

March 10th:

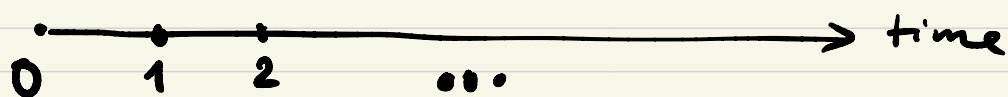
§2. Average number of visits to a state

Setting: MC  $\{X_n\}_{n=0}^{\infty}$ ,

$S$ : state space,  $\#S \leq \infty$

$P$ : Markov transition function

Recall:  $y \in S$



$$\text{r.v. } N(y) = \sum_{m=1}^{\infty} 1_y(X_m)$$

total no of visits to  $y$   
for all positive time



r.v.: For  $n \geq 1$ ,

$$N_n(y) \stackrel{\text{def.}}{=} \sum_{m=1}^n 1_y(X_m)$$

this is the no of visits to  $y \in S$   
in  $n$ -steps (i.e. at positive times  
1 to  $n$ )

Recall :

$$E_x(N(y)) = \sum_{m=1}^{\infty} P^m(x, y)$$

So, similarly,

$$E_x(N_n(y)) = \sum_{m=1}^n P^m(x, y)$$

We want to study :

$\frac{N_n(y)}{n}$  : proportion / visit frequency  
(i.e. no of visits to  $y$  / unit time)

$E_x\left(\frac{N_n(y)}{n}\right)$  : (average) visit frequency

long-run behavior  
 $n \rightarrow \infty$

Remark: If  $y$  is transient, then answer is easy:  
( $\Leftrightarrow P_x(N(y)=\infty) = 0$ )

$$\lim_{n \rightarrow \infty} N_n(y) \stackrel{\text{prob.}}{=} N(y) (< \infty)$$

(Ref.:  $\exists_n \xrightarrow{P} \exists$  as  $n \rightarrow \infty$ )

$$\Leftrightarrow \lim_{n \rightarrow \infty} P(|\exists_n - \exists| > \epsilon) = 0, \forall \epsilon > 0$$

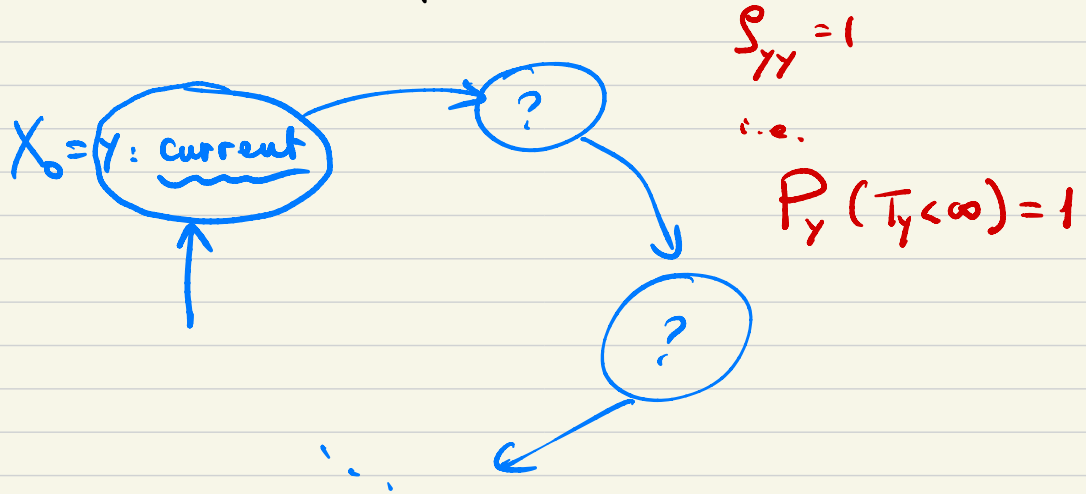
$$\Rightarrow \lim_{n \rightarrow \infty} \frac{N_n(y)}{n} \stackrel{\text{prob.}}{=} 0$$

$$\lim_{n \rightarrow \infty} E_x\left(\frac{N_n(y)}{n}\right) = E_x\left(\lim_{n \rightarrow \infty} \frac{N_n(y)}{n}\right) = 0$$

this means :

the visit frequency to any transient state must be zero in the long-run.

