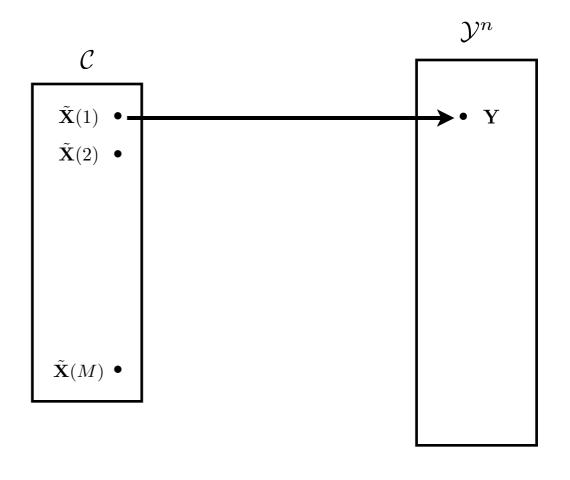
Assume without loss of generality that the message 1 is chosen.

1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

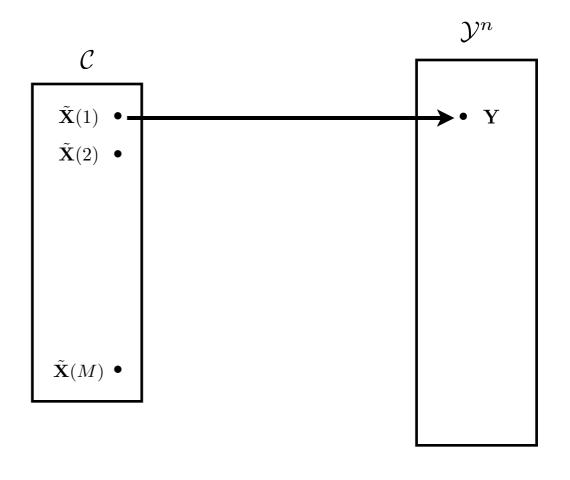


1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

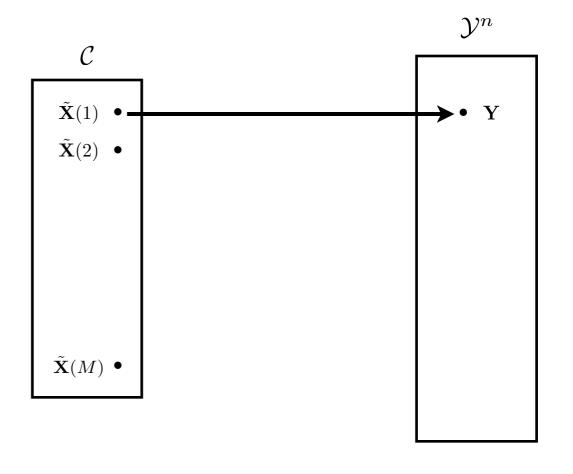
2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

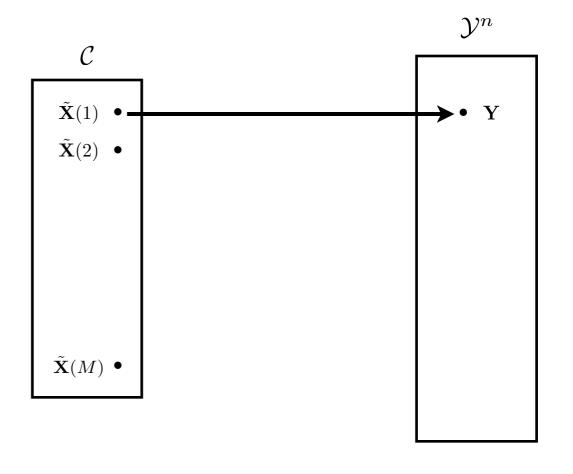
2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

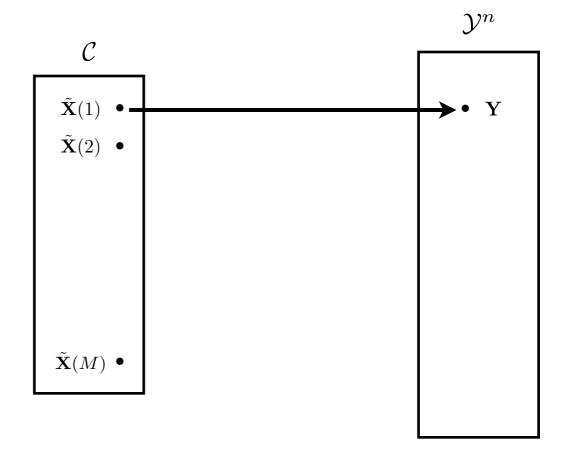
$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

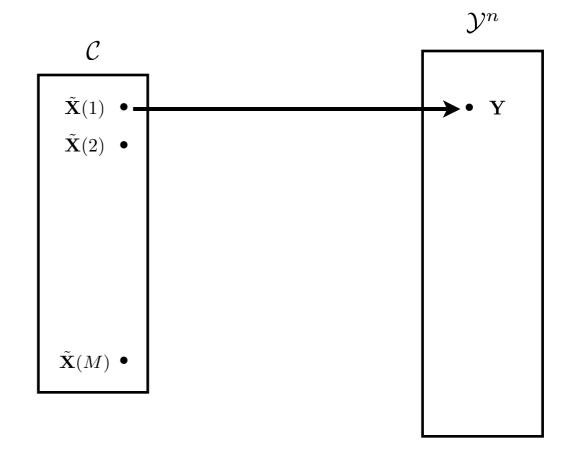
Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{\frac{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

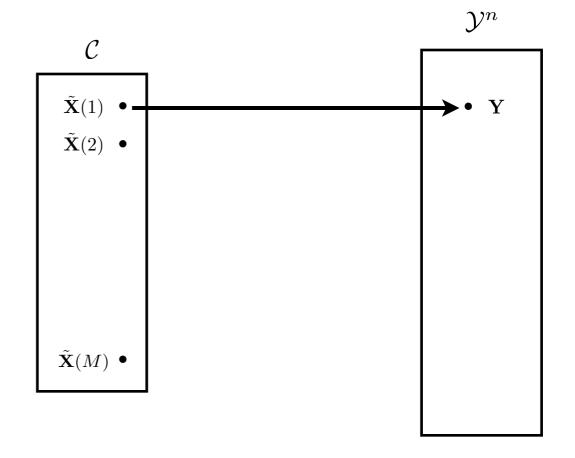
Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

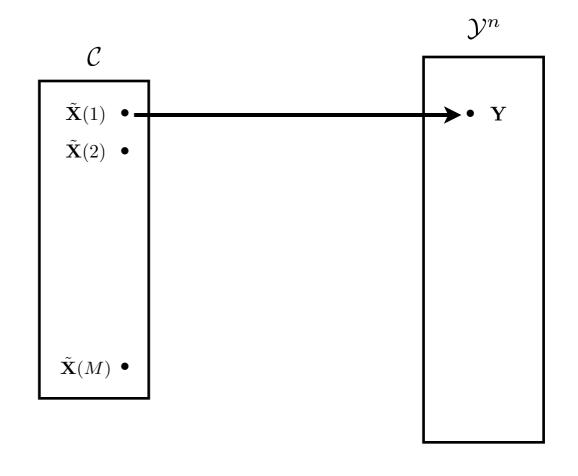
$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

 $\Pr\{Err|W=1\}$ 



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

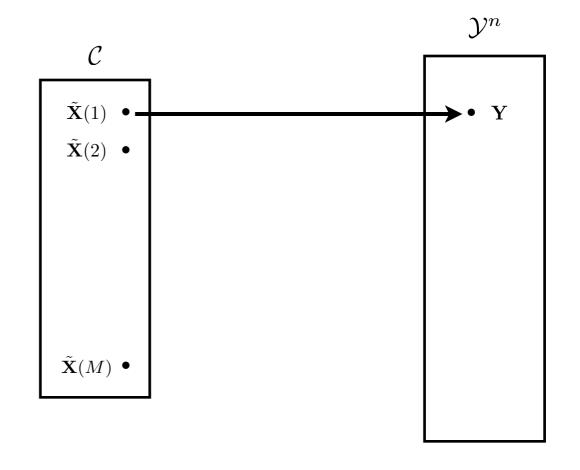
3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$\Pr\{Err|W = 1\}$$
$$= 1 - \Pr\{Err^{c}|W = 1\}$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

