

2 decks of cards

1 deck Red deck: 28 red cards
24 black cards

Black deck: 28 black cards
24 red cards

R B B R R B R

① conditioned on "red deck" what is the
prob of this sequence $\left(\frac{28}{52}\right)^4 \left(\frac{24}{52}\right)^3$

conditioned on "black deck" what is the
prob. of this sequence $\left(\frac{24}{52}\right)^4 \left(\frac{28}{52}\right)^3$

→ ~~Let~~ Let $\theta = 0$ if "red deck"
 $= 1$ if "black deck"

$$x_1, \dots, x_n \quad P(x_1, \dots, x_n | \theta = 0) = \left(\frac{28}{52}\right)^4 \left(\frac{24}{52}\right)^3$$

$$P(x_1, \dots, x_n | \theta = 1) = \left(\frac{24}{52}\right)^4 \left(\frac{28}{52}\right)^3$$

$$P(\theta = 0 | x_1, \dots, x_n) = \frac{P(\theta = 0, x_1, \dots, x_n)}{P(x_1, \dots, x_n)}$$

$$= \frac{P(\theta = 0, x_1, \dots, x_n)}{P(\theta = 0, x_1, \dots, x_n) + P(\theta = 1, x_1, \dots, x_n)}$$

$$= \frac{P(\theta = 0) P(x_1, \dots, x_n | \theta = 0)}{P(\theta = 0) P(x_1, \dots, x_n | \theta = 0) + P(\theta = 1) P(x_1, \dots, x_n | \theta = 1)}$$

$$P(\theta=0) = \frac{1}{2}$$

(2)

$$\begin{aligned} P(\theta=0 | x_1, \dots, x_7) &= \frac{(28)^4 (24)^3 \cdot \frac{1}{2}}{\frac{1}{2} [(28)^4 (24)^3 + (28)^3 (24)^4] \cdot \frac{1}{2}} \\ &= \frac{28}{52} \end{aligned}$$

$$P(\theta=1 | x_1, \dots, x_7) = \frac{24}{52}$$

R B B R R B R B B

$$\begin{aligned} P(\theta=0 | x_1, \dots, x_9) &= \frac{(28)^4 (24)^5}{(28)^4 (24)^5 + (28)^5 (24)^4} \\ &= \frac{24}{52} \end{aligned}$$

$$f_{\theta|x}(\theta|x) = \frac{f(x, \theta)}{f(x)} = \frac{f(\theta) f(x|\theta)}{f(x)}$$

$$\frac{f_{\theta|x}(\theta=i|x)}{f_{\theta|x}(\theta=j|x)} = \frac{f(\theta=i) f(x|\theta=i)}{f(\theta=j) f(x|\theta=j)}$$

if conditioned on θ , x_1, \dots, x_n are independent

$$f(x_1, \dots, x_n | \theta) = \prod_{k=1}^n f(x_k | \theta)$$

$$\therefore \frac{f_{\theta|x}(\theta=i|\vec{x})}{f_{\theta|x}(\theta=j|\vec{x})} = \frac{f(\theta=i) f(x_1, \dots, x_n | \theta=i)}{f(\theta=j) f(x_1, \dots, x_n | \theta=j)}$$

$$\begin{aligned}
 & \textcircled{3} \\
 & = \frac{f(\theta=i) \prod_k f(x_k | \theta=i)}{f(\theta=j) \prod_k f(x_k | \theta=j)} \\
 & = \frac{f(\theta=i)}{f(\theta=j)} \prod_k \frac{f(x_k | \theta=i)}{f(x_k | \theta=j)}
 \end{aligned}$$

$\theta = 1$ uniform $[0, 18 \text{ minutes}]$
 $\theta = 2$ uniform $[0, 48 \text{ minutes}]$
 $\theta = 3$ uniform $(0, 36 \text{ minutes})$

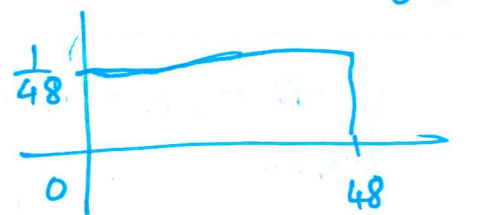
$x_1 = 30 \text{ minutes}$

$\textcircled{4}$ Prior $\rightarrow P(\theta=1) = P(\theta=2) = P(\theta=3) = \frac{1}{3}$

$$P(\theta=1 | x_1=30 \text{ minutes}) = \frac{P(\theta=1) \cdot f(x_1=30 | \theta=1)}{P(x_1=30)} = 0$$

$$P(\theta=2, f(x_1=30 | \theta=2)) = \frac{1}{48}$$

$$f(x_1=30 | \theta=3) = \frac{1}{36}$$



$$\frac{P(\theta=2 | x_1=30)}{P(\theta=3 | x_1=30)} = \frac{\frac{1}{48}}{\frac{1}{36}} = \frac{36}{48}$$

$$\begin{aligned}
 P(\theta=2 | x_1=30) &= \frac{P(\theta=2) \cdot P(x_1=30 | \theta=2)}{P(\theta=3) \cdot P(x_1=30 | \theta=3) + P(\theta=2) \cdot P(x_1=30 | \theta=2)} \\
 &= \frac{\frac{1}{48}}{\frac{1}{48} + \frac{1}{36}} = \frac{36}{84} = \frac{3}{7} \\
 P(\theta=3 | x_1=30) &= \frac{4}{7}
 \end{aligned}$$

Puzzle

Two envelopes, one contains number x , another contains number y . I don't know x or y .

