The Chinese University of Hong Kong Department of Mathematics MATH3280A Introductory Probability Solutions to Assignment 2

p.102 Q1

Let *E* be the event that at least one of the dice lands on 6. Let *F* be the event that the dice land on different numbers. Since $P(E \cap F) = 10/36$ and P(F) = 30/36, the conditional probability of *E* given *F* is

$$P(E|F) = \frac{P(E \cap F)}{P(F)} = \frac{1}{3}.$$

p.102 Q4

Let E be the event that at least one of the dice lands on 6. Let F_i be the event that the sum of the dice is i, where i = 2, 3, ..., 12.

$$P(E \cap F_i) = \begin{cases} 0 & i = 2, 3, \dots, 6\\ 2/36 & i = 7, 8, \dots, 11\\ 1/36 & i = 12. \end{cases}$$
$$P(F_i) = \begin{cases} 6/36 & i = 7\\ 5/36 & i = 8\\ 4/36 & i = 9\\ 3/36 & i = 10\\ 2/36 & i = 11\\ 1/36 & i = 12. \end{cases}$$

Hence the conditional probability of E given F_i is

$$P(E|F_i) = \frac{P(E \cap F_i)}{P(F_i)} = \begin{cases} 0 & i = 2, 3, \dots, 6\\ 1/3 & i = 7\\ 2/5 & i = 8\\ 1/2 & i = 9\\ 2/3 & i = 10\\ 1 & i = 11, 12. \end{cases}$$

or

(a)

$$P(B|A_s) = \frac{P(B \cap A_s)}{P(A_s)} = \frac{\frac{1}{52} \times \frac{3}{51} + \frac{3}{52} \times \frac{1}{51}}{\frac{1}{52} \times 1 + \frac{51}{52} \times \frac{1}{51}} = \frac{1}{17}$$
$$P(B|A_s) = \frac{\binom{3}{1} / \binom{52}{2}}{\binom{51}{1} / \binom{52}{2}} = \frac{1}{17}.$$

(b)

$$P(B|A) = \frac{P(B \cap A)}{P(A)} = \frac{\frac{4}{52} \times \frac{3}{51}}{1 - \frac{48}{52} \times \frac{47}{51}} = \frac{1}{33}$$
$$P(B|A) = \frac{\binom{4}{2}}{1 - \left[\binom{48}{2} + \binom{52}{2}\right]} = \frac{1}{33}.$$

p.103 Q 23

or

Let W be the event that white ball is drawn from urn II. Let T be the event that the transferred ball is white.

(a) By the law of total probability,

$$P(W) = P(T)P(W|T) + P(T^c)P(W|T^c) = \frac{2}{6} \times \frac{2}{3} + \frac{4}{6} \times \frac{1}{3} = \frac{4}{9}$$

(b) By the Bayes' formula,

$$P(T|W) = \frac{P(T)P(W|T)}{P(W)} = \frac{\frac{2}{6} \times \frac{2}{3}}{\frac{4}{9}} = \frac{1}{2}$$

p.109 Q73

(a) The probability that all children are of the same sex is equal to

$$\frac{1}{2^5} + \frac{1}{2^5} = \frac{1}{16}.$$

(b) The probability that the 3 eldest children are boys and the others are girls is equal to

$$\frac{1}{2^5} = \frac{1}{32}.$$

(c) The probability that exactly 3 children are boys is equal to

$$\frac{1}{2^5} \times \binom{5}{3} = \frac{5}{16}.$$

(d) The probability that the 2 eldest children are girls is equal to

$$1 \times 1 \times 1 \times \frac{1}{2} \times \frac{1}{2} = \frac{1}{4}.$$

(e) The probability that there is at least one girl is equal to

$$1 - \frac{1}{2^5} = \frac{31}{32}.$$

p.110 Q1

Note that

$$P(AB|A) = \frac{P(AB \cap A)}{P(A)} = \frac{P(AB)}{P(A)}$$

and

$$P(AB|A \cup B) = \frac{AB \cap (A \cup B)}{P(A \cup B)} = \frac{P(AB)}{P(A \cup B)}$$

Since $A \subseteq A \cup B$ implies that $P(A) \leq P(A \cup B)$, it follows that $P(AB|A) \geq P(AB|A \cup B)$.

p.112 Q16

Let E_n be the event that there are even number of successes among the first *n* trials. Let S_n be the event that the *n*-th trial is a success. Then for $n \ge 1$,

$$P_n = P(E_n) = P(S_n | E_{n-1}^c) P(E_{n-1}^c) + P(S_n^c | E_{n-1}) P(E_{n-1})$$

= $P(S_n) P(E_{n-1}^c) + P(S_n^c) P(E_{n-1})$ (the trials are independent)
= $p(1 - P_{n-1}) + (1 - p) P_{n-1}.$

Next, we prove that for all $n \in \mathbb{N}$,

$$P_n = \frac{1 + (1 - 2p)^n}{2}.$$
 (*)

For n = 1, we have $P_1 = 1 - p = \frac{1 + (1 - 2p)}{2}$. Suppose (*) holds for some $n \in \mathbb{N}$. Then

$$P_{n+1} = p(1 - P_n) + (1 - p)P_n$$

= $p(1 - \frac{1 + (1 - 2p)^n}{2}) + (1 - p)\frac{1 + (1 - 2p)^n}{2}$
= $\frac{2p - p + p(1 - 2p)^n + 1 - p + (1 - p)(1 - 2p)^n}{2}$
= $\frac{1 + (1 - 2p)^{n+1}}{2}$.

Hence the result follows from Mathematical induction.

$$P(E|FG)P(G|F) + P(E|FG^{c})P(G^{c}|F) = \frac{P(EFG)}{P(FG)}\frac{P(FG)}{P(FG)} + \frac{P(EFG^{c})}{P(FG^{c})}\frac{P(FG^{c})}{P(F)}$$
$$= \frac{P(EFG) + P(EFG^{c})}{P(F)}$$
$$= \frac{P(EF)}{P(F)}$$
$$= P(E|F).$$