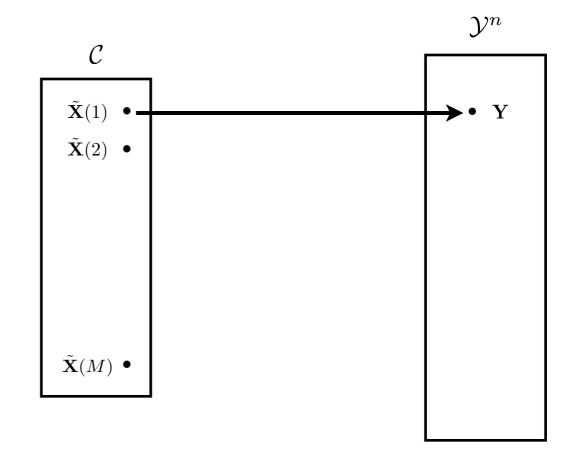
3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$\Pr \{ Err | W = 1 \}$$
$$= 1 - \Pr \{ Err^{c} | W = 1 \}$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

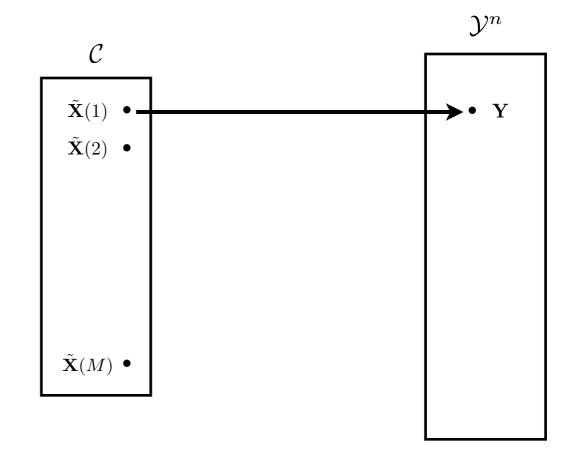
3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\} \le 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

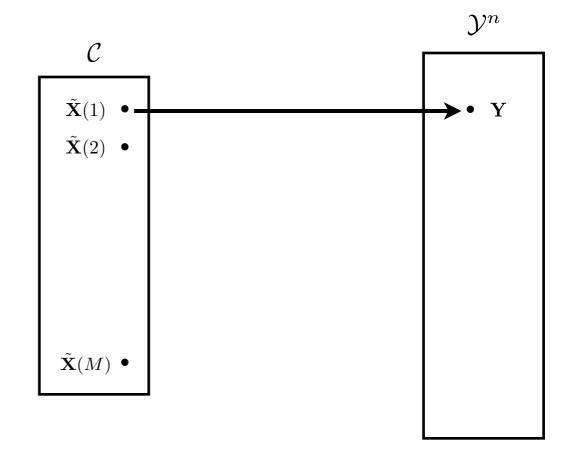
3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$\Pr\{Err|W = 1\} \\ = 1 - \Pr\{Err^{c}|W = 1\} \\ \leq 1 - \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\} \\ = \Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

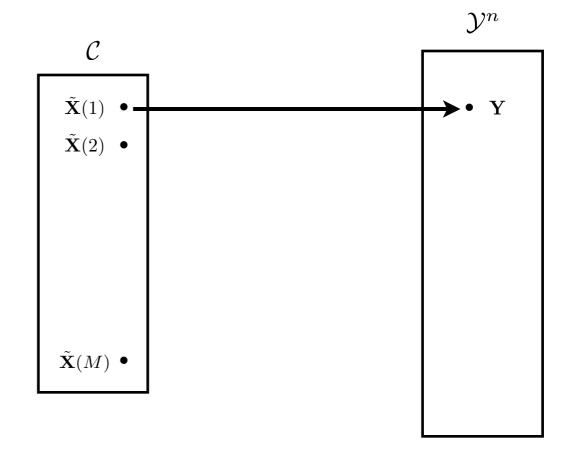
4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$\Pr\{Err|W = 1\} = 1 - \Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= \Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

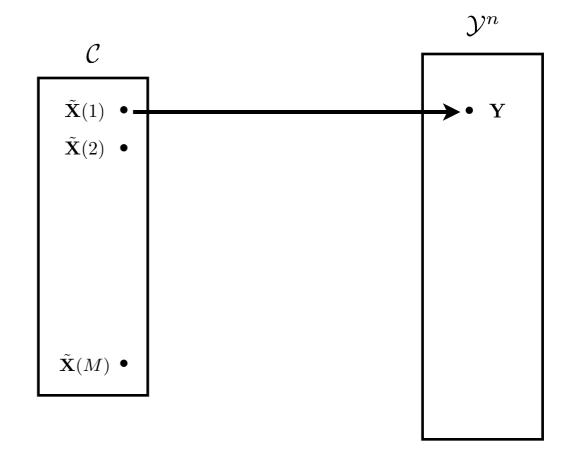
$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$\Pr\{Err|W = 1\} = 1 - \Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= \Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= \Pr\{\underline{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}}|W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

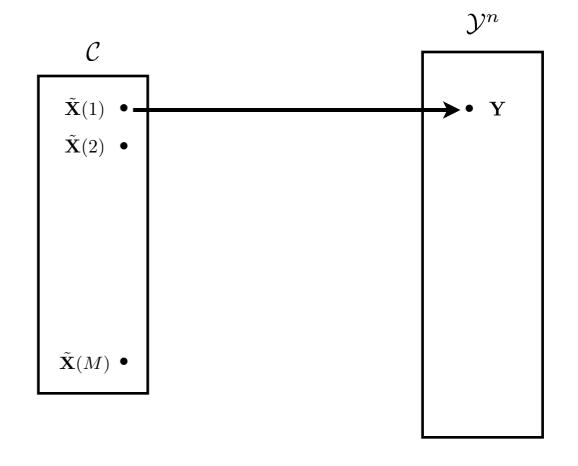
$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

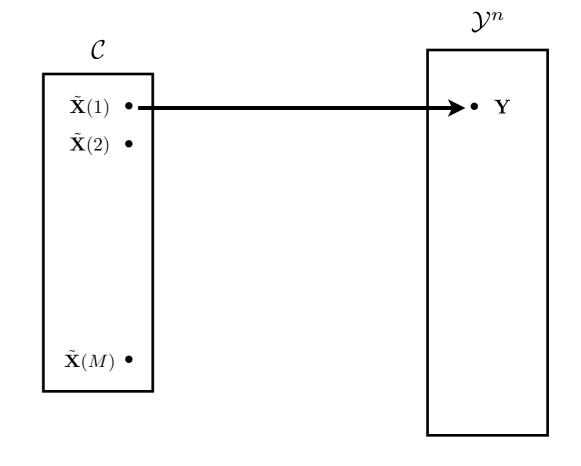
$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

6. By the union bound,

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

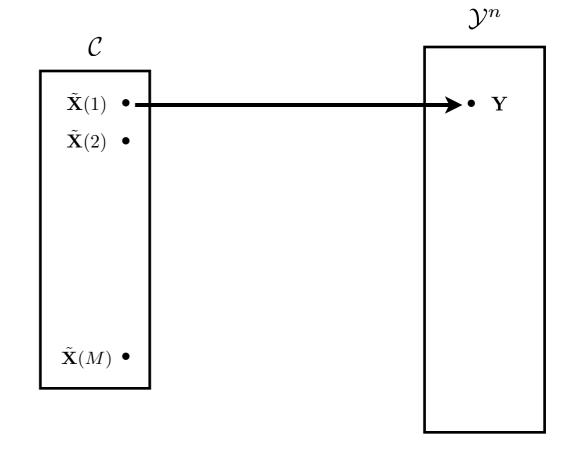
$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{\frac{E_{1}^{c}}{2} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

$$\Pr\{Err|W=1\} \le \Pr\{\frac{E_1^c}{W}|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{\frac{E_w}{|w=1\}} = 1\}$$

