

Chapter 10 Differential Entropy

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- Symmetric, positive definite, and covariance matrices

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where the right hand side is a Lebesgue-Stieltjes integration which covers all cases (i.e., discrete, continuous, and mixed) for the CDF $F_X(x)$.

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- 2. If $X \perp Y$, then cov(X, Y) = 0, or X and Y are uncorrelated. However, the converse is not true.
- 3. If X_1, X_2, \dots, X_n are mutually independent, then

$$\operatorname{var}\left(\sum_{i=1}^{n} X_{i}\right) = \sum_{i=1}^{n} \operatorname{var} X_{i}$$

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10.1 Preliminaries

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Proposition 10.3 A covariance matrix is both symmetric and positive semidefinite.

• A symmetric matrix K can be diagonalized as

 $K = Q \Lambda Q^{\top}$

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where Λ is a diagonal matrix and Q (also Q^{\top}) is an orthogonal matrix, i.e.,

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Let \mathbf{q}_i be the *i*th column of Q. Then

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$$QQ^{\top} = \begin{bmatrix} | & | \\ \mathbf{q}_1 & \cdots & \mathbf{q}_n \\ | & | \end{bmatrix} \begin{bmatrix} - & \mathbf{q}_1^{\top} & - \\ \vdots & \\ - & \mathbf{q}_n^{\top} & - \end{bmatrix} = \begin{bmatrix} \mathbf{q}_i \mathbf{q}_j^{\top} \end{bmatrix} = I.$$

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Therefore,

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Hence, $Q^{-1} = Q^{\top}$.

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$$KQ = K \begin{bmatrix} \mathbf{q}_1 & \cdots & \mathbf{q}_n \end{bmatrix} = \begin{bmatrix} \underline{K}\mathbf{q}_1 & \cdots & K\mathbf{q}_n \end{bmatrix}$$
and

$$Q\Lambda = \begin{bmatrix} \begin{vmatrix} \mathbf{q}_1 & \cdots & \mathbf{q}_n \\ \mathbf{q}_1 & \cdots & \mathbf{q}_n \end{bmatrix} \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{bmatrix} = \begin{bmatrix} \underline{\lambda}_1\mathbf{q}_1 & \cdots & \lambda_n\mathbf{q}_n \end{bmatrix}.$$

Therefore $KQ = Q\Lambda$ is equivalent to

 $K\mathbf{q}_i = \lambda_i \mathbf{q}_i$

for all i.

• A symmetric matrix K can be diagonalized as

$$K = Q \Lambda Q^{\top}$$

$$Q^{-1} = Q^{\top}.$$

- $|Q| = |Q^{\top}| = \pm 1.$
- Let $\lambda_i = i$ th diagonal element of Λ and $\mathbf{q}_i = i$ th column of Q.
- Then $KQ = (Q\Lambda Q^{\top})Q = Q\Lambda(Q^{\top}Q) = Q\Lambda$, or

$$K\mathbf{q}_i = \lambda_i \mathbf{q}_i.$$

• A symmetric matrix K can be diagonalized as

$$K = Q \Lambda Q^{\top}$$

where Λ is a diagonal matrix and Q (also Q^{\top}) is an orthogonal matrix, i.e.,

$$Q^{-1} = Q^{\top}.$$

- $|Q| = |Q^{\top}| = \pm 1.$
- Let $\lambda_i = i$ th diagonal element of Λ and $\mathbf{q}_i = i$ th column of Q.
- Then $KQ = (Q\Lambda Q^{\top})Q = Q\Lambda(Q^{\top}Q) = Q\Lambda$, or

$$K\mathbf{q}_i = \lambda_i \mathbf{q}_i.$$

• That is, \mathbf{q}_i is an eigenvector of K with eigenvalue λ_i .

Proposition 10.4 The eigenvalues of a positive semidefinite matrix are non-negative.

