

## Review

- Joint distribution of two r.v.'s
- Independence of two r.v.'s.

Remark: The above concepts can be generalized to finitely many r.v.'s.

Let  $X_1, \dots, X_n$  be r.v.'s.

① The joint CDF of  $X_1, \dots, X_n$  is

$$F(a_1, \dots, a_n) = P\{X_1 \leq a_1, \dots, X_n \leq a_n\}, \quad a_1, \dots, a_n \in \mathbb{R}.$$

② Say that  $X_1, \dots, X_n$  are jointly cts if  $\exists f: \mathbb{R}^n \rightarrow [0, \infty)$   
such that

$$P\{(X_1, \dots, X_n) \in C\} = \iint_C f(x_1, \dots, x_n) dx_1 \dots dx_n$$

for all "measurable" sets  $C \subset \mathbb{R}^n$ .

③ Say  $X_1, \dots, X_n$  are independent if

$$P\{X_1 \in A_1, \dots, X_n \in A_n\} = P\{X_1 \in A_1\} \dots P\{X_n \in A_n\}$$

(\*) for all  $A_1, \dots, A_n \subset \mathbb{R}$ .

Also (\*) is equivalent to

$$(**) F(a_1, \dots, a_n) = F_{X_1}(a_1) \cdots F_{X_n}(a_n), \quad \forall a_1, \dots, a_n \in \mathbb{R}.$$

### § 6.3 Sums of independent r.v.'s.

Q: Let  $X, Y$  be independent r.v.'s.

How to determine the distribution of  $X+Y$ ?

1. The case when  $X, Y$  are both continuous.

Suppose  $X$  has a density  $f_X(x)$ ,  $Y$  has a density  $f_Y(y)$ .

Then  $X, Y$  have a joint density

$$f(x, y) = f_X(x) f_Y(y).$$

For  $a \in \mathbb{R}$ ,

$$\begin{aligned} F_{X+Y}(a) &= P\{X+Y \leq a\} = \iint_{\{x+y \leq a\}} f_X(x) f_Y(y) dx dy \\ &= \int_{-\infty}^{\infty} \int_0^{a-y} f_X(x) f_Y(y) dx dy \end{aligned}$$

$$= \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy.$$

$$=: F_X * f_Y^0(a) \quad \textcircled{1}$$

( where  $g * h(a) := \int_{-\infty}^{\infty} g(a-y) h(y) dy.$  )

In particular

$$\begin{aligned} f_{X+Y}^0(a) &= \frac{d F_{X+Y}(a)}{da} = \int_{-\infty}^{\infty} \frac{d}{da} (F_X(a-y)) f_Y^0(y) dy \\ &= \int_{-\infty}^{\infty} f_X(a-y) f_Y^0(y) dy \\ &= f_X * f_Y^0(a). \quad \textcircled{2} \end{aligned}$$

Example 1. If  $X, Y$  are independent, and both unif. dist. on  $[0,1]$ , calculate the joint density of  $X+Y$ .

Solution:

$$\begin{aligned} f_{X+Y}^0(a) &= f_X * f_Y^0(a) \\ &= \int_{-\infty}^{\infty} f_X(a-y) f_Y^0(y) dy. \end{aligned}$$

$$= \int_0^1 f_X(a-y) dy$$

$$= \int_{a-1}^a f_X(z) dz.$$

If  $0 < a \leq 1$

$$f_{X+Y}(a) = \int_0^a 1 dz = a$$

If  $1 < a \leq 2$

$$f_{X+Y}(a) = \int_{a-1}^1 1 dz = 2-a$$

If  $a < 0$  or  $a > 2$ ,

$$f_{X+Y}(a) = 0.$$

i.e.

$$f_{X+Y}(a) = \begin{cases} a & \text{if } 0 < a \leq 1 \\ 2-a & \text{if } 1 < a \leq 2 \\ 0 & \text{otherwise.} \end{cases}$$

Example 2. Let  $X, Y$  be independent normal r.v.'s with parameters  $(0, 1)$  and  $(0, \sigma^2)$ , respectively. Find the density of  $X+Y$ .