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

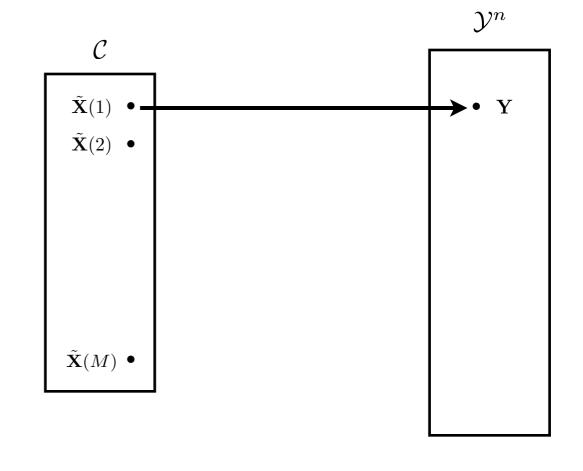
3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$\Pr\{Err|W = 1\} \\ = 1 - \Pr\{Err^{c}|W = 1\} \\ \leq 1 - \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\} \\ = \Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\} \\ = \Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

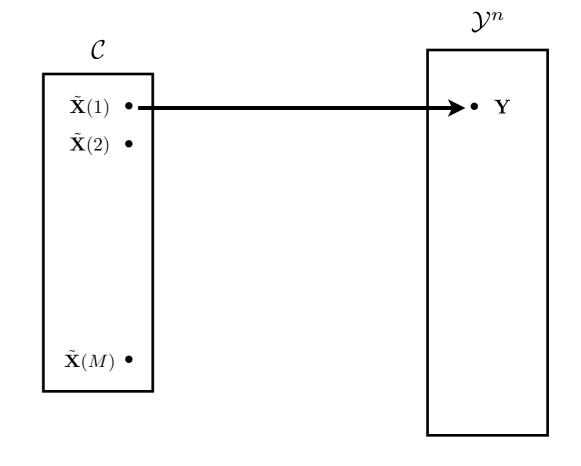
$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

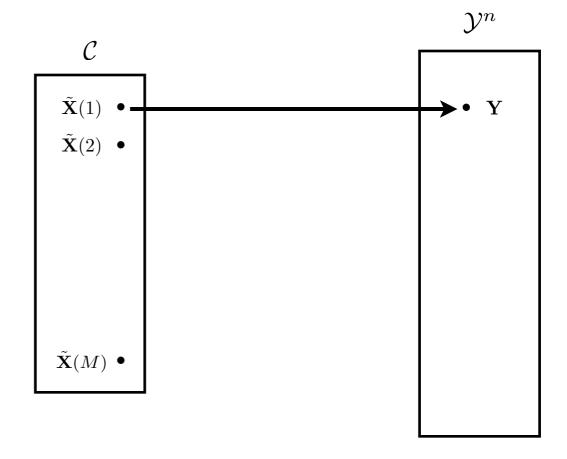
$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

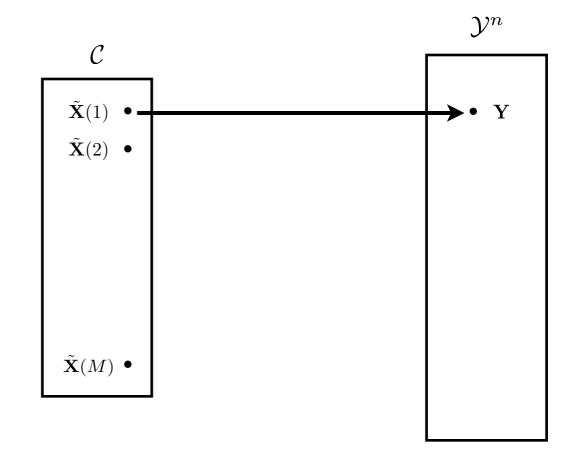
$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \frac{\Pr\{E_1^c|W=1\}}{|W|=1} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr{Err} = \sum_{w=1}^{M} \Pr{Err|W = w} \Pr{W = w}$$
$$= \Pr{Err|W = 1} \sum_{w=1}^{M} \Pr{W = w}$$
$$= \Pr{Err|W = 1}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

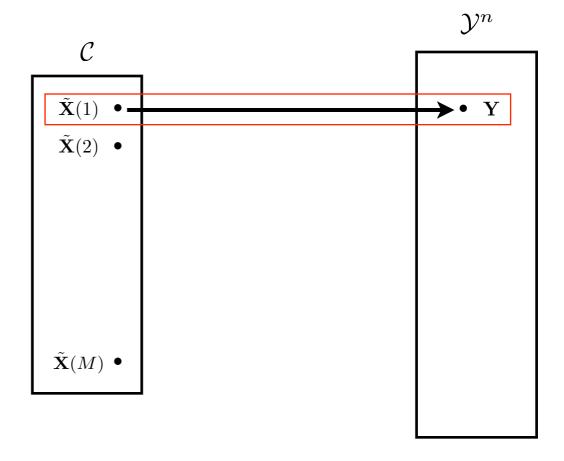
$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \frac{\Pr\{E_1^c|W=1\}}{|W|=1} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

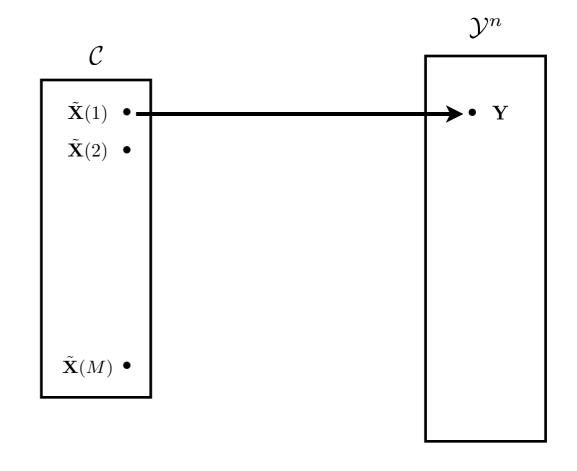
$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

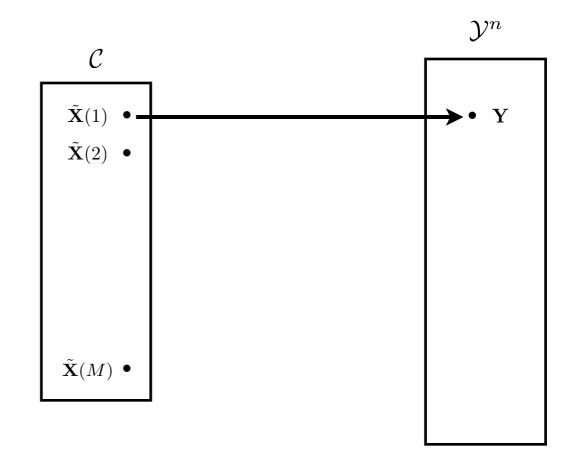
$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$



1. We need to show that  $\Pr{Err} = \Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

2. Consider

$$\Pr\{Err\} = \sum_{w=1}^{M} \Pr\{Err|W = w\} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\} \sum_{w=1}^{M} \Pr\{W = w\}$$
$$= \Pr\{Err|W = 1\}.$$

Assume without loss of generality that the message 1 is chosen.

3. For  $1 \leq w \leq M$ , define the event

$$E_w = \{ (\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} \}.$$

4. If  $E_1$  occurs but  $E_w$  does not occur for all  $2 \le w \le M$ , then no decoding error. Therefore,

$$\Pr\{Err^{c} | W = 1\} \ge \Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \cdots \cap E_{M}^{c} | W = 1\}.$$

5. Consider

$$Pr\{Err|W = 1\} = 1 - Pr\{Err^{c}|W = 1\}$$
  

$$\leq 1 - Pr\{E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c}|W = 1\}$$
  

$$= Pr\{(E_{1} \cap E_{2}^{c} \cap E_{3}^{c} \cap \dots \cap E_{M}^{c})^{c}|W = 1\}$$
  

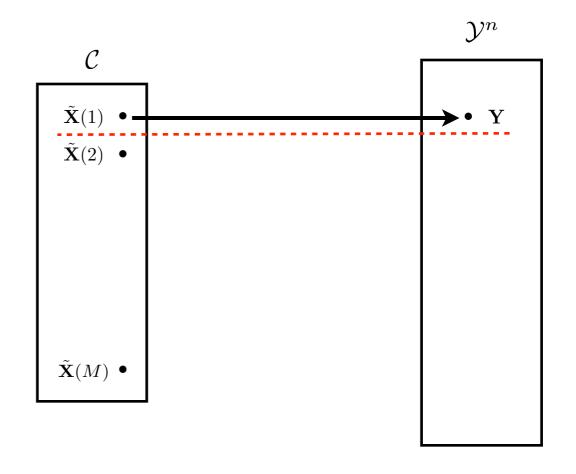
$$= Pr\{E_{1}^{c} \cup E_{2} \cup E_{3} \cup \dots \cup E_{M}|W = 1\}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$