Question: What if  $y$  recurrent?



Expect:

$$\text{“ visit frequency } = \frac{1}{\text{returning time}} \text{ ”}$$

i.e. no of visits / unit time  
 to  $y$

i.e. the waiting time for the chain from  $y$  to return back to  $y$ .

Thm: Let  $\{X_n\}_{n=0}^{\infty}$  be an irreducible recurrent MC, then

$$\lim_{n \rightarrow \infty} \frac{N_n(y)}{n} \stackrel{\text{prob.}}{=} \frac{1}{m_y}$$

$$\lim_{n \rightarrow \infty} E_x \left( \frac{N_n(y)}{n} \right) = \frac{1}{m_y}, \quad \forall x \in S$$

where  $m_y \stackrel{\text{def.}}{=} E_y(T_y)$

a recurrent state

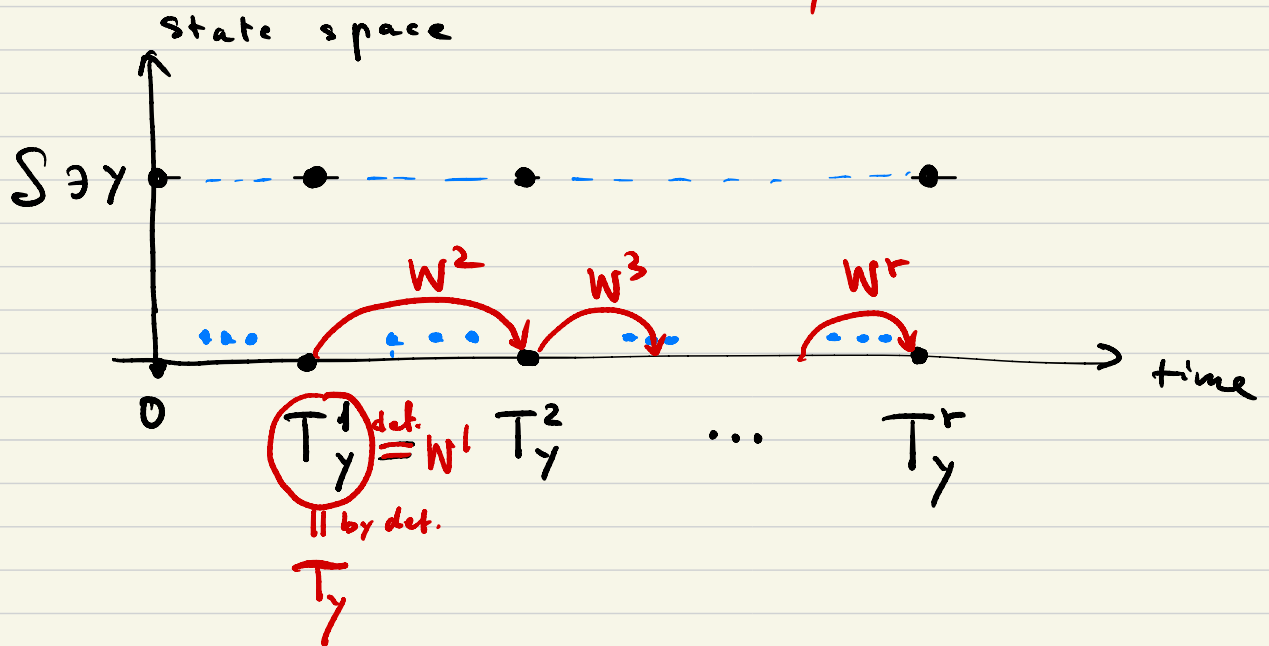
means the average time of returning to  $y$  for the chain starting from  $y$ .

Proof:  $X_0 = y$  (current).