Proof

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Proof

1. Consider eigenvector $\mathbf{q} \neq 0$ and corresponding eigenvalue λ of K, i.e.,

 $K\mathbf{q} = \lambda \mathbf{q}.$
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Proof

1. Consider eigenvector $\mathbf{q} \neq 0$ and corresponding eigenvalue λ of K, i.e.,

 $K\mathbf{q} = \lambda \mathbf{q}.$

2. Since K is positive semidefinite,

$$0 \leq \mathbf{q}^{\top} K \mathbf{q} = \mathbf{q}^{\top} (\lambda \mathbf{q}) = \lambda (\mathbf{q}^{\top} \mathbf{q}).$$

3. $\lambda \ge 0$ because $\mathbf{q}^{\top}\mathbf{q} = \|\mathbf{q}\|^2 > 0$.

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3. $\lambda \ge 0$ because $\mathbf{q}^{\top}\mathbf{q} = \|\mathbf{q}\|^2 > 0$.

Remark Since a covariance matrix is both symmetric and positive semidefinite, it is diagonalizable and its eigenvalues are nonnegative.

$$K_{\mathbf{Y}} = \mathbf{A} K_{\mathbf{X}} \mathbf{A}^{\mathsf{T}}$$

$$K_{\mathbf{Y}} = \mathbf{A} K_{\mathbf{X}} \mathbf{A}^{\top}$$

$$\tilde{K}_{\mathbf{Y}} = \mathbf{A}\tilde{K}_{\mathbf{X}}\mathbf{A}^{\mathsf{T}}.$$

$$K_{\mathbf{Y}} = \mathbf{A} K_{\mathbf{X}} \mathbf{A}^{\mathsf{T}}$$

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$$K_{\mathbf{Y}} = \mathbf{A} K_{\mathbf{X}} \mathbf{A}^{\mathsf{T}}$$

$$\tilde{K}_{\mathbf{Y}} = A \tilde{K}_{\mathbf{X}} A^{\top}.$$

$$K_{\mathbf{Y}} = E\mathbf{Y}\mathbf{Y}^{\top} - (E\mathbf{Y})(E\mathbf{Y})^{\top}$$

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= $\underline{A}(E\mathbf{X}\mathbf{X}^{\top})A^{\top} - \underline{A}(E\mathbf{X})(E\mathbf{X}^{\top})A^{\top}$
= $\underline{A}[E\mathbf{X}\mathbf{X}^{\top} - (E\mathbf{X})(E\mathbf{X})^{\top}]A^{\top}$

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= $A[E\mathbf{X}\mathbf{X}^{\top} - (E\mathbf{X})(E\mathbf{X})^{\top}]A^{\top}$
= $AK_{\mathbf{X}}A^{\top}$

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

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$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

and

$$\tilde{K}_{\mathbf{Y}} = A \tilde{K}_{\mathbf{X}} A^{\top}.$$

Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$.

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

and

$$\tilde{K}_{\mathbf{Y}} = A \tilde{K}_{\mathbf{X}} A^{\top}.$$

Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e.,

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

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Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e.,

1. the random variables in \mathbf{Y} are uncorrelated

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

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- 1. the random variables in \mathbf{Y} are uncorrelated
- 2. var $Y_i = \lambda_i$ for all i

Lemma 10.6 Let **X** and **Y** be column vectors of n random variables such that

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

 \mathbf{Proof}

Lemma 10.6 Let **X** and **Y** be column vectors of n random variables such that

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

\mathbf{Proof}

1. By Proposition 10.5,
$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

\mathbf{Proof}

$$K_{\mathbf{Y}} = Q^{\top} K_{\mathbf{X}} Q$$

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

\mathbf{Proof}

$$K_{\mathbf{Y}} = Q^{\top} \underline{K_{\mathbf{X}}} Q$$

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

\mathbf{Proof}

$$K_{\mathbf{Y}} = Q^{\top} \underline{K}_{\mathbf{X}} Q$$
$$= Q^{\top} (Q \Lambda Q^{\top}) Q$$

