

THE CHINESE UNIVERSITY OF HONG KONG  
Department of Mathematics

## MMAT 5340 Probability and Stochastic Analysis

### Suggested Solution of Homework 2

**11.29** Let  $x(n) = [P(X_n = 1) \ P(X_n = 2)]$ . Then

$$x(2) = x(0)A^2 = [0.4 \ 0.6] \begin{bmatrix} 0.4 & 0.6 \\ 0.3 & 0.7 \end{bmatrix}^2 = [0.334 \ 0.666].$$

Hence  $P(X_2 = 1) = 0.334$ .

**11.30** The stationary distribution  $p = [p_1 \ p_2]$  satisfies

$$p = pA \iff \begin{cases} p_1 = 0.4p_1 + 0.3p_2 \\ p_2 = 0.6p_1 + 0.7p_2 \end{cases} \iff 2p_1 = p_2.$$

Since  $p_1 + p_2 = 1$ , we have  $p = [\frac{1}{3} \ \frac{2}{3}]$ .

**11.31** Solving

$$0 = |A - \lambda I| = \begin{vmatrix} 0.4 - \lambda & 0.6 \\ 0.3 & 0.7 - \lambda \end{vmatrix} = \lambda^2 - 1.1\lambda + 0.1,$$

the eigenvalues of  $A$  are given by

$$\lambda_1 = 1, \quad \lambda_2 = \frac{1}{10}.$$

Hence the rate of convergence is  $|\lambda_2|^n = \frac{1}{10^n}$ .

**11.32** Let  $p_x$  be the probability that, starting from  $x$ , the process hits 2 before 3. Then

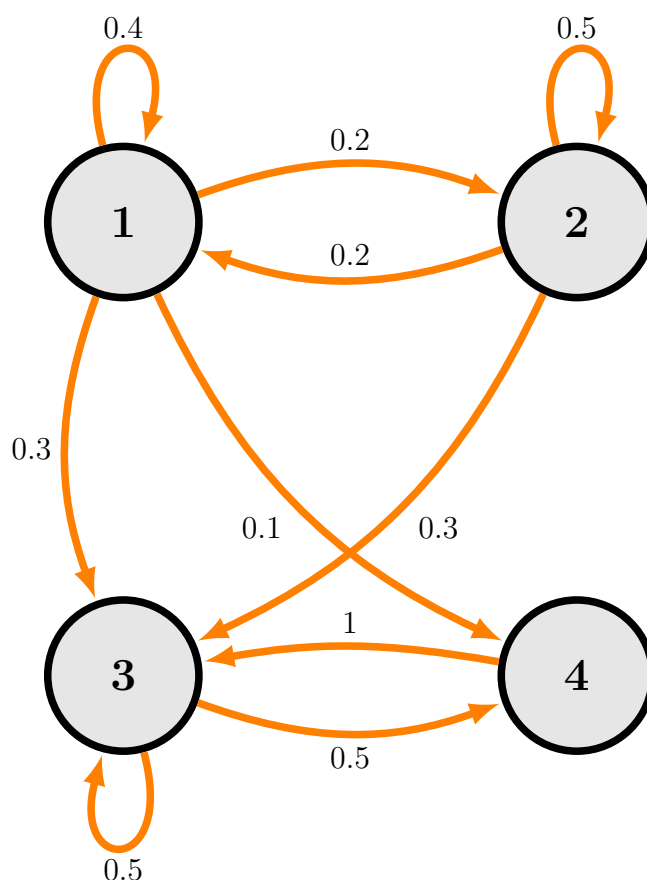
$$p_2 = 1, \quad p_3 = 0,$$

and

$$\begin{aligned} p_1 &= 0.4p_1 + 0.4p_2 + 0.2p_3 \\ &= 0.4p_1 + 0.4 + 0, \end{aligned}$$

so that  $p_1 = \frac{2}{3}$ .

**11.36** The graph of the Markov chain is given below:



From the graph, we can see that 1 and 2 are transient states, while 3 and 4 are recurrent states.

**11.40** Let  $X = Z - 1$ . Then

$$P(X = 0) = P(Z = 1) = 0.5, \quad P(X = 1) = P(Z = 2) = 0.5^2 = 0.25,$$

$$P(X \geq 2) = 1 - P(X = 0) - P(X = 1) = 0.25$$

Hence, the transition matrix is

$$P = \begin{bmatrix} 0.5 & 0.5 & 0 \\ 0.5 & 0.25 & 0.25 \\ 0 & 0.5 & 0.5 \end{bmatrix}$$

We can find the stationary distribution  $p = [p_0 \ p_1 \ p_2]$  by solving  $pP = p$ : By performing column operations,

$$P - I = \begin{bmatrix} -0.5 & 0.5 & 0 \\ 0.5 & -0.75 & 0.25 \\ 0 & 0.5 & -0.5 \end{bmatrix} \sim \begin{bmatrix} -0.5 & 0 & 0 \\ 0.5 & -0.25 & 0.25 \\ 0 & 0.5 & -0.5 \end{bmatrix}$$

$$\sim \begin{bmatrix} -0.5 & 0 & 0 \\ 0.5 & -0.25 & 0 \\ 0 & 0.5 & 0 \end{bmatrix} \sim \begin{bmatrix} 1 & 0 & 0 \\ -1 & 1 & 0 \\ 0 & -2 & 0 \end{bmatrix}.$$

That is  $p_0 = p_1 = 2p_2$ . Hence  $p = [0.4 \ 0.4 \ 0.2]$ .