For  $r = 1, 2, \dots$

$$T_y^r \stackrel{\text{def.}}{=} \min \{ n \geq 1 : N_n(y) = r \}$$

r.v. denoting the min positive  $n$  for the  $r^{\text{th}}$  visits to  $y$



$$\begin{cases} W^1 \stackrel{\text{def.}}{=} T_y^1 = \min \{ n \geq 1, N_n(y) = 1 \} = T_y \\ W^r \stackrel{\text{def.}}{=} T_y^r - T_y^{r-1}, \quad r = 1, 2, \dots \end{cases} \quad (= \min \{ n \geq 1 : X_n = y \})$$

$\Rightarrow$  a sequence of r.v.s

$$W^1, W^2, \dots, W^r, \dots$$

$\stackrel{\parallel}{=} T_y$  is i.i.d.

note:  $T_y^r = W^1 + W^2 + \dots + W^r$

Strong law  
of large number  
→

$$\lim_{n \rightarrow \infty} \frac{T_y^r}{n} = \lim_{n \rightarrow \infty} \frac{W^1 + W^2 + \dots + W^r}{n} \stackrel{\text{prob.}}{=} E_y(W^1) = E_y(T_y) = m_y$$

Note:

Given  $N_n(y) = r$ ,

(i.e.  $r$  visits to  $y$  in  $n$ -steps)

(otherwise,  $T_y^{r+1} \leq n$ , then  $N_n(y) \geq r+1$ )

then  $T_y^r \leq n < T_y^{r+1}$

∴

$$m_y \left( \frac{T_y^r}{r} \right) \leq \frac{n}{N_n(y)} = \frac{n}{r} < \left( \frac{T_y^{r+1}}{r} \right)$$

consider the event  $N_n(y) = r$

Squeezing law  $\Rightarrow$

$\frac{T_y^{r+1}}{r} = \frac{T_y^{r+1}}{r+1} \cdot \frac{r+1}{r} \rightarrow 1$

prob.  $\rightarrow m_y$

$$\lim_{n \rightarrow \infty} \frac{n}{N_n(y)} \stackrel{\text{prob.}}{=} m_y \quad \#$$

Moreover,

due to dominated convergence thm

$$\lim_{n \rightarrow \infty} E_x \left( \frac{N_n(y)}{n} \right) \stackrel{?}{=} E_x \left( \lim_{n \rightarrow \infty} \frac{N_n(y)}{n} \right) = E_x \left( \frac{1}{m_y} \right)$$

YES

$$= \frac{1}{m_y} \cdot \#$$

FYI:

D.C.T.:

$$\left( \begin{array}{l} * \quad \bar{z}_n(\omega) \xrightarrow{n \rightarrow \infty} \bar{z}(\omega) \quad \forall \omega \in \Omega \\ * \quad \exists \eta \text{ (r.v.) s.t.} \\ \quad |\bar{z}_n| \leq \eta, \quad E(\eta) < \infty \end{array} \right.$$

$$\Rightarrow \lim_{n \rightarrow \infty} E(|\bar{z}_n - \bar{z}|) = 0$$

particularly

$$\lim_{n \rightarrow \infty} E(\bar{z}_n) = E(\bar{z})$$

$$= E\left(\lim_{n \rightarrow \infty} \bar{z}_n\right)$$

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Final remark: General situation

$$\underline{MC}, \quad S = S_R \cup S_T = \left( \bigcup_{i=1}^k C_i \right) \cup S_T$$

$\neq \emptyset$

$$y \in S_R$$

$$\lim_{n \rightarrow \infty} \frac{N_n(y)}{n} \stackrel{\text{prob.}}{=} \frac{\mathbb{1}_{\{T_y < \infty\}}}{m_y}$$

$$\lim_{n \rightarrow \infty} E_x \left( \frac{N_n(y)}{n} \right) = \frac{p_{xy}}{m_y}, \quad \forall x \in S$$

$$\parallel$$
$$E_x \left( \frac{1_{\{T_y < \infty\}}}{m_y} \right) = \frac{E_x(1_{\{T_y < \infty\}})}{m_y}$$

$$= \frac{1 \cdot P_x(T_y < \infty)}{m_y}$$

$$\parallel \frac{p_{xy}}{m_y}$$