6. By the union bound,

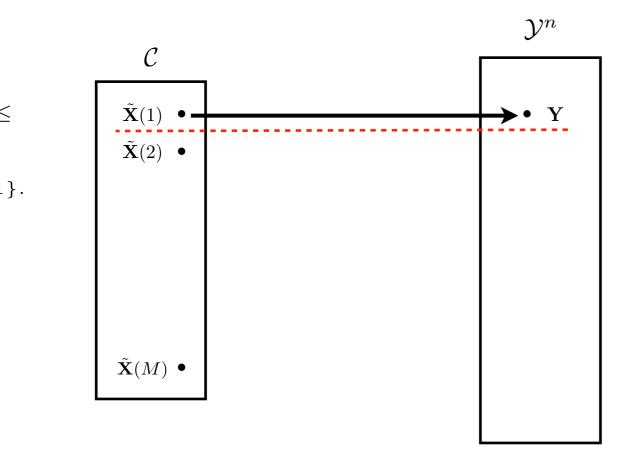
n

1

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

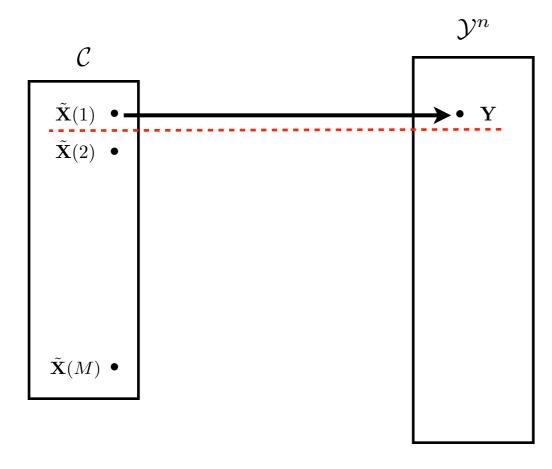


6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$



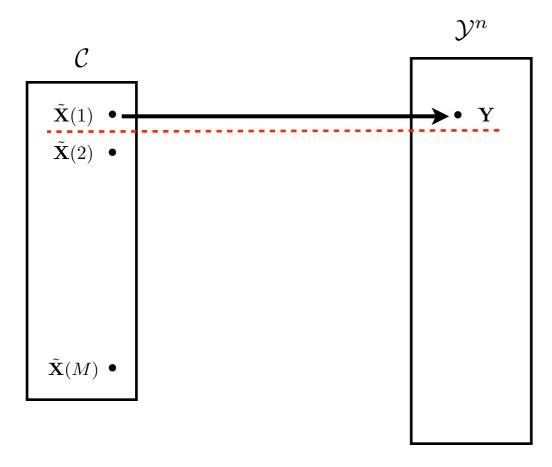
6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .



9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

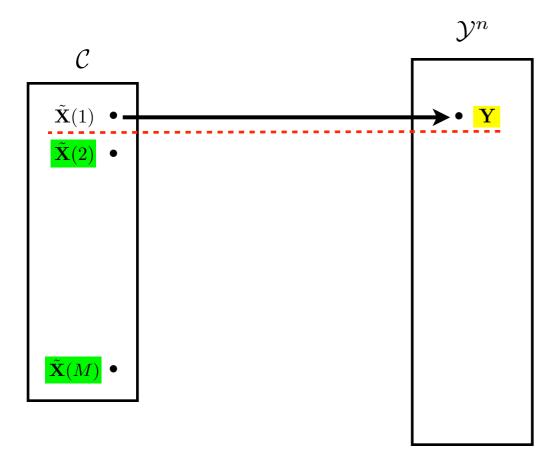
6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .



9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

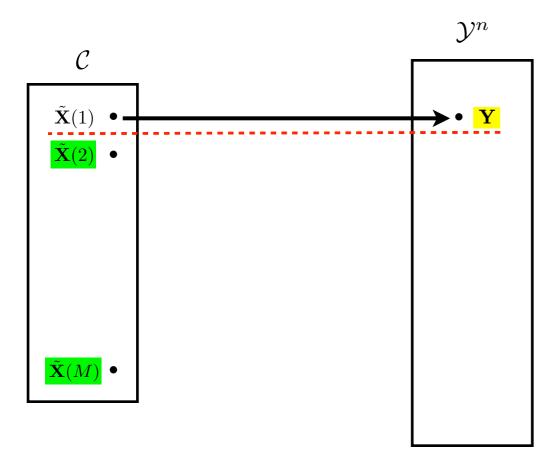
6. By the union bound,

 $\Pr\{Err|W = 1\} \le \Pr\{E_1^c|W = 1\} + \sum_{w=2}^M \Pr\{E_w|W = 1\}$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .



9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

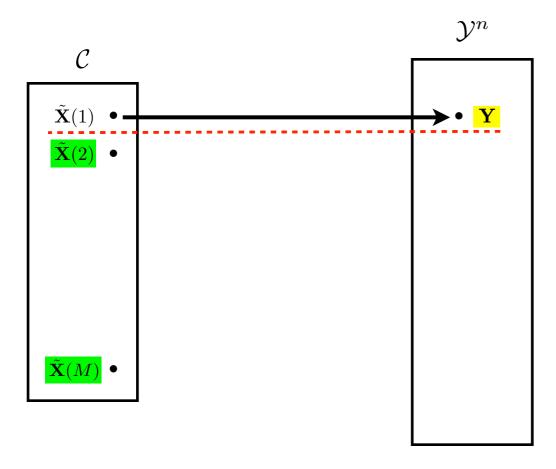
6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .



9. Since a DMC is memoryless, X' and Y' are independent dent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For 
$$2 \leq w \leq M$$
,

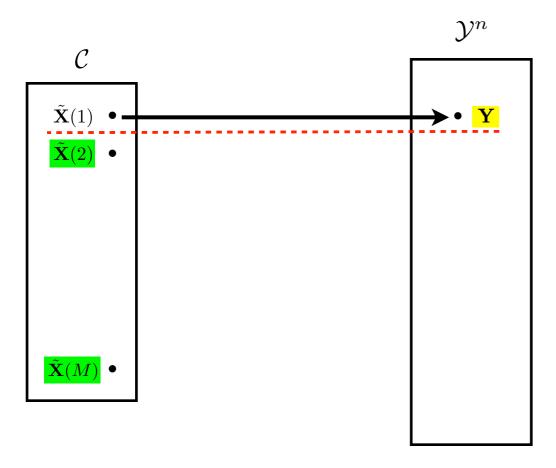
6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .



9. Since a DMC is memoryless, X' and Y' are independent dent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_w | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} | W = 1\}$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

 $\mathcal{C}$ 

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

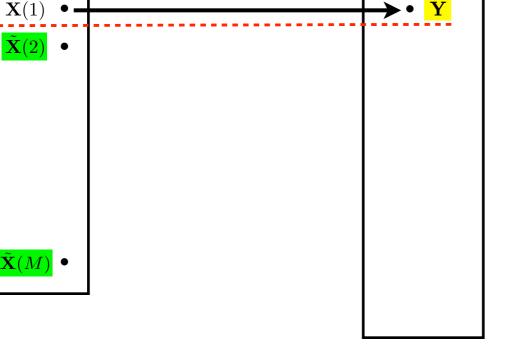
8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

 $\mathcal{Y}^n$ 

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_w | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} | W = 1\}$$



Lemma 7.17

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

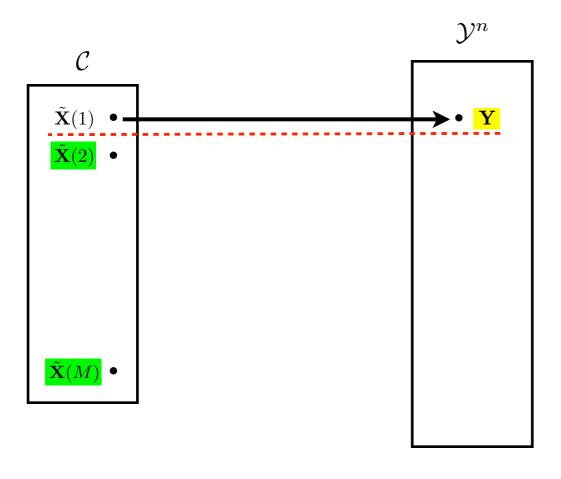
$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$



Lemma 7.17

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

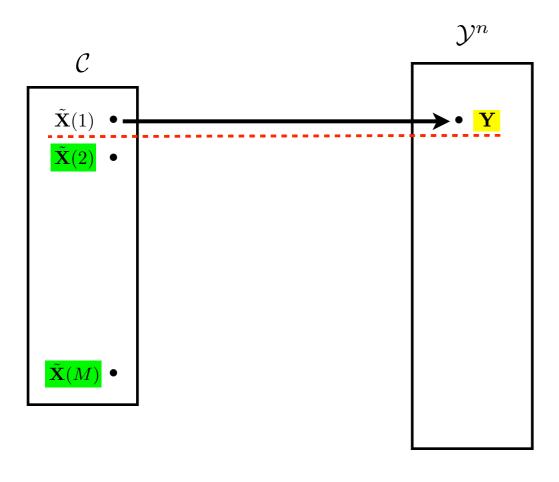
8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .



