

ENGG 2430 / ESTR 2004: Probability and Statistics
Spring 2019

10. Bayesian Statistics

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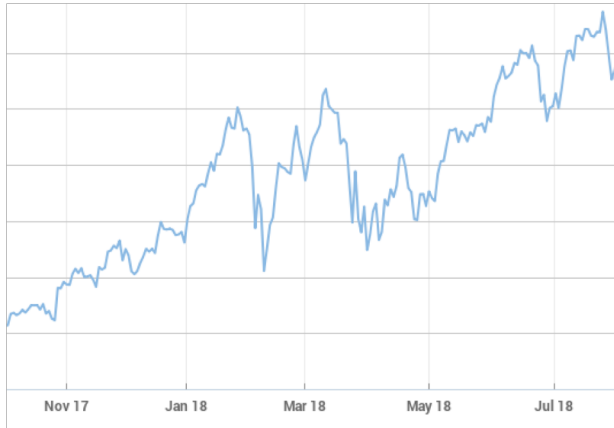
The Central Dogma of Statistics

**data = independent samples
from some random variable
(or several random variables)**

...but we don't know PDF/PMF



Poisson(λ)



Normal(μ, σ)

| | |
|----------------|-------------|
| Alice | PASS |
| Bob | PASS |
| Charlie | FAIL |

Binomial($200, p$)

Please pass
me the



?

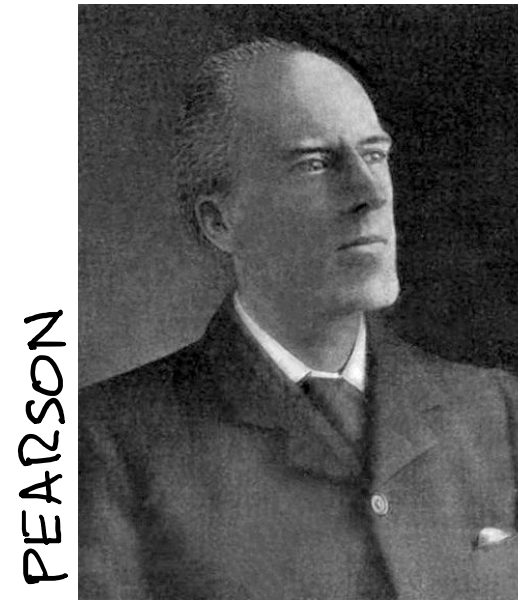
SALT 20%
BALL 30%
BAZOOKA 2%

OBAMA 1%
KIM 30%
XI 15%
MERCHEL 10%

parameters λ, μ, σ, p etc. are



BAYESIAN
RANDOM VAR.



UNKNOWN

Bayesian inference

1. Assign **prior probabilities** to params
2. **Observe** data
3. Update probabilities via **Bayes' rule**

Bayes' rule

$$f_{\Theta|X}(\theta | x) = \frac{f_{X|\Theta}(x | \theta) f_{\Theta}(\theta)}{f_X(x)}$$

POSTERIOR \nearrow

$\propto f_{X|\Theta}(x|\theta) f_{\Theta}(\theta)$ \longleftarrow PRIOR

$$f_{\Theta|X_1 \dots X_n}(\theta | x_1 \dots x_n) \propto f_{X_1|\Theta}(x_1 | \theta) \dots f_{X_n|\Theta}(x_n | \theta) f_{\Theta}(\theta)$$

if X_1, \dots, X_n are independent

Romeo is waiting for Juliet on their first date.



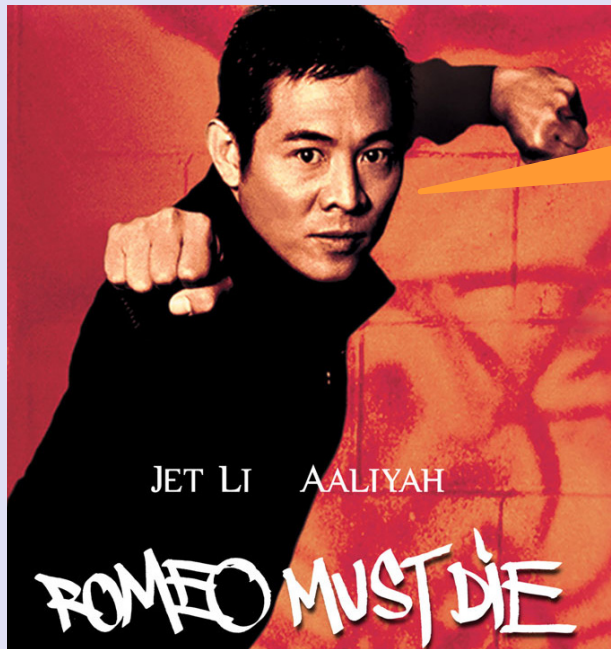
$X = \text{Uniform}(0, .3)$



$\text{Uniform}(0, .8)$



$\text{Uniform}(0, .6)$

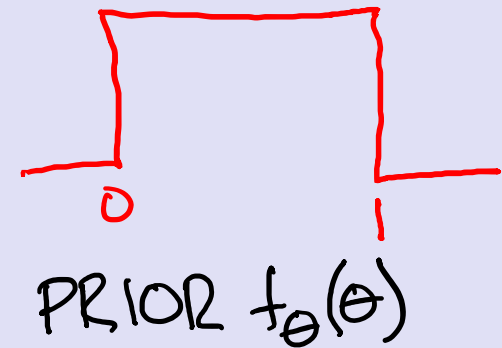


girls are $\text{Uniform}(0, \Theta)$ late

Romeo's model

$$X = \text{Uniform}(0, \Theta)$$

$$\Theta = \text{Uniform}(0, 1)$$



On her first date, Juliet arrives $\frac{1}{2}$ hour late.