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

Proof

$$K_{\mathbf{Y}} = Q^{\top} K_{\mathbf{X}} Q$$
$$= Q^{\top} (Q \Lambda Q^{\top}) Q$$
$$= (Q^{\top} Q) \Lambda (Q^{\top} Q)$$

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

Proof

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Proof

$$K_{\mathbf{Y}} = Q^{\top} K_{\mathbf{X}} Q$$
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$$= (Q^{\top} Q) \Lambda (Q^{\top} Q)$$
$$= \Lambda.$$

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

Proof

1. By Proposition 10.5,

$$K_{\mathbf{Y}} = Q^{\top} K_{\mathbf{X}} Q$$
$$= Q^{\top} (Q \Lambda Q^{\top}) Q$$
$$= (Q^{\top} Q) \Lambda (Q^{\top} Q)$$
$$= \Lambda.$$

2. Since $K_{\mathbf{Y}} = \Lambda$ is a diagonal matrix, the random variables in \mathbf{Y} are uncorrelated because

$$\operatorname{cov}(Y_i, Y_j) = 0$$

for $i \neq j$.

$$\mathbf{Y} = Q^{\top} \mathbf{X},$$

where $Q\Lambda Q^{\top}$ is a diagonalization of $K_{\mathbf{X}}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e., the random variables in \mathbf{Y} are uncorrelated and var $Y_i = \lambda_i$, the *i*th diagonal element of Λ .

Proof

1. By Proposition 10.5,

$$K_{\mathbf{Y}} = Q^{\top} K_{\mathbf{X}} Q$$
$$= Q^{\top} (Q \Lambda Q^{\top}) Q$$
$$= (Q^{\top} Q) \Lambda (Q^{\top} Q)$$
$$= \Lambda.$$

2. Since $K_{\mathbf{Y}} = \Lambda$ is a diagonal matrix, the random variables in \mathbf{Y} are uncorrelated because

$$\operatorname{cov}(Y_i, Y_j) = 0$$

for $i \neq j$.

3. Furthermore, the variance of Y_i is given by the *i*th diagonal element of $K_{\mathbf{Y}} = \Lambda$, i.e., λ_i . The proposition is proved.

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

and

$$\tilde{K}_{\mathbf{Y}} = A \tilde{K}_{\mathbf{X}} A^{\top}.$$

Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e.,

- 1. the random variables in \mathbf{Y} are uncorrelated
- 2. var $Y_i = \lambda_i$ for all i

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

and

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Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e.,

- 1. the random variables in \mathbf{Y} are uncorrelated
- 2. var $Y_i = \lambda_i$ for all i

Corollary 10.7 Any random vector \mathbf{X} can be written as a linear transformation of an uncorrelated vector. Specifically, $\mathbf{X} = Q\mathbf{Y}$, where $K_{\mathbf{X}} = Q\Lambda Q^{\top}$.

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

and

$$\tilde{K}_{\mathbf{Y}} = A \tilde{K}_{\mathbf{X}} A^{\top}.$$

Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e.,

- 1. the random variables in \mathbf{Y} are uncorrelated
- 2. var $Y_i = \lambda_i$ for all i

Corollary 10.7 Any random vector \mathbf{X} can be written as a linear transformation of an uncorrelated vector. Specifically, $\mathbf{X} = Q\mathbf{Y}$, where $K_{\mathbf{X}} = Q\Lambda Q^{\top}$.

Proof In Proposition 10.6, $\mathbf{Y} = Q^{\top} \mathbf{X}$ implies $Q \mathbf{Y} = Q Q^{\top} \mathbf{X}$, or $\mathbf{X} = Q \mathbf{Y}$.

$$K_{\mathbf{Y}} = A K_{\mathbf{X}} A^{\top}$$

and

$$\tilde{K}_{\mathbf{Y}} = A \tilde{K}_{\mathbf{X}} A^{\top}.$$

Proposition 10.6 (Decorrelation) Let $\mathbf{Y} = Q^{\top} \mathbf{X}$, where $K_{\mathbf{X}} = Q \Lambda Q^{\top}$. Then $K_{\mathbf{Y}} = \Lambda$, i.e.,

