

**The Chinese University of Hong Kong**  
**Department of Mathematics**  
MATH3280A Introductory Probability  
Solutions to Assignment 4

**p.224 Q4**

(a)

$$P(X > 20) = \int_{20}^{\infty} \frac{10}{x^2} dx = \frac{1}{2}.$$

(b) The cumulative distribution function of  $X$  is

$$\begin{aligned} F(t) &= \begin{cases} \int_{10}^t \frac{10}{x^2} dx & t \geq 10 \\ 0 & t < 10 \end{cases} \\ &= \begin{cases} 1 - \frac{10}{t} & t \geq 10 \\ 0 & t < 10 \end{cases} \end{aligned}$$

(c) Assume that the lifetimes of the electronic devices are independent. Let  $Y$  be the random variable of the number of devices that will function for at least 15 hours. Then  $Y$  has a binomial distribution with parameters  $n = 6$  and  $p$ , where

$$p = P(X \geq 15) = \int_{15}^{\infty} \frac{10}{x^2} dx = \frac{2}{3}.$$

The required probability is

$$P(Y \geq 3) = 1 - \sum_{k=0}^2 \binom{6}{k} p^k (1-p)^{6-k} = \frac{656}{729} \approx 0.8999$$

**p.225 Q7**

First, note that

$$1 = \int_{-\infty}^{\infty} f(x) dx = \int_0^1 (a + bx^2) dx = a + \frac{1}{3}b. \quad (1)$$

Moreover, we have

$$\frac{3}{5} = E(X) = \int_0^1 x(a + bx^2) dx = \frac{1}{2}a + \frac{1}{4}b. \quad (2)$$

Solving (1) and (2), we have  $a = 3/5$  and  $b = 6/5$ .

**p.225 Q11**

We may identify the line segment with the interval  $[0, L]$ . A point  $x$  is randomly chosen on the interval and let  $X$  be the random variable given by the value of  $x$ . Then  $X$  is the uniform random variable over  $[0, L]$ . Note that

$$\left( X < L - X \text{ and } \frac{X}{L - X} < \frac{1}{4} \right) \Leftrightarrow X < \frac{L}{5}$$

and

$$\left( X > L - X \text{ and } \frac{L - X}{X} < \frac{1}{4} \right) \Leftrightarrow X > \frac{4L}{5}.$$

The required probability is

$$P(X < L/5) + P(X > 4L/5) = 1/5 + 1/5 = 2/5.$$

**p.225 Q16**

Assume that the annual rainfalls are independent from year to year. Let  $X$  be the random variable of annual rainfall. Then  $X \sim N(40, 4^2)$ .

$$P(X \leq 50) = P\left(\frac{X - 40}{4} \leq 2.5\right) = \Phi(2.5) \approx 0.9938$$

The required probability is  $P(X \leq 50)^{10} \approx 0.9397$ .

**p.225 Q18**

Suppose  $X \sim N(5, \sigma^2)$ . Then we have

$$0.2 = P(X > 9) = P\left(\frac{X - 5}{\sigma} > \frac{4}{\sigma}\right) = 1 - \Phi(4/\sigma)$$

i.e.  $\Phi(4/\sigma) = 0.8$ . Checking the standard normal table, we have  $\Phi(0.84) \approx 0.8$  and thus  $4/\sigma \approx 0.84$ . Hence  $Var(X) = \sigma^2 \approx 22.7$ .

**p.226 Q32**

(a)

$$P(X > 2) = \int_2^{\infty} \frac{1}{2} e^{-\frac{1}{2}x} dx = e^{-1}.$$

(b)

$$P(X \geq 10 | X > 9) = \frac{P(X \geq 10)}{P(X > 9)} = \frac{\int_{10}^{\infty} \frac{1}{2} e^{-\frac{1}{2}x} dx}{\int_9^{\infty} \frac{1}{2} e^{-\frac{1}{2}x} dx} = \frac{e^{-10/2}}{e^{-9/2}} = e^{-1/2}.$$

**p.227 Q8**

We assume that  $X$  is a continuous random variable with density  $f(x)$ .

$$\begin{aligned} E(X^2) &= \int_0^c x^2 f(x) dx \leq c \int_0^c x f(x) dx = cE(X). \\ \text{Var}(X) &= E(X^2) - E(X)^2 \\ &\leq cE(X) - E(X)^2 \\ &= -\left(E(X) - \frac{c}{2}\right)^2 + \frac{c^2}{4} \\ &\leq \frac{c^2}{4}. \end{aligned}$$

**p.228 Q11**

(a) By integration by parts, we have

$$\begin{aligned} E(g'(Z)) &= \int_{-\infty}^{\infty} g'(x) \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \frac{1}{\sqrt{2\pi}} \left[ g(x) e^{-\frac{x^2}{2}} \right]_{-\infty}^{\infty} + \int_{-\infty}^{\infty} x g(x) \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \end{aligned}$$

Here, we need the additional assumption that

$$\lim_{x \rightarrow \pm\infty} g(x) e^{-\frac{x^2}{2}} = 0.$$

Then we have  $E(g'(Z)) = E(Zg(Z))$ .

(b) For  $n = 1$ ,  $E(Z^2) = \text{Var}(Z) + E(Z)^2 = 1 = 1 \cdot E(Z^0)$ .

For  $n \geq 2$ , put  $g(x) = x^n$ . Note that the additional assumption in (a) is satisfied. Then by (a), we have  $E(Z^{n+1}) = E(Zg(Z)) = E(g'(Z)) = E(nZ^{n-1}) = nE(Z^{n-1})$ .

(c) By (b), we have  $E(Z^4) = 3E(Z^2) = 3$ .

**p.228 Q15**

Let  $F_X$  and  $F_{cX}$  be the distribution of  $X$  and  $cX$  respectively. Let  $f_X$  and  $f_{cX}$  be the density of  $X$  and  $cX$  respectively. For  $t > 0$ ,

$$F_{cX}(t) = P(cX \leq t) = P(X \leq t/c) = F_X(t/c),$$

$$f_{cX}(t) = F'_{cX}(t) = \frac{1}{c} f_X(t/c) = \frac{\lambda}{c} e^{-\frac{\lambda}{c}t}.$$

For  $t < 0$ ,  $F_{cX}(t) = 0$  and  $f_{cX}(t) = 0$ . Hence  $cX$  is exponential with parameter  $\lambda/c$ .