The long-term average premium is

$$\begin{aligned} p_0r_0 + p_1r_1 + p_2r_2 &= (0.4)(0.5 \cdot 0 + 1) + (0.4)(0.5 \cdot 1 + 1) + (0.2)(0.5 \cdot 2 + 1) \\ &= 1.4 \text{ thousands of dollars.} \end{aligned}$$

**12.12** Let  $n$  and  $m$  be the quantity of up and down steps. Then

$$\begin{cases} n + m = 13 - 2 = 11 \\ n - m = 3 - 2 = 1. \end{cases}$$

Hence  $n = 6, m = 5$ . Thus

$$P(S_{13} = 3 | S_2 = 2) = \binom{11}{6} (0.7)^6 (0.3)^5 \approx 0.1321.$$

**12.13** By the reflection principle, every path from  $(2, 2)$  to  $(13, 3)$  which hits  $y = 1$  corresponds to a path from  $(2, 2)$  to  $(13, -1)$ . Solving

$$\begin{cases} n + m = 13 - 2 = 11 \\ n - m = -1 - 2 = -3, \end{cases}$$

we have  $n = 4, m = 7$ . The number of such path is  $\binom{11}{4}$ . Thus

$$P(S_{13} = 3, S_n > 1, n = 2, 3, \dots, 13 | P_2 = 2) = \left[ \binom{11}{6} - \binom{11}{4} \right] (0.7)^6 (0.3)^5 \approx 0.0377.$$

**12.14** Solving

$$\begin{cases} n + m = 10 - 0 = 10 \\ n - m = 2 - 0 = 2, \end{cases} \implies \begin{cases} n = 6 \\ m = 4. \end{cases}$$

So there are  $\binom{10}{6}$  paths from  $(0, 0)$  to  $(10, 2)$ . By the reflection principle, every path from  $(0, 0)$  to  $(10, 2)$  which hits  $y = -2$  corresponds to a path from  $(0, 0)$  to  $(10, -2 - (2 - (-2))) = (10, -6)$ . Solving

$$\begin{cases} n + m = 10 - 0 = 10 \\ n - m = -6 - 0 = -6, \end{cases} \implies \begin{cases} n = 2 \\ m = 8. \end{cases}$$

So there are  $\binom{10}{2}$  paths from  $(0, 0)$  to  $(10, -6)$ . Thus,

$$P(S_{10} = 2, S_n > -2, n = 0, \dots, 10) = \left[ \binom{10}{6} - \binom{10}{2} \right] (0.5)^{10} \approx 0.1611.$$

**12.27** The risk-neutral probability  $p_0, q_0$  can be found by  $\mathbf{E}_0 P_1 = P_0$ :

$$\begin{cases} p_0 \cdot 2 + q_0 \cdot 0.3 = 1 \\ p_0 + q_0 = 1 \end{cases} \implies p_0 = \frac{7}{17}, q_0 = \frac{10}{17}.$$

Hence, the fair price is given by

$$\begin{aligned} v &= \mathbf{E}_0(P_3 - K)_+ \\ &= p_0^3(2^3 - 1) + 3p_0^2q_0(2^2 \cdot 0.3 - 1) + 3p_0q_0^2(0) + q_0^3(0) \\ &\approx 0.5485. \end{aligned}$$

**13.20** Since  $X_n$  is the number of Heads during the first  $n$  tosses, we have

$$\mathbf{E}(X_{n+1} | X_n) = \frac{1}{2}(X_n + 1) + \frac{1}{2}X_n = X_n + \frac{1}{2}.$$

If  $Y_n := 3X_n - cn$  is a martingale, then

$$\begin{aligned} Y_n &= \mathbf{E}(Y_{n+1} | Y_0, \dots, Y_n) = \mathbf{E}(Y_{n+1} | X_0, \dots, X_n) \\ &= \mathbf{E}(3X_{n+1} - c(n+1) | X_n) \\ &= 3\mathbf{E}(X_{n+1} | X_n) - c(n+1) \\ &= 3X_n + \frac{3}{2} - c(n+1) \\ &= Y_n + \frac{3}{2} - c. \end{aligned}$$

Hence  $c = \frac{3}{2}$ .