$$\begin{aligned} f_{\theta|x}(\theta | x = \frac{1}{2}) &\propto f_{x|\theta}(\frac{1}{2} | \theta) f_{\theta}(\theta) \\ &= \frac{1}{\theta} \cdot 1 \quad \text{IF } \theta \geq \frac{1}{2} \end{aligned}$$

$$f_{\theta|x}(\theta | x = \frac{1}{2}) = \frac{\frac{1}{\theta}}{\int_{\frac{1}{2}}^1 \frac{1}{\theta'} d\theta'} = \frac{\frac{1}{\theta}}{-\ln \frac{1}{2}} = \frac{1}{\theta \ln 2}$$

POSTERIOR $f_{\theta|x}(\theta | \frac{1}{2})$

On her first 3 dates, Juliet is late by x_1, x_2, x_3 hours.

$$\begin{aligned}
 f_{\theta|x_1x_2x_3}(\theta|x_1, x_2, x_3) &\propto f_{x_1x_2x_3|\theta}(x_1, x_2, x_3|\theta) f_{\theta}(\theta) \\
 &= f_{x_1|\theta}(x_1|\theta) f_{x_2|\theta}(x_2|\theta) f_{x_3|\theta}(x_3|\theta) f_{\theta}(\theta) \\
 &= \frac{1}{\theta} \cdot \frac{1}{\theta} \cdot \frac{1}{\theta} \cdot 1 \\
 &\quad x_1 \leq \theta \quad x_2 \leq \theta \quad x_3 \leq \theta \\
 &= \begin{cases} \frac{1}{\theta^3} & \text{IF } x_1, x_2, x_3 \leq \theta \\ 0 & \text{IF NOT} \end{cases}
 \end{aligned}$$

EX. $x_1 = \frac{1}{2}, x_2 = x_3 = \frac{1}{4}$

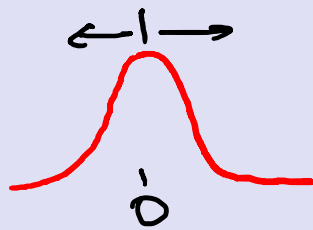
$$\int_{1/2}^1 \frac{c}{\theta^3} d\theta = 1 \rightarrow c = \frac{3}{8}$$



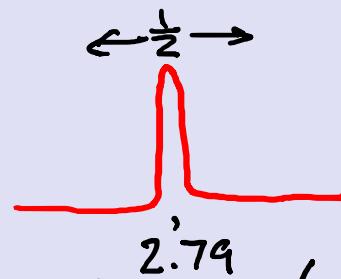
Three independent Normal(Θ , 1) RVs take values 3.97, 4.09, 3.11. What is Θ ?

PRIOR $\Theta = \text{Normal}(0, 1)$

$$\begin{aligned}
 f_{\Theta|x_1x_2x_3}(\Theta|x_1x_2x_3) &\propto f(x_1|\Theta)f(x_2|\Theta)f(x_3|\Theta)f(\Theta) \\
 &\propto e^{-\frac{(x_1-\Theta)^2}{2}} e^{-\frac{(x_2-\Theta)^2}{2}} e^{-\frac{(x_3-\Theta)^2}{2}} e^{-\frac{\Theta^2}{2}} \\
 &\propto e^{-\frac{\left(\frac{x_1+x_2+x_3+\Theta}{4}-\Theta\right)^2}{2 \cdot \left(\frac{1}{4}\right)^2}} \\
 &= \text{PDF OF Normal}\left(\frac{0+x_1+x_2+x_3}{4}, \sigma=\frac{1}{4}\right)
 \end{aligned}$$



PRIOR f_{Θ}



POSTERIOR $f_{\Theta|x_1x_2x_3}(\Theta|3.97, 4.09, 3.11)$

Inference for normals

$X_i = \text{Normal}(\Theta, \sigma_i)$ independent given Θ

Θ is $\text{Normal}(x_0, \sigma_0)$

$(\Theta \mid X_1 = x_1, \dots, X_n = x_n)$ is $\text{Normal}(x, \sigma)$ where

$$1/\sigma^2 = 1/\sigma_0^2 + \dots + 1/\sigma_n^2$$

$$x/\sigma^2 = \frac{x_0/\sigma_0^2 + \dots + x_n/\sigma_n^2}{n + 1}$$

A coin of unknown bias flips HHTH.

What is the bias?

PRIOR: BIAS θ IS Uniform $(0,1)$

$$\begin{aligned} f_{\theta|x}(\theta | \text{HHTH}) &\propto \frac{f(\text{HHTH}|\theta)}{x!} \cdot \frac{f(\theta)}{\theta} \\ &= \theta^3 (1-\theta) \cdot 1 \end{aligned}$$

The Beta(α , β) random variable

$$f_{\Theta}(\theta) = \frac{1}{\mathbf{B}(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1} \quad \text{when } 0 < \theta < 1$$