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1. If $\mathbf{Y} = \sum_{i=1}^{n} \mathbf{X}_{i}$ where $\mathbf{X}_{1}, \mathbf{X}_{2}, \cdots, \mathbf{X}_{n}$ are mutually independent, then

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$$K_{\mathbf{Y}} = \sum_{i=1}^{n} K_{\mathbf{X}_{i}}.$$

2. When \mathbf{X}_i are scalars, this reduces to

$$\operatorname{var} Y = \sum_{i=1}^{n} \operatorname{var} X_i.$$

Proposition 10.9 (Preservation of Energy) Let $\mathbf{Y} = Q\mathbf{X}$, where Q is an orthogonal matrix. Then

$$E\sum_{i=1}^{n} Y_i^2 = E\sum_{i=1}^{n} X_i^2.$$

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 \mathbf{Proof}

$$\sum_{i=1}^{n} Y_i^2 = \underline{\mathbf{Y}}^{\top} \mathbf{Y}$$
$$= (\underline{Q} \mathbf{X})^{\top} (Q \mathbf{X})$$

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 \mathbf{Proof}

$$\sum_{i=1}^{n} Y_i^2 = \mathbf{Y}^\top \mathbf{Y}$$
$$= (Q\mathbf{X})^\top (Q\mathbf{X})$$
$$= (\mathbf{X}^\top Q^\top) (Q\mathbf{X})$$

$$E \sum_{i=1}^{n} Y_i^2 = E \sum_{i=1}^{n} X_i^2.$$

 \mathbf{Proof}

$$\sum_{i=1}^{n} Y_i^2 = \mathbf{Y}^\top \mathbf{Y}$$
$$= (Q\mathbf{X})^\top (Q\mathbf{X})$$
$$= (\mathbf{X}^\top Q^\top) (Q\mathbf{X})$$
$$= \mathbf{X}^\top (Q^\top Q) \mathbf{X}$$

$$E \sum_{i=1}^{n} Y_i^2 = E \sum_{i=1}^{n} X_i^2.$$

 \mathbf{Proof}

$$\sum_{i=1}^{n} Y_i^2 = \mathbf{Y}^\top \mathbf{Y}$$
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 \mathbf{Proof}

$$\sum_{i=1}^{n} Y_i^2 = \mathbf{Y}^\top \mathbf{Y}$$
$$= (Q\mathbf{X})^\top (Q\mathbf{X})$$
$$= (\mathbf{X}^\top Q^\top) (Q\mathbf{X})$$
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$$= \mathbf{X}^\top \mathbf{X}$$

$$E \sum_{i=1}^{n} Y_i^2 = E \sum_{i=1}^{n} X_i^2.$$

 \mathbf{Proof}

$$\sum_{i=1}^{n} Y_i^2 = \mathbf{Y}^\top \mathbf{Y}$$
$$= (Q\mathbf{X})^\top (Q\mathbf{X})$$
$$= (\mathbf{X}^\top Q^\top) (Q\mathbf{X})$$
$$= \mathbf{X}^\top (Q^\top Q) \mathbf{X}$$
$$= \mathbf{X}^\top \mathbf{X}$$
$$= \sum_{i=1}^{n} X_i^2.$$

$$E \sum_{i=1}^{n} Y_i^2 = E \sum_{i=1}^{n} X_i^2.$$

 \mathbf{Proof}

1. Consider

$$\sum_{i=1}^{n} Y_i^2 = \mathbf{Y}^\top \mathbf{Y}$$
$$= (Q\mathbf{X})^\top (Q\mathbf{X})$$
$$= (\mathbf{X}^\top Q^\top) (Q\mathbf{X})$$
$$= \mathbf{X}^\top (Q^\top Q) \mathbf{X}$$
$$= \mathbf{X}^\top \mathbf{X}$$
$$= \sum_{i=1}^{n} X_i^2.$$

2. The proposition is proved upon taking expectation on both sides.