Lemma 7.17

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

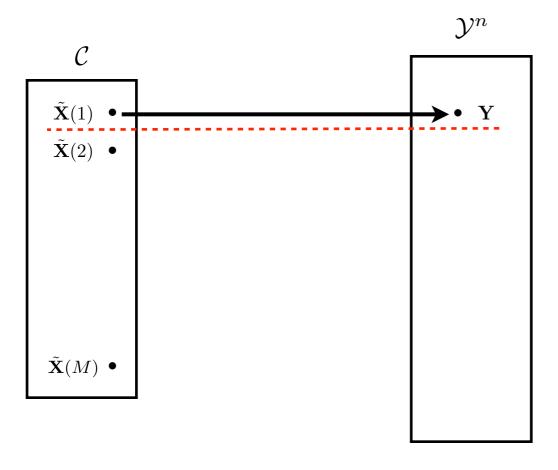
8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .



$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

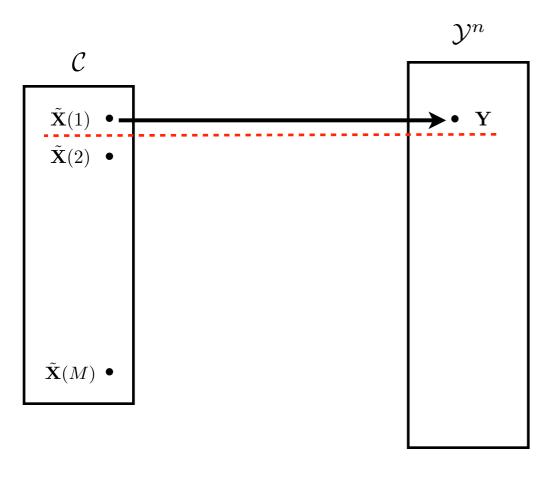
10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

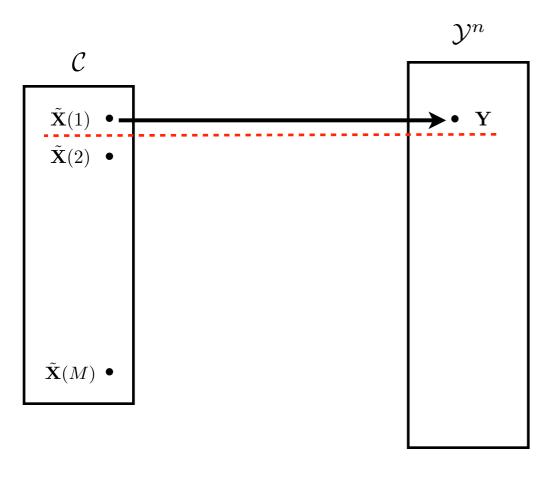
10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

 $\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$ 

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

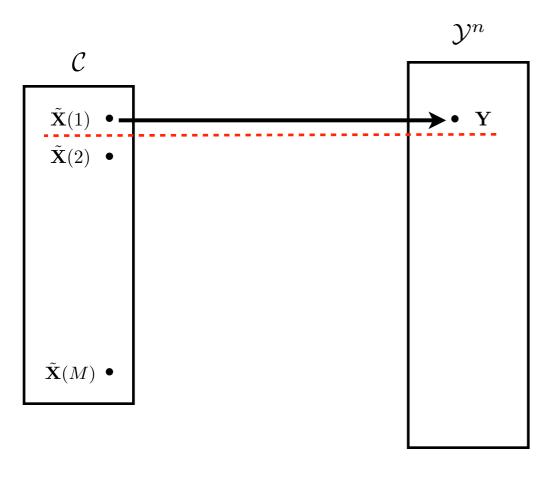
10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$

6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

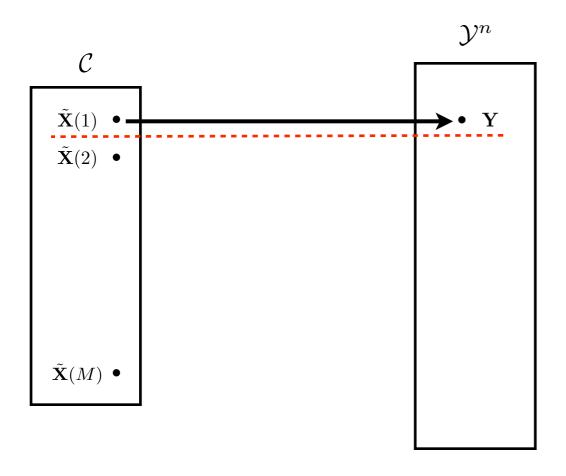
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\frac{\Pr\{Err\}}{\left\{Err\right\}} < \nu + 2^{n\left(I(X;Y) - \frac{\epsilon}{4}\right)} \cdot 2^{-n\left(I(X;Y) - \tau\right)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

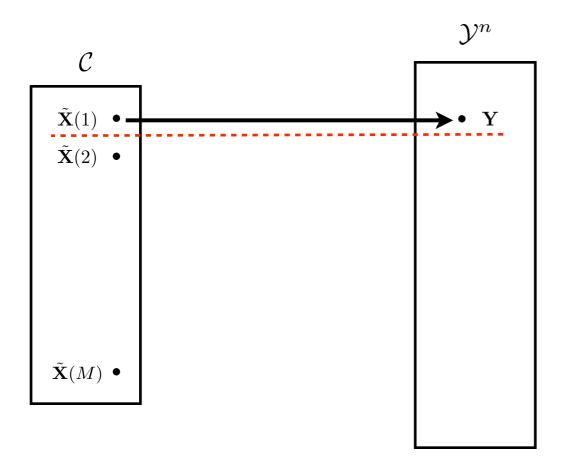
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\frac{\Pr\{Err\}}{\left\{Err\right\}} < \nu + 2^{n\left(I(X;Y) - \frac{\epsilon}{4}\right)} \cdot 2^{-n\left(I(X;Y) - \tau\right)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

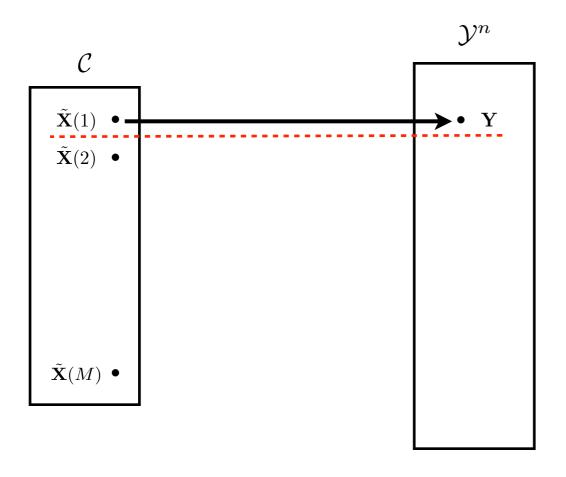
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\frac{\Pr\{Err\}}{\left\{Err\right\}} < \nu + 2^{n\left(I(X;Y) - \frac{\epsilon}{4}\right)} \cdot 2^{-n\left(I(X;Y) - \tau\right)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