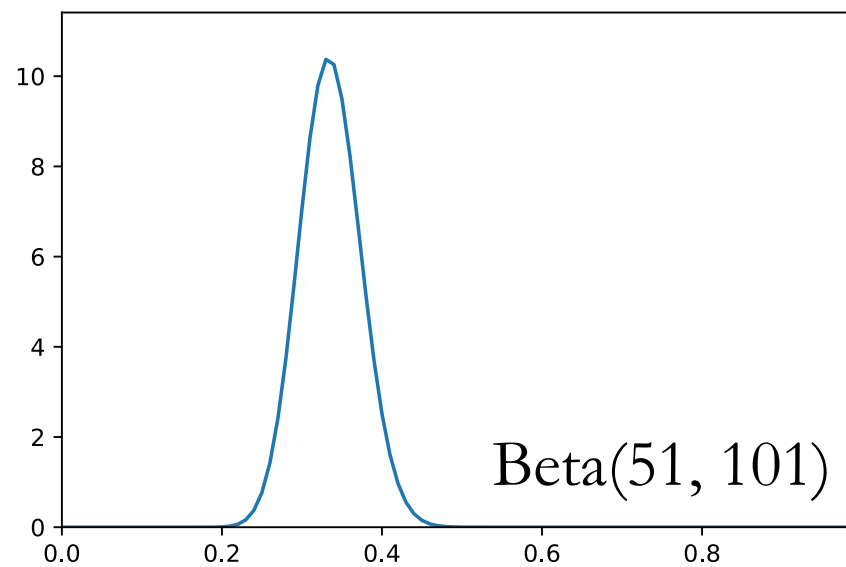
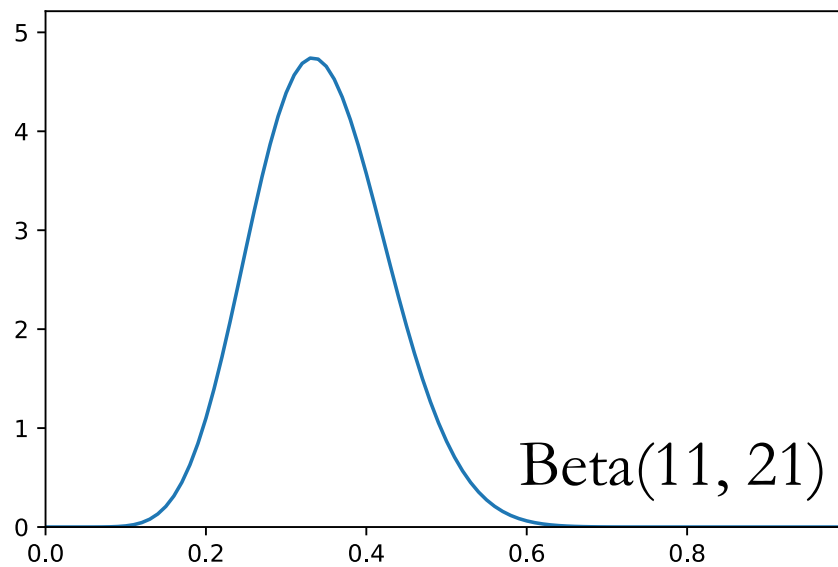
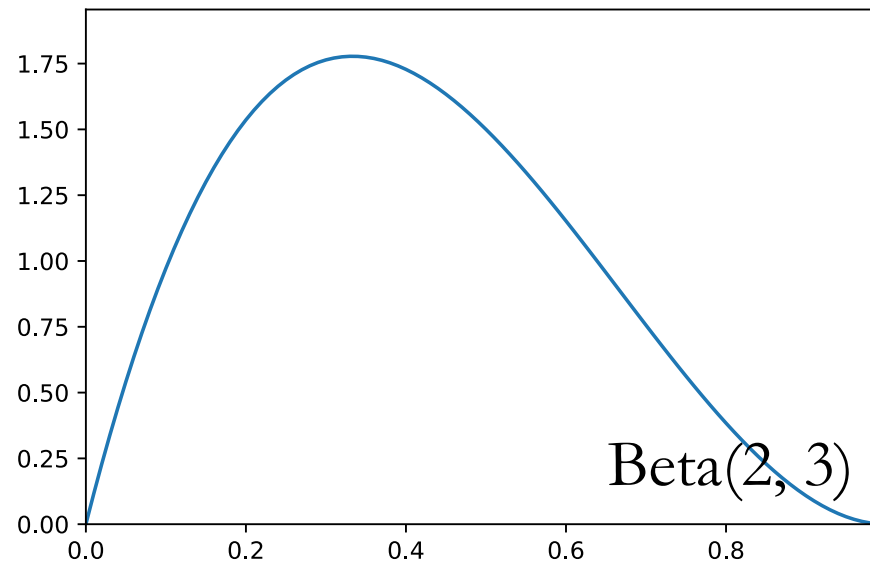
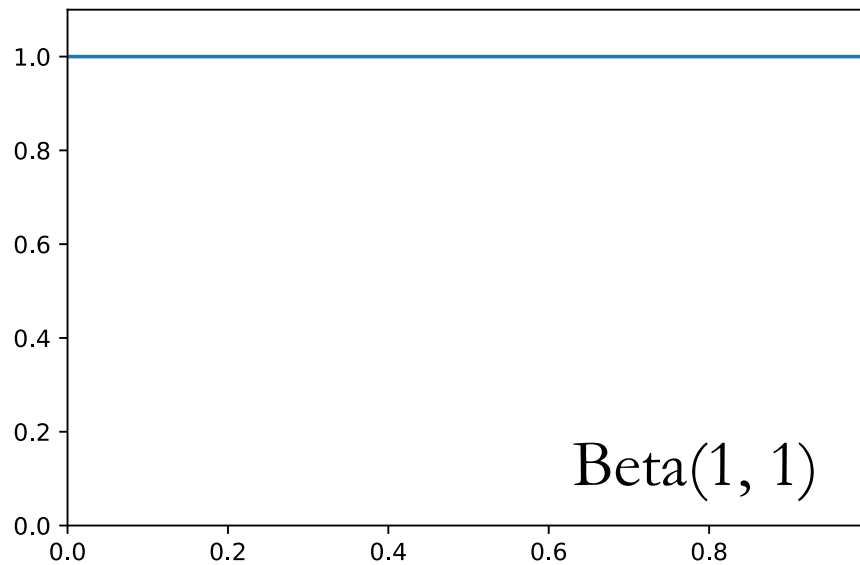
$$\mathbf{B}(\alpha, \beta) = (\alpha - 1)! (\beta - 1)! / (\alpha + \beta - 1)!$$

The Beta(α, β) random variable

$$\text{Beta}(1, 1) = \text{Uniform}(0, 1)$$

Θ is Beta(1, 1)

$(\Theta \mid b \text{ heads}, t \text{ tails})$ is Beta($1 + b, 1 + t$)



CONGRATULATIONS!! 3/02

from: F48E5F6BRT@vega.ocn.ne.jp

to: andrejb@cse.cuhk.edu.hk

Dear Customer,

My name is Sandra Davis, Board of Directors of United Nations and Chief Executive Officer, effective April 16, 2018,

The United Nations {UN} has giving you extra three working days to receive your fund from Citibank Plc, New York or you will lose the opportunity for ever. So you are advised to comply immediately to avoid the cancellation of your fund, follow the instruction immediately for your own good and future

The Citibank controlling department controlling of the security transfer CODE which is (CI201), the Authentication section code of this bank concludes the verification of your file. After going through all the documents of claim received by this department with justification and verification from the global strategy United States we are completely satisfied and you have been confirmed.