**13.30** Need to find  $c$  such that  $M_n = e^{S_n - cn}$  is a martingale:

$$\mathbf{E}e^{-1+X_1-c} = e^{-1} \implies e^{4/2} = F_{N(0,4)}(1) = e^c \implies c = 2.$$

For  $x > 0$ ,  $f(x) = x^3$  is convex, since  $f''(x) = 6x > 0$ . Note that  $M_n > 0$ . By Doob's martingale inequality, for  $\lambda > 0$ , we have,

$$\mathbf{P}\left(\max_{0 \leq n \leq 100} M_n \geq \lambda\right) \leq \frac{\mathbf{E}M_{100}^3}{\lambda^3}.$$

Now

$$M_{100}^3 = e^{-3} e^{3S_{100}} e^{-300c} = e^{-3} e^{-300c} e^{3X_1} \dots e^{3X_{100}},$$

so that

$$\mathbf{E}M_{100}^3 = e^{-3} e^{-300c} (\mathbf{E}e^{3X_1})^{100} = e^{-3} e^{-300(2)} \left(e^{\frac{(3)^2 4}{2}}\right)^{100} = e^{1197}.$$

Hence

$$\mathbf{P}\left(\max_{0 \leq n \leq 100} M_n \geq \lambda\right) \leq \frac{e^{1197}}{\lambda^3}.$$

**14.12** Note that  $\tau_4 - \tau_3$  and  $\tau_3 - \tau_2$  are i.i.d.  $\text{Exp}(\lambda)$  random variables, with  $\lambda = 3$ . Hence

$$\tau_4 - \tau_2 = (\tau_4 - \tau_3) + (\tau_3 - \tau_2) \sim \Gamma(2, \lambda) = \Gamma(2, 3).$$

Therefore,

$$\begin{aligned} \mathbf{E}(\tau_4 - \tau_2) &= 2\lambda^{-1} = \frac{2}{3}, \\ \text{Var}(\tau_4 - \tau_2) &= 2\lambda^{-2} = \frac{2}{9}. \end{aligned}$$

**14.13** Since  $N(t) - N(s)$  is independent of  $N(u)$ ,  $u \leq s$ , and  $N(t) - N(s) \sim \text{Poi}(\lambda(t - s))$ , we have

$$\begin{aligned} \mathbf{P}(N(5/2) = 3 \mid N(1) = 1) &= \mathbf{P}(N(5/2) - N(1) = 2 \mid N(1) = 1) \\ &= \mathbf{P}(N(5/2) - N(1) = 2) \\ &= \frac{(\lambda(3/2))^2}{2!} e^{-\lambda(3/2)} \\ &= \frac{81}{8} e^{-9/2}. \end{aligned}$$

**14.18** Note that  $\mathbf{E}N(t) = \text{Var}N(t) = \lambda t = t$ . By (35) and (36),

$$\mathbf{E}X(t) = \mathbf{E}N(t) \cdot \mathbf{E}Z_k = t \cdot 2 = 2t,$$

$$\text{Var}X(t) = \mathbf{E}N(t) \cdot \text{Var}Z_k + \text{Var}N(t) \cdot (\mathbf{E}Z_k)^2 = t \cdot 3 + t \cdot 2^2 = 7t.$$

**15.24** The stationary distribution  $\rho = [\rho_1 \ \rho_2 \ \rho_3]$  satisfies  $\rho A = 0$ . Since  $\rho_1 + \rho_2 + \rho_3 = 1$ , we have

$$\rho = \frac{1}{1 + 5 + 2} [1 \ 5 \ 2] = \left[ \frac{1}{8} \ \frac{5}{8} \ \frac{1}{4} \right].$$

**15.25** Since  $0 > \lambda_2 > \lambda_3$ , the rate of convergence is  $e^{\lambda_2 t} = e^{-2t}$ .

**15.26** The transition matrix for the corresponding discrete-time Markov chain is

$$P = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 0 & 0 & 1 \\ \frac{1}{3} & \frac{2}{3} & 0 \end{bmatrix}.$$

**15.27** Let  $\lambda'_i$ ,  $1 \leq i \leq 3$ , be the intensity of exit from state  $i$ , that is, the negative of the diagonal entries of  $A$ . Then the stationary distribution  $\pi$  for the corresponding discrete-time Markov chain is given by

$$\begin{aligned} \pi &= \frac{1}{\lambda'_1 \rho_1 + \lambda'_2 \rho_2 + \lambda'_3 \rho_3} [\lambda'_1 \rho_1 \ \lambda'_2 \rho_2 \ \lambda'_3 \rho_3] \\ &= \frac{8}{13} \left[ \frac{1}{4} \ \frac{5}{8} \ \frac{3}{4} \right] \\ &= \left[ \frac{2}{13} \ \frac{5}{13} \ \frac{6}{13} \right]. \end{aligned}$$