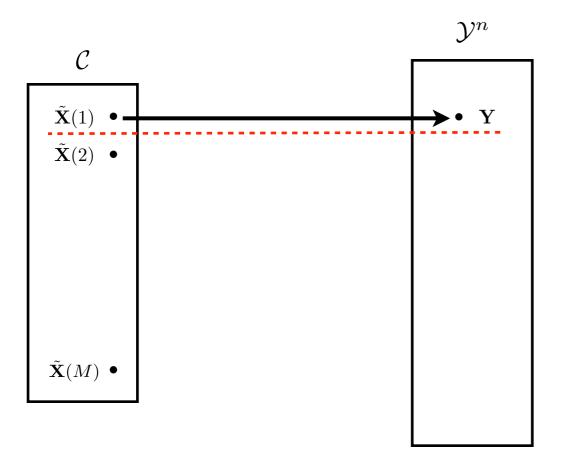
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\frac{\Pr\{Err\}}{\left\{Err\right\}} < \nu + 2^{n\left(I(X;Y) - \frac{\epsilon}{4}\right)} \cdot 2^{-n\left(I(X;Y) - \tau\right)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

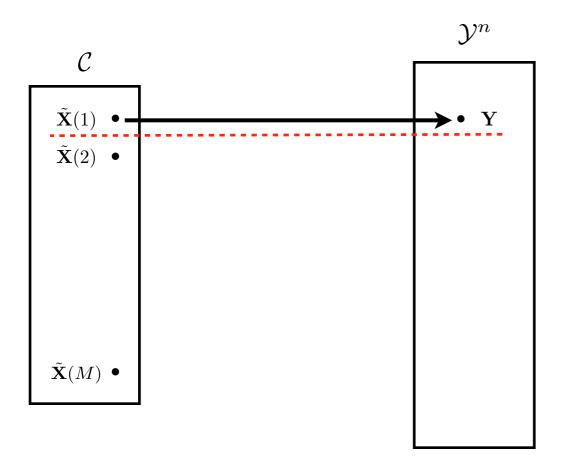
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\frac{\Pr\{Err\}}{\nu} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

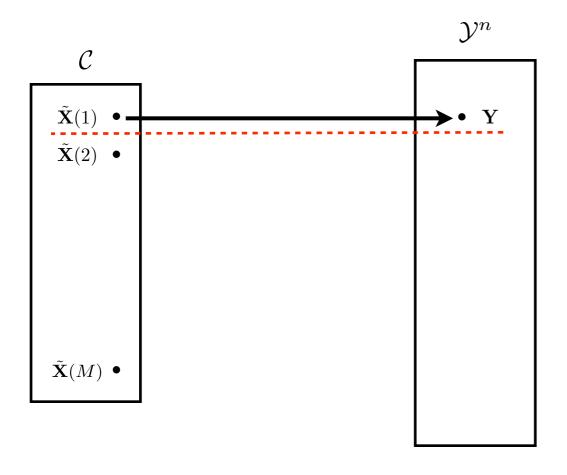
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

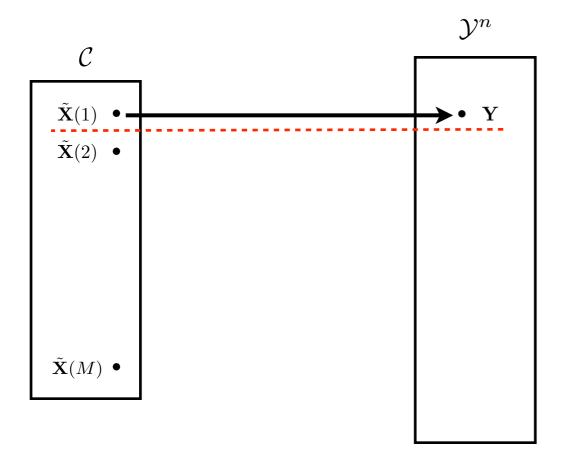
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\frac{\Pr\{Err\}}{\nu} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \le 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

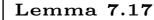
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where 
$$\tau \to 0$$
 as  $\delta \to 0$ 

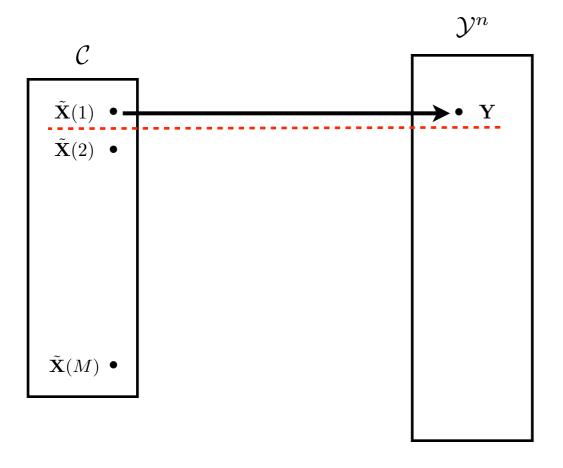
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

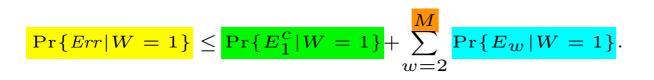
$$\frac{\Pr\{Err\}}{\Pr\{Err\}} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot \frac{2^{-n(I(X;Y) - \tau)}}{2^{-n(I(X;Y) - \tau)}}$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,



7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

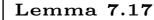
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where 
$$\tau \to 0$$
 as  $\delta \to 0$ 

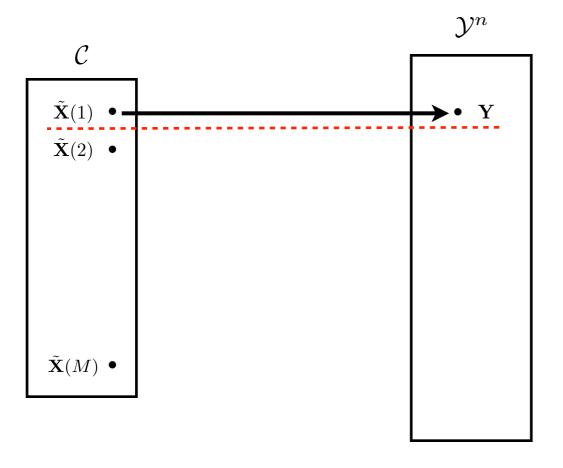
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

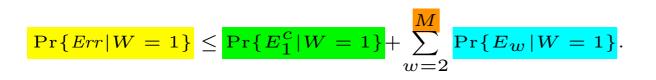
$$\frac{\Pr\{Err\}}{\Pr\{Err\}} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot \frac{2^{-n(I(X;Y) - \tau)}}{2^{-n(I(X;Y) - \tau)}}$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,



7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

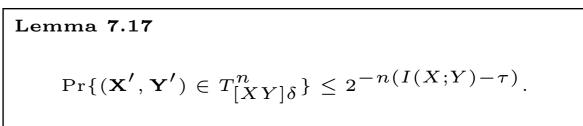
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

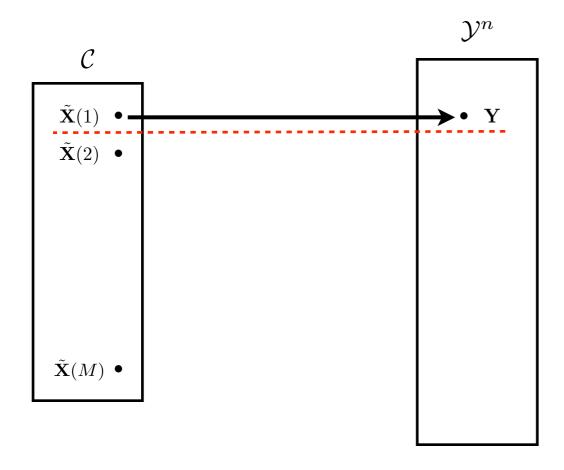
where 
$$\tau \to 0$$
 as  $\delta \to 0$ 

11. Note that

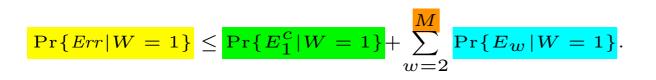
$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}$$

$$\frac{\Pr\{Err\}}{\Pr\{Err\}} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot \frac{2^{-n(I(X;Y) - \tau)}}{2^{-n(I(X;Y) - \tau)}}$$





6. By the union bound,



7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

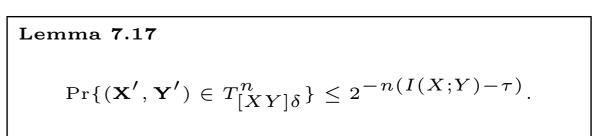
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