The Citibank concerning wire transfers of your fund. Your letter has been referred to the (JMCB) Legal Division for Funds (US\$2.8 Million Dollars) Transferred code (.). We are satisfied using Electronic Wire Transfer or Swift Wire Transfer and the rights and liabilities of using of electronic and Swift fund transfer systems are defined by the Electronic Fund Transfer Act... The regulation, however, which implements this statute, Regulation E. specifically states that its provisions are inapplicable to a situation such we must ensure your Funds Transferred to your destination Bank Account between 72 hours.

Considering the volume of your payment, it is right for us to seek for the approval of some money regulatory Boards here in United States before we can carry out the Transfer of an amount of such magnitude to anybody, otherwise any such transfer will be stopped by the Authorities, and the International Monetary Fund (IMF), since your

Transfer is Electronic Transfer or Swift Wire transfer is almost activated with our bank and the only thing holding the final activation of your Account are some certain Approval Documents from the concerned Authorities here in United States

NB: THIS TRANSACTION IS BEING MONITORED BY THE UNITED STATES GOVERNMENT IN ORDER TO GUARDS US FROM INTERNET IMPOSTORS.

Provide your designated bank account details for Electronic Transfer, to avoid mistake(s).

Bank Name and Address

Account Number:

Account Name:

Routing Number:

your home address and phone number,
place of work and address.

send it the citibank remittance manager. on her email
: ombes2@gmx.com

UN gives you only 3 working days to receive your fund from our bank or no more so follow the instruction by sending email to us back with the bank detail details along with your personal details.

Thank you for giving us the opportunity to serve your banking needs. ombes2@gmx.com

Yours sincerely

Board of Directors of Citibank

Sandra Davis

Chief Executive Officer, effective April 16, 2018

Θ = spam indicator = $\begin{cases} 1 & \text{IF SPAM} \\ 0 & \text{IF NOT} \end{cases}$

X_1 = contains "million dollars"

X_2 = contains "Nigerian princess"

$$P(X_1 = 1 | \Theta = 1) = 10\%$$

$$P(X_1 = 1 | \Theta = 0) = 3\%$$

$$P(X_2 = 1 | \Theta = 1) = 1\%$$

$$P(X_2 = 1 | \Theta = 0) = 0.01\%$$

$$P(\Theta = 1) = 20\%$$

$$P(\Theta = 0) = 80\%$$

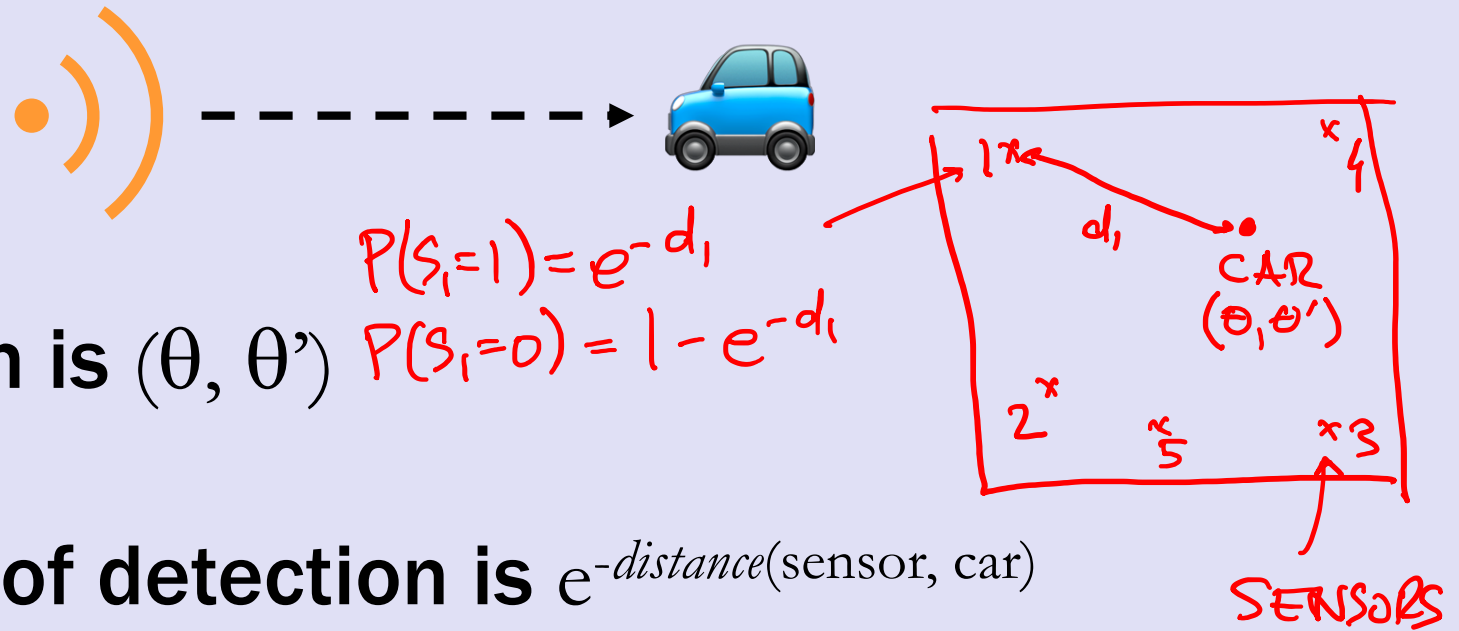
ASSUME

X_1, X_2 INDEPENDENT
GIVEN Θ .