Solution.

$$\begin{aligned}
 f_{X+Y}(a) &= f_X * f_Y(a) \\
 &= \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy \\
 &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{(a-y)^2}{2}} \cdot \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{y^2}{2\sigma^2}} dy \\
 &= \frac{1}{2\pi\sigma} \int_{-\infty}^{\infty} e^{-\frac{(a-y)^2}{2} - \frac{y^2}{2\sigma^2}} dy
 \end{aligned}$$

Notice that

$$\frac{(a-y)^2}{2} + \frac{y^2}{2\sigma^2} = \frac{\left(\sqrt{\sigma^2+1}y - \frac{a\sigma^2}{\sqrt{\sigma^2+1}}\right)^2}{2\sigma^2} + \frac{a^2}{2(\sigma^2+1)}$$

Hence

$$f_{X+Y}(a) = \frac{1}{2\pi\sigma} e^{-\frac{a^2}{2(\sigma^2+1)}} \int_{-\infty}^{\infty} e^{-\frac{(\sqrt{\sigma^2+1}y - \frac{a\sigma^2}{\sqrt{\sigma^2+1}})^2}{2\sigma^2}} dy$$

$$\left( \text{Letting } z = \frac{\sqrt{\sigma^2+1}y - \frac{a\sigma^2}{\sqrt{\sigma^2+1}}}{\sigma} \right)$$

$$= \frac{1}{2\pi\sqrt{\sigma^2+1}} e^{-\frac{a^2}{2(\sigma^2+1)}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz$$

$$= \frac{1}{\sqrt{2\pi(\sigma^2+1)}} e^{-\frac{a^2}{2(\sigma^2+1)}}.$$

Hence  $X+Y$  is a normal r.v. with parameters  $(0, \sigma^2+1)$ .

Remark: In general, if  $X, Y$  are ind. normal r.v.'s with parameters  $(\mu_1, \sigma_1^2), (\mu_2, \sigma_2^2)$ . Then

$X+Y$  is a norm r.v. with parameters  $(\mu_1+\mu_2, \sigma_1^2+\sigma_2^2)$

2. The Case that  $X, Y$  are both discrete.

In this case

$$\begin{aligned} P\{X+Y=k\} &= \sum_y P\{X=k-y\} P\{Y=y\} \\ &= \sum_y p_X(k-y) p_Y(y) \\ &= \sum_x p_X(x) p_Y(k-x). \end{aligned}$$

Example 3. Suppose that  $X, Y$  are independent

Poisson r.v's with parameters  $\lambda_1, \lambda_2$ , resp.

Find the distribution of  $X+Y$ .

Solution: Notice that  $X, Y$  both take values

in  $0, 1, 2, \dots$ . Hence the range of  $X+Y$   
is the set  $\{0, 1, 2, \dots\}$

For  $n=0, 1, 2, \dots$

$$P\{X+Y=n\} = \sum_{k=0}^n P\{X=k\} P\{Y=n-k\}$$

$$= \sum_{k=0}^n \cdot e^{-\lambda_1} \cdot \frac{\lambda_1^k}{k!} e^{-\lambda_2} \cdot \frac{\lambda_2^{n-k}}{(n-k)!}$$

$$= \frac{e^{-\lambda_1-\lambda_2}}{n!} \cdot \sum_{k=0}^n \cdot \binom{n}{k} \lambda_1^k \lambda_2^{n-k}$$

$$= \frac{e^{-\lambda_1-\lambda_2}}{n!} \cdot (\lambda_1 + \lambda_2)^n.$$

Hence,  $X+Y$  is a Poisson r.v. with parameter  $\lambda_1 + \lambda_2$ .

## § 6.4. Conditional distributions

### 1. Discrete case.

Def. Let  $X, Y$  be two discrete r.v.'s.

The conditional prob. mass function of  $X$  given

$Y=y$  is

$$\begin{aligned} p_{X|Y}(x|y) &:= P\{X=x \mid Y=y\} \\ &= \frac{P\{X=x, Y=y\}}{P\{Y=y\}} \\ &= \frac{p(x,y)}{p_Y(y)}, \end{aligned}$$

if  $P_Y(y) \neq 0$ .

Example 4. If  $X, Y$  are independent Poisson r.v's with parameters  $\lambda_1, \lambda_2$ . Calculate the conditional distribution of  $X$  given  $X+Y=n$ .

Solution:

$$\begin{aligned}
 & P\{X=k \mid X+Y=n\} \\
 &= \frac{P\{X=k, X+Y=n\}}{P\{X+Y=n\}} \\
 &= \frac{P\{X=k\} P\{Y=n-k\}}{P\{X+Y=n\}} \\
 &= \frac{e^{-\lambda_1} \frac{\lambda_1^k}{k!} \cdot e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!}}{e^{-\lambda_1-\lambda_2} \frac{(\lambda_1+\lambda_2)^n}{n!}}
 \end{aligned}$$

$$= \binom{n}{k} \left( \frac{\lambda_1}{\lambda_1+\lambda_2} \right)^k \left( \frac{\lambda_2}{\lambda_1+\lambda_2} \right)^{n-k}, \quad \text{for } k=0, 1, \dots, n.$$

Hence the conditional distribution of  $X$

Given  $X+Y=n$

is the binomial distribution with parameters

$n$  and  $\frac{\lambda_1}{\lambda_1+\lambda_2}$ .