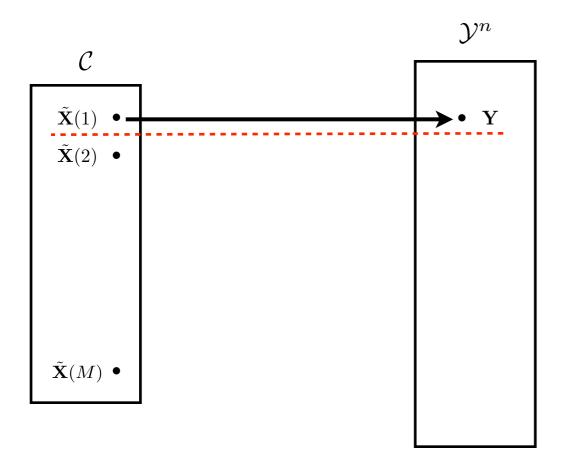
where 
$$\tau \to 0$$
 as  $\delta \to 0$ 

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}$$

$$\frac{\Pr\{Err\}}{\Pr\{Err\}} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot \frac{2^{-n(I(X;Y) - \tau)}}{2^{-n(I(X;Y) - \tau)}}$$





6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

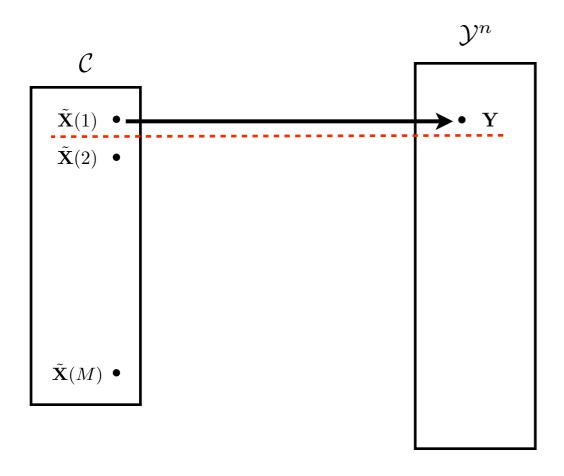
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$

$$\Pr\{(\mathbf{X}',\mathbf{Y}') \in T^n_{[XY]\delta}\} \leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

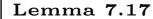
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

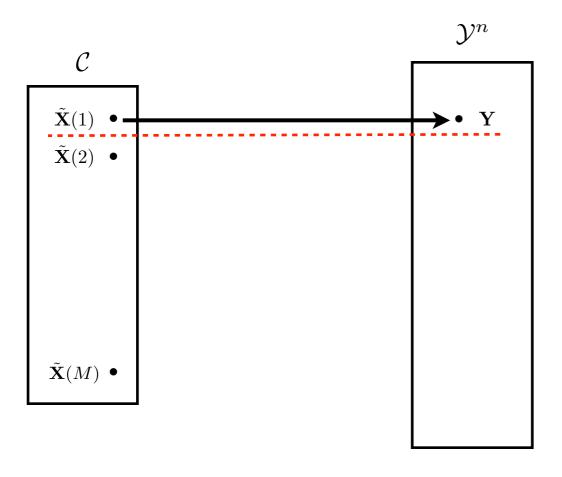
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X,Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

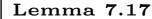
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

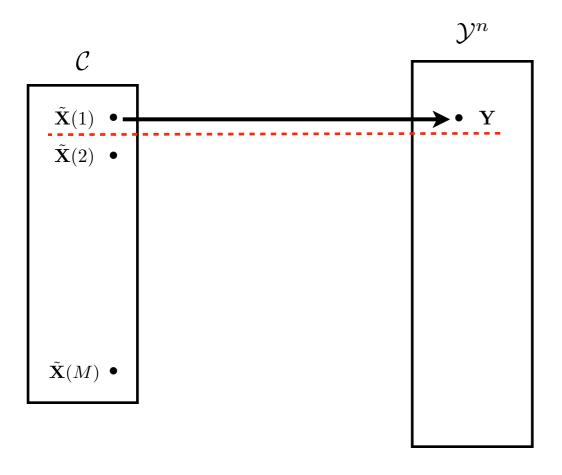
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X,Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X,Y) - \tau)}$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



6. By the union bound,

$$\Pr\{Err|W=1\} \le \Pr\{E_1^c|W=1\} + \sum_{w=2}^M \Pr\{E_w|W=1\}.$$

7. By strong JAEP,

$$\Pr\{E_1^c | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(1), \mathbf{Y}) \notin T_{[XY]\delta}^n | W = 1\} < \nu.$$

8. Conditioning on  $\{W = 1\}$ , for  $2 \leq w \leq M$ ,  $(\tilde{\mathbf{X}}(w), \mathbf{Y})$  are *n* i.i.d. copies of the pair of generic random variables (X', Y'), where  $X' \sim X$  and  $Y' \sim Y$ .

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

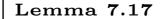
$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

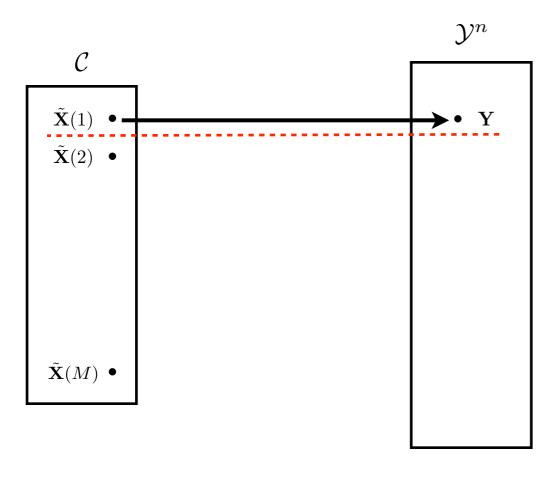
11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X,Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X,Y) - \tau)}$$
$$= \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$



$$\Pr\{(\mathbf{X}',\mathbf{Y}')\in T^n_{[XY]\delta}\}\leq 2^{-n(I(X;Y)-\tau)}.$$



9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

 $\langle \nu.$ 

[, 1-

:1}.

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)} \\ = \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$
$$= \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)} \\ = \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_w | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} | W = 1\}$$
$$\leq 2^{-n(I(X;Y) - \tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)} \\ = \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_w | W = 1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^n_{[XY]\delta} | W = 1\}$$
$$\leq 2^{-n(I(X;Y) - \tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)} \\ = \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

14. Then

$$2^{-n\left(\frac{\epsilon}{4}-\tau\right)}\to 0$$

as  $n \to \infty$ .

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)} \\ = \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

14. Then

$$2^{-n\left(\frac{\epsilon}{4}-\tau\right)}\to 0$$

as  $n \to \infty$ .

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$
$$= \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)} \\ = \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

14. Then

$$2^{-n(\frac{\epsilon}{4}-\tau)} \to 0$$

as  $n \to \infty$ . 15. Let  $\nu < \frac{\epsilon}{3}$  to obtain

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$
$$= \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$
$$< \frac{\epsilon}{3}$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

14. Then

$$2^{-n\left(\frac{\epsilon}{4}-\tau\right)} \to 0$$

as  $n \to \infty$ . 15. Let  $\nu < \frac{\epsilon}{3}$  to obtain

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$
$$= \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$
$$< \frac{\epsilon}{3} \qquad (<\frac{\epsilon}{6})$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

14. Then

$$2^{-n(\frac{\epsilon}{4}-\tau)} \to 0$$

as  $n \to \infty$ . 15. Let  $\nu < \frac{\epsilon}{3}$  to obtain

9. Since a DMC is memoryless, X' and Y' are independent because  $\tilde{\mathbf{X}}(1)$  and  $\tilde{\mathbf{X}}(w)$  are independent and the generation of  $\mathbf{Y}$  depends only on  $\tilde{\mathbf{X}}(1)$ . See textbook for a formal proof.

10. For  $2 \leq w \leq M$ ,

$$\Pr\{E_{w}|W=1\} = \Pr\{(\tilde{\mathbf{X}}(w), \mathbf{Y}) \in T^{n}_{[XY]\delta}|W=1\}$$
$$\leq 2^{-n(I(X;Y)-\tau)}$$

where  $\tau \to 0$  as  $\delta \to 0$ .