$X_1 = 1$ OBSERVED "MILLION \$"
 $X_2 = 0$ DID NOT SEE "NIGERIAN PRINCESS"

$$\begin{aligned} P(\theta = 1 \mid X_1 = 1, X_2 = 0) &\propto P(X_1 = 1, X_2 = 0 \mid \theta = 1) P(\theta = 1) \\ &= P(X_1 = 1 \mid \theta = 1) P(X_2 = 0 \mid \theta = 1) P(\theta = 1) \\ &= 0.1 \cdot 0.99 \cdot 0.2 \end{aligned}$$

$$\begin{aligned} P(\theta = 0 \mid X_1 = 1, X_2 = 0) &\propto P(X_1 = 1 \mid \theta = 0) P(X_2 = 0 \mid \theta = 0) P(\theta = 0) \\ &= 0.03 \cdot 0.9999 \cdot 0.8 \end{aligned}$$



Car position is (θ, θ')

Probability of detection is $e^{-\text{distance}(\text{sensor}, \text{car})}$

There are 5 sensors at different positions

Given that sensors 1, 3, 4 reported detection and 2, 5 didn't, where is the car?

PROBABILITY MODEL

ASSUME S_1, \dots, S_5 ARE INDEPENDENT GIVEN θ, θ'

PRIOR: θ, θ' INDEPENDENT Normal(0, 1)

$S_1 = S_3 = S_4 = 1, S_2 = S_5 = 0$: WHAT IS θ, θ' ?

$$f_{\theta\theta'}(\theta, \theta' \mid S_1 = S_3 = S_4 = 1, S_2 = S_5 = 0)$$

$$\propto P(S_1 = S_3 = S_4 = 1, S_2 = S_5 = 0 \mid \theta = \theta, \theta' = \theta') \cdot f_{\theta\theta'}(\theta, \theta')$$

$$= P(S_1 = 1 \mid \theta = \theta, \theta' = \theta') P(S_2 = 0 \mid \theta = \theta, \theta' = \theta') \dots \cdot f_{\theta\theta'}(\theta, \theta')$$

$$= e^{-\sqrt{(x_1 - \theta)^2 + (y_1 - \theta')^2}} \cdot (1 - e^{-\sqrt{(x_2 - \theta)^2 + (y_2 - \theta')^2}}) \cdot \dots \cdot \frac{1}{2\pi} e^{-(\theta^2 + \theta'^2)/2}$$

TO CALCULATE VALUE, DIVIDE BY

$$\iint_{-\infty}^{\infty} e^{-\sqrt{(x_1 - \theta)^2 + (y_1 - \theta')^2}} \cdot (1 - e^{-\sqrt{(x_2 - \theta)^2 + (y_2 - \theta')^2}}) \cdot \dots \cdot \frac{1}{2\pi} e^{-(\theta^2 + \theta'^2)/2} d\theta d\theta'$$

Point estimation

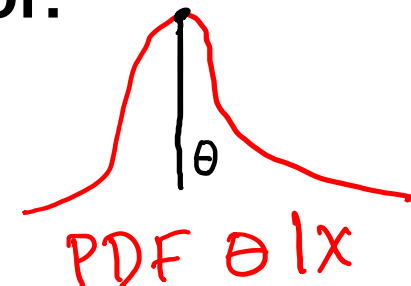
How to turn conditional PDF/PMF $f_{\Theta|X}(\theta | x)$ estimate into one number?

Conditional expectation (CE) estimator:

$$E[\theta | X = x]$$

Maximum *a posteriori* (MAP) estimator:

$$\operatorname{argmax}_{\theta} f_{\Theta|X}(\theta | x)$$



Point estimation for normals

$X_i = \text{Normal}(\Theta, 1)$ independent given Θ

Θ is $\text{Normal}(x_0, 1)$

$(\Theta \mid X_1 = x_1, \dots, X_n = x_n)$ is $\text{Normal}(x, 1/\sqrt{n})$

$$x = \frac{x_0 + x_1 + \dots + x_n}{n+1}$$

CE estimate: $E[\Theta \mid X_1 = x_1, \dots, X_n = x_n] = x$

MAP estimate: x

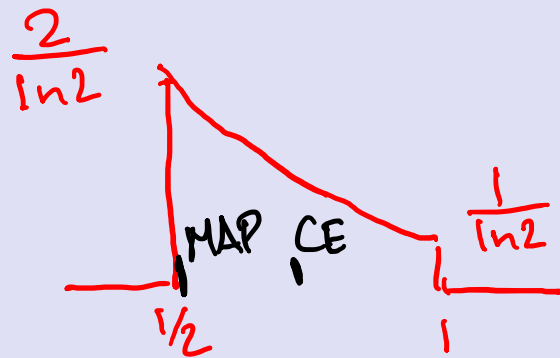


Romeo's model

$$X = \text{Uniform}(0, \Theta)$$

$$\Theta = \text{Uniform}(0, 1)$$

On her first date, Juliet arrives $\frac{1}{2}$ hour late.



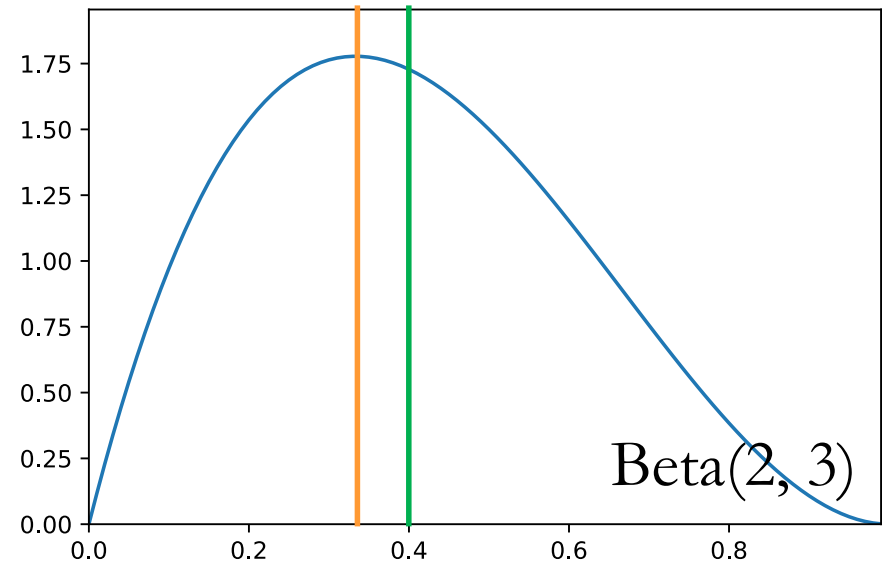
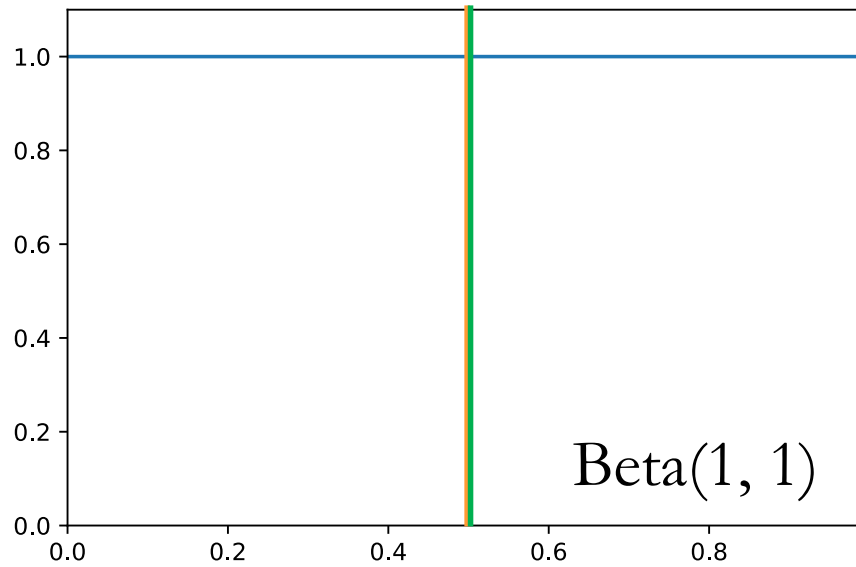
$$f_{\Theta|X}(\theta | \frac{1}{2}) = \frac{1}{\theta \cdot \ln 2}$$

CE estimate:

$$E[\Theta | X = \frac{1}{2}] = \int_{\frac{1}{2}}^1 \theta' \frac{1}{\theta' \ln 2} d\theta' = \frac{1}{2 \ln 2} \approx 0.72$$

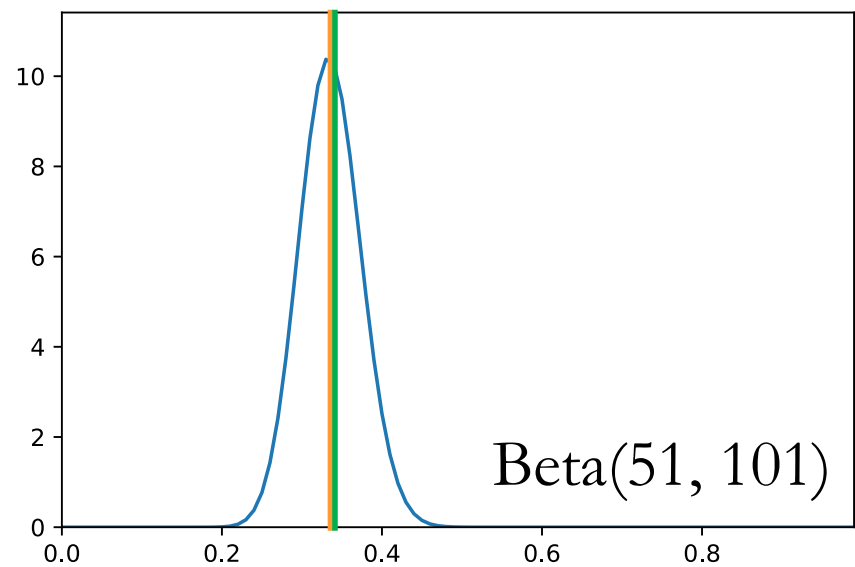
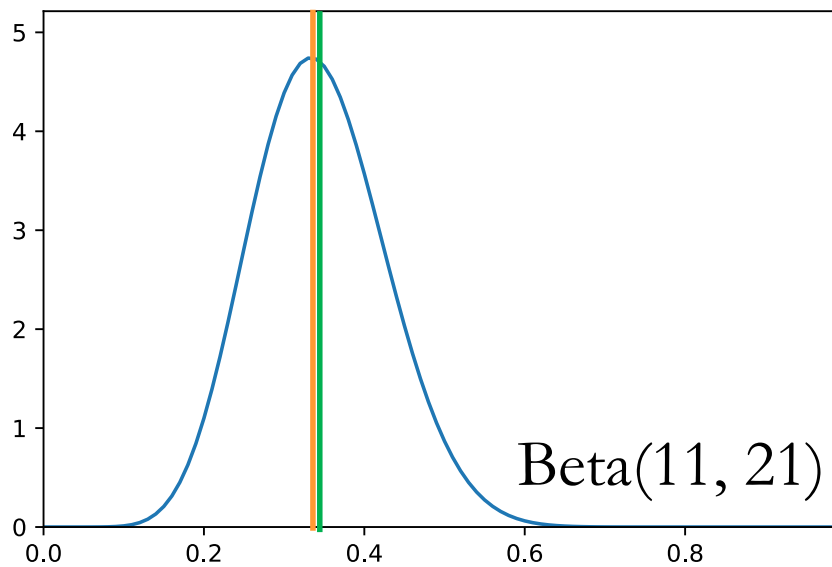
MAP estimate:

$$\frac{1}{2}$$



$$\text{CE} = \alpha / \beta = (k + 1) / (n + 2)$$

$$\text{MAP} = k / n$$



Hypothesis testing

Suppose Θ takes two values (e.g. spam / legit)

$$\mathbf{MAP} = \operatorname{argmax}_{\theta} f_{\Theta|X}(\theta | x)$$

Choose the one for which $f_{\Theta|X}(\theta | x)$ is larger

$\Theta = 80\%$ legit, 20% spam

| θ | $\mathbf{P}(X_1 \theta)$ | $\mathbf{P}(X_2 \theta)$ |
|----------|----------------------------|----------------------------|
| legit | 0.03 | 0.0001 |
| spam | 0.1 | 0.01 |

The Citibank concerning wire transfers of your fund. Your letter has been referred to the (JMCB) Legal Division for Funds (US\$2.8 Million Dollars)

$$P(\text{SPAM} | X_1=1, X_2=0) = 0.1 \cdot 0.99 \cdot 0.2 \approx 0.0198$$

$$P(\text{LEGIT} | X_1=1, X_2=0) = 0.03 \cdot 0.9999 \cdot 0.8 \approx 0.0240$$

MAP ESTIMATE: LEGIT

Coin A is heads with probability $1/3$.