11. Note that

$$\frac{1}{n}\log M < I(X;Y) - \frac{\epsilon}{4} \iff M < 2^{n(I(X;Y) - \frac{\epsilon}{4})}.$$

12. Therefore,

$$\Pr\{Err\} < \nu + 2^{n(I(X;Y) - \frac{\epsilon}{4})} \cdot 2^{-n(I(X;Y) - \tau)}$$
$$= \nu + 2^{-n(\frac{\epsilon}{4} - \tau)}.$$
$$< \frac{\epsilon}{3} \qquad (<\frac{\epsilon}{6})$$

13. Recall that  $\epsilon$  is fixed. Since  $\tau \to 0$  as  $\delta \to 0$ , we can choose  $\delta$  to be sufficiently small so that

$$\frac{\epsilon}{4} - \tau > 0.$$

14. Then

$$2^{-n\left(\frac{\epsilon}{4}-\tau\right)} \to 0$$

as  $n \to \infty$ . 15. Let  $\nu < \frac{\epsilon}{3}$  to obtain

$$\Pr\{Err\} < \frac{\epsilon}{2}$$

for sufficiently large n.

• Let n be large.

- Let n be large.
- $\Pr{\{\tilde{\mathbf{X}}(1) \text{ jointly typical with } \mathbf{Y}\}} \to 1.$

- Let n be large.
- $\Pr{\{\tilde{\mathbf{X}}(1) \text{ jointly typical with } \mathbf{Y}\}} \to 1.$
- For  $w \neq 1$ ,  $\Pr{\{\tilde{\mathbf{X}}(w) \text{ jointly typical with } \mathbf{Y}\}} \approx 2^{-nI(X;Y)}$ .

- Let n be large.
- $\Pr{\{\tilde{\mathbf{X}}(1) \text{ jointly typical with } \mathbf{Y}\}} \rightarrow 1.$
- For  $w \neq 1$ ,  $\Pr{\{\tilde{\mathbf{X}}(w) \text{ jointly typical with } \mathbf{Y}\}} \approx 2^{-nI(X;Y)}$ .
- If  $|\mathcal{C}| = M$  grows at a rate  $\langle I(X; Y)$ , then

 $\Pr{\{\tilde{\mathbf{X}}(w) \text{ jointly typical with } \mathbf{Y} \text{ for some } w \neq 1 \}}$ 

can be made arbitrarily small.

- Let n be large.
- $\Pr{\{\tilde{\mathbf{X}}(1) \text{ jointly typical with } \mathbf{Y}\}} \rightarrow 1.$
- For  $w \neq 1$ ,  $\Pr{\{\tilde{\mathbf{X}}(w) \text{ jointly typical with } \mathbf{Y}\}} \approx 2^{-nI(X;Y)}$ .
- If  $|\mathcal{C}| = M$  grows at a rate  $\langle I(X; Y)$ , then

 $\Pr{\{\tilde{\mathbf{X}}(w) \text{ jointly typical with } \mathbf{Y} \text{ for some } w \neq 1 \}}$ 

can be made arbitrarily small.

• Then  $\Pr{\{\hat{W} \neq W\}}$  can be made arbitrarily small.

• According to the random coding scheme,

$$\Pr\{Err\} = \sum_{\mathcal{C}} \Pr\{\mathcal{C}\} \Pr\{Err|\mathcal{C}\}.$$

• According to the random coding scheme,

$$\Pr\{Err\} = \sum_{\mathcal{C}} \Pr\{\mathcal{C}\} \Pr\{Err|\mathcal{C}\}.$$

• Then there exists at least one codebook  $\mathcal{C}^*$  such that

$$P_e = \Pr\{Err|\mathcal{C}^*\} \le \Pr\{Err\} < \frac{\epsilon}{2}.$$

• According to the random coding scheme,

$$\Pr\{Err\} = \sum_{\mathcal{C}} \Pr\{\mathcal{C}\} \Pr\{Err|\mathcal{C}\}.$$

• Then there exists at least one codebook  $\mathcal{C}^*$  such that

$$P_e = \Pr\{Err|\mathcal{C}^*\} \le \Pr\{Err\} < \frac{\epsilon}{2}.$$

• By construction, this codebook has rate

$$\frac{1}{n}\log M > I(X;Y) - \frac{\epsilon}{2}.$$

• We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2}$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2+1}^{M} \lambda_w \le \sum_{w=1}^{M} \lambda_w < \left(\frac{M}{2}\right) \epsilon$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2+1}^{M} \lambda_w \le \sum_{w=1}^{M} \lambda_w < \left(\frac{M}{2}\right) \epsilon$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2}^{M} \lambda_w \le \sum_{w=1}^{M} \lambda_w < \left(\frac{M}{2}\right) \epsilon$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2+1}^{M} \lambda_w \le \sum_{w=1}^{M} \lambda_w < \left(\frac{M}{2}\right) \epsilon$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2+1}^{M} \lambda_w < \left(\frac{M}{2}\right)\epsilon$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2+1}^{M} \lambda_w \qquad \qquad < \left(\frac{M}{2}\right)\epsilon \implies \frac{1}{M/2} \sum_{w=M/2+1}^{M} \lambda_w < \epsilon.$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

$$\sum_{w=M/2+1}^{M} \lambda_w \qquad \qquad < \left(\frac{M}{2}\right)\epsilon \implies \frac{1}{M/2}\sum_{w=M/2+1}^{M} \lambda_w < \epsilon.$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

• Since M is even, M/2 is an integer. Then

$$\sum_{w=M/2+1}^{M} \lambda_w \qquad \qquad < \left(\frac{M}{2}\right)\epsilon \implies \frac{1}{M/2} \sum_{w=M/2+1}^{M} \lambda_w < \epsilon.$$

• Hence,

 $\lambda_{M/2\,+1} < \epsilon$ 

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

• Since M is even, M/2 is an integer. Then

$$\sum_{w=M/2+1}^{M} \lambda_w \qquad \qquad < \left(\frac{M}{2}\right) \epsilon \implies \frac{1}{M/2} \sum_{w=M/2+1}^{M} \lambda_w < \epsilon.$$

• Hence,

$$\lambda_{M/2+1} < \epsilon \implies \lambda_{M/2} < \epsilon.$$

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

• Since M is even, M/2 is an integer. Then

$$\sum_{w=M/2+1}^{M} \lambda_w \qquad \qquad < \left(\frac{M}{2}\right)\epsilon \implies \frac{1}{M/2} \sum_{w=M/2+1}^{M} \lambda_w < \epsilon.$$

• Hence,

$$\lambda_{M/2+1} < \epsilon \implies \lambda_{M/2} < \epsilon.$$

• Conclusion: If  $P_e < \epsilon/2$ , then  $\lambda_1 \le \lambda_2 \le \cdots \le \lambda_{M/2} < \epsilon$ .

- We want a code with  $\lambda_{max} < \epsilon$ , not just  $P_e < \epsilon/2$ .
- Without loss of generality, assume  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_M$ . Consider

$$\frac{1}{M}\sum_{w=1}^{M}\lambda_w < \frac{\epsilon}{2} \quad \Longleftrightarrow \quad \sum_{w=1}^{M}\lambda_w < \left(\frac{M}{2}\right)\epsilon.$$

• Since M is even, M/2 is an integer. Then

$$\sum_{w=M/2+1}^{M} \lambda_w \qquad \qquad < \left(\frac{M}{2}\right) \epsilon \implies \frac{1}{M/2} \sum_{w=M/2+1}^{M} \lambda_w < \epsilon.$$

• Hence,

$$\lambda_{M/2+1} < \epsilon \implies \lambda_{M/2} < \epsilon.$$

- Conclusion: If  $P_e < \epsilon/2$ , then  $\lambda_1 \le \lambda_2 \le \cdots \le \lambda_{M/2} < \epsilon$ .
- Discard the worst half of the codewords in  $\mathcal{C}^*$  to achieve  $\lambda_{max} < \epsilon$ .

• After discarding the worse half of  $\mathcal{C}^*$ , the rate of the code becomes

$$\frac{1}{n}\log\frac{M}{2} = \frac{1}{n}\log M - \frac{1}{n}$$

$$> \left(I(X;Y) - \frac{\epsilon}{2}\right) - \frac{1}{n}$$

$$> I(X;Y) - \epsilon$$

for sufficiently large n.

• After discarding the worse half of  $\mathcal{C}^*$ , the rate of the code becomes

$$\frac{1}{n}\log\frac{M}{2} = \frac{1}{n}\log M - \frac{1}{n}$$

$$> \left(I(X;Y) - \frac{\epsilon}{2}\right) - \frac{1}{n}$$

$$> I(X;Y) - \epsilon$$

for sufficiently large n.

• Here we assume that the decoding function is unchanged, so that deletion of worst half of the codewords does not affect the conditional probabilities of error of the remaining codewords.