Coin B is tails with probability $1/3$.

HHHT are 4 flips of a random coin. Which coin was it?

PRIOR: $\Theta = A$ WITH PROB $1/2$
 B WITH PROB $1/2$

$$P(A | \text{HHHT}) \propto P(\text{HHHT} | A) \cdot P(A)$$

$$= \left(\frac{1}{3}\right)^3 \cdot \frac{2}{3} \cdot \frac{1}{2}$$

MAP: COIN B.

$$P(B | \text{HHHT}) \propto \left(\frac{2}{3}\right)^3 \cdot \frac{1}{3} \cdot \frac{1}{2}$$

MAP: MORE HEADS $\rightarrow B$

MORE TAILS $\rightarrow A$

2H, 2T

$\rightarrow A$

(DOESN'T MATTER)

What is the probability you are wrong, given the outcome is HHHT?

W = WRONG DECISION

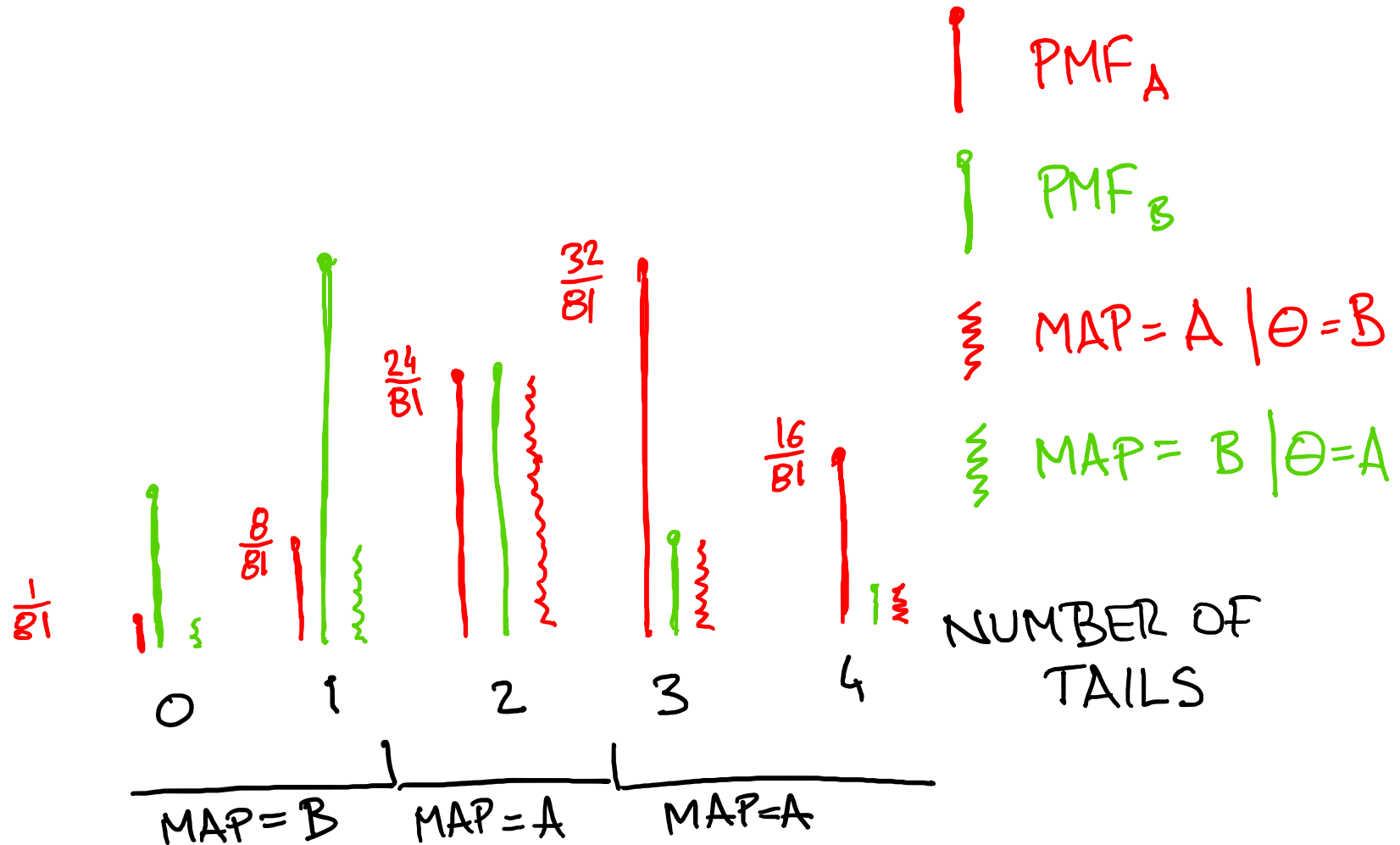
$$\begin{aligned} P(W | HHHT) &= P(\theta = A | HHHT) \\ &\propto P(HHHT | \theta = A) P(\theta = A) \\ &= \left(\frac{1}{3}\right)^3 \cdot \frac{2}{3} \cdot \frac{1}{2} \end{aligned}$$

$$P(W | HHHT) = \frac{\left(\frac{1}{3}\right)^3 \cdot \frac{2}{3} \cdot \frac{1}{2}}{\left(\frac{1}{3}\right)^3 \cdot \frac{2}{3} \cdot \frac{1}{2} + \left(\frac{2}{3}\right)^3 \cdot \frac{1}{3} \cdot \frac{1}{2}} = \frac{1}{1+4} = 20\%$$

What is the probability you are wrong **on average**?

$$\begin{aligned} P(W) &= P(\text{MAP} = A, \theta = B) + P(\text{MAP} = B, \theta = A) \\ &= \frac{1}{2} \cdot P(\text{MAP} = A | \theta = B) + \frac{1}{2} P(\text{MAP} = B | \theta = A) \\ &= \frac{1}{2} \left(\binom{1}{2}^4 + 4 \cdot \frac{2}{3} \cdot \left(\frac{1}{3}\right)^3 + 6 \cdot \left(\frac{2}{3}\right)^2 \cdot \left(\frac{1}{3}\right)^2 \right) + \frac{1}{2} \left(\left(\frac{1}{3}\right)^1 + 4 \cdot \left(\frac{1}{3}\right)^3 \cdot \frac{2}{3} \right) \\ &= \frac{21}{81} \approx 26\% \end{aligned}$$

Hypothesis testing error



An car-jack **detector** X outputs $\text{Normal}(0, 1)$ if there is no intruder and $\text{Normal}(1, 1)$ if there is. When should **alarm** activate?

MODEL: $\Theta = \begin{cases} 1 \text{ (INTRUDER)} & p = 10\% \\ 0 \text{ (LEGITIMATE)} & 1-p = 90\% \end{cases}$

$$f_{x|\theta}(x|0) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

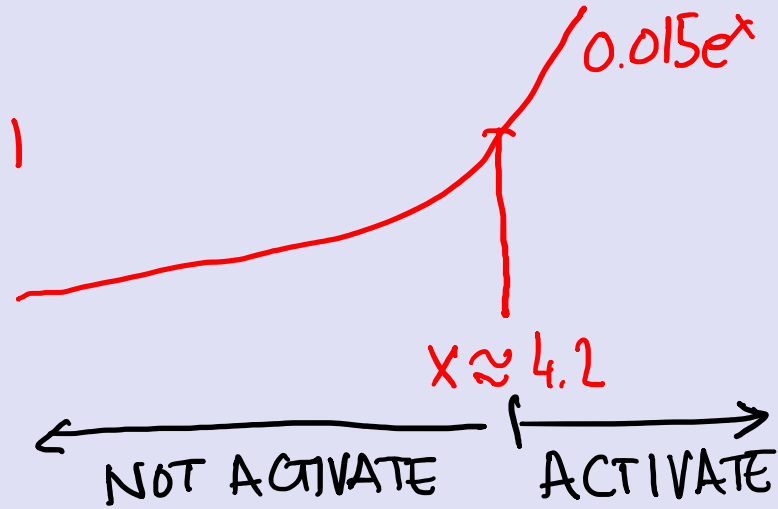
$$f_{x|\theta}(x|1) = \frac{1}{\sqrt{2\pi}} e^{-(x-1)^2/2}$$

POSTERIORS: $P(\theta=1|X=x) \propto f_{x|\theta}(x|1) \cdot P(\theta=1) \propto e^{-(x-1)^2/2} \cdot p$
 $P(\theta=0|X=x) \propto f_{x|\theta}(x|0) \cdot P(\theta=0) \propto e^{-x^2/2} \cdot (1-p)$

MAP: INTRUSION IF $P(\theta=1|X=x) > P(\theta=0|X=x)$
 $\frac{P(\theta=1|X=x)}{P(\theta=0|X=x)} = \frac{e^{-(x-1)^2/2}}{e^{-x^2/2}} \cdot \frac{p}{1-p} = e^{x-\frac{1}{2}} \cdot \frac{p}{1-p}$

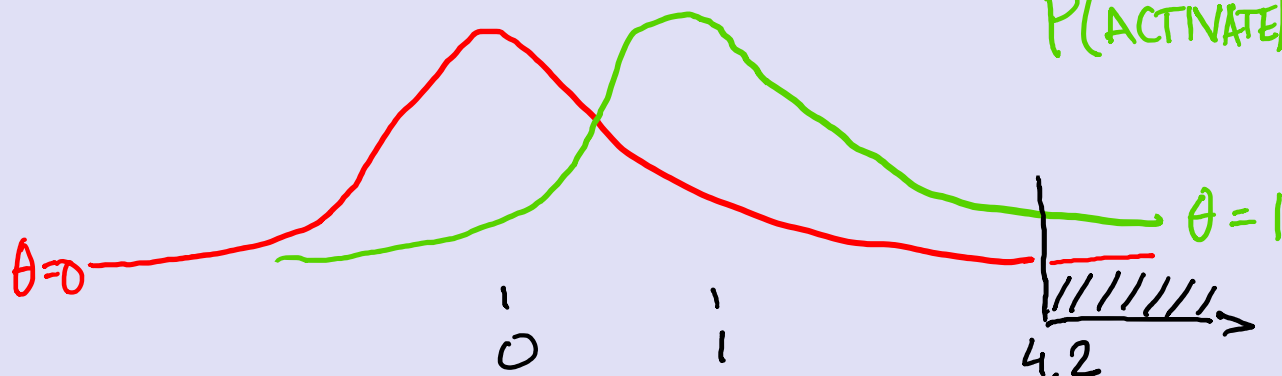
WANT TO KNOW IF

$$\frac{e^{x - \frac{1}{2}} \cdot \frac{0.1}{0.9}}{0.015e^x} > 1$$
$$0.015e^x > 1$$



$$P(\text{ACTIVATE} | \theta = 0) = P(\text{Normal}(0, 1) > 4.2)$$
$$\approx 0$$

$$P(\text{ACTIVATE} | \theta = 1) = P(\text{Normal}(1, 1) > 4.2)$$
$$\approx 0.0007$$



$$P(\text{ERROR}) \approx 0.9 \cdot 0 + 0.1 \cdot 99.93$$
$$\approx 0.1$$