Goal. Open an envelope, look at the value and then decide whether to switch or not?

Qn: Can you devise a strategy that beats an expected value of $\frac{x+y}{2}$.

Inference for Normals

$$f(\theta | x_1 = x_1, \dots, x_n = x_n) = f(\theta | x_1, x_2)$$

$$= \frac{f(\theta) f(x_1 = x_1, x_2 = x_2 | \theta)}{f(x_1 = x_1, x_2 = x_2)}$$

$$f(\theta = \theta, x_1 = x_1, x_2 = x_2) = f(\theta) f_{x_1 x_2 | \theta}(x_1 = x_1, x_2 = x_2 | \theta)$$

$$= \frac{1}{\sqrt{2\pi\sigma_0^2}} e^{-\frac{(\theta - \mu_0)^2}{2\sigma_0^2}} \cdot \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{(x_1 - \mu_1)^2}{2\sigma_1^2}} \cdot \frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{(x_2 - \mu_2)^2}{2\sigma_2^2}}$$

$$f(x_1=x_1, x_2=x_2) = \int_{-\infty}^{\infty} f(\theta=\theta, x_1=x_1, x_2=x_2) d\theta$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{1}{\sigma_0} e^{-\frac{1}{2} \left[\frac{\theta^2}{\sigma_0^2} + \frac{\theta^2}{\sigma_1^2} + \frac{\theta^2}{\sigma_2^2} - \frac{2\theta x_0}{\sigma_0^2} - \frac{2\theta x_1}{\sigma_1^2} - \frac{2\theta x_2}{\sigma_2^2} \right]} (2\pi)^{3/2} \sqrt{\sigma_0^2 \sigma_1^2 \sigma_2^2} d\theta$$

$$\rightarrow x e^{-\frac{1}{2} \left[\frac{x_0^2}{\sigma_0^2} + \frac{x_1^2}{\sigma_1^2} + \frac{x_2^2}{\sigma_2^2} \right]} d\theta$$

$$= \frac{e^{-\frac{1}{2} \left[\frac{x_0^2}{\sigma_0^2} + \frac{x_1^2}{\sigma_1^2} + \frac{x_2^2}{\sigma_2^2} \right]}}{(2\pi)^{3/2} \sqrt{\sigma_0^2 \sigma_1^2 \sigma_2^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} \left[\theta^2 \left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right) - 2 \left(\frac{\theta x_0}{\sigma_0^2} + \frac{\theta x_1}{\sigma_1^2} + \frac{\theta x_2}{\sigma_2^2} \right) \right]} d\theta$$

$$= \frac{e^{-\frac{1}{2} \left[\frac{x_0^2}{\sigma_0^2} + \frac{x_1^2}{\sigma_1^2} + \frac{x_2^2}{\sigma_2^2} \right]}}{(2\pi)^{3/2} \sqrt{\sigma_0^2 \sigma_1^2 \sigma_2^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} \left[\theta - \frac{\left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right)}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)} \right]^2}$$

$$\times e^{-\frac{1}{2} \frac{\left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right)^2}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)}} d\theta$$

$$= \sqrt{2\pi} \frac{1}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)} \times \frac{e^{-\frac{1}{2} \left[\frac{x_0^2}{\sigma_0^2} + \frac{x_1^2}{\sigma_1^2} + \frac{x_2^2}{\sigma_2^2} \right]}}{(2\pi)^{3/2} \sqrt{\sigma_0^2 \sigma_1^2 \sigma_2^2}} \times$$

$$e^{-\frac{1}{2} \frac{\left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right)^2}{\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2}}}$$

(6)

$$f(\Theta=0 | X_1=x_1, X_2=x_2) = \frac{f(\Theta=0, X_1=x_1, X_2=x_2)}{f(X_1=x_1, X_2=x_2)}$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \frac{(\Theta - \mu)^2}{\sigma^2}}$$

where $\frac{1}{\sigma^2} = \frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2}$

and $\frac{\mu}{\sigma^2} = \frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2}$



$$\frac{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right) \Theta - \left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right)}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)}$$

$$\frac{\left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right) - \left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right) \Theta}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)}$$

$$\frac{\left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right) - \left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right) \Theta}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)}$$

$$\frac{\left(\frac{x_0}{\sigma_0^2} + \frac{x_1}{\sigma_1^2} + \frac{x_2}{\sigma_2^2} \right) - \left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right) \Theta}{\left(\frac{1}{\sigma_0^2} + \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)}$$

Two coins θ

(7)

HHTH

$$C_1: P_T(H) = \frac{1}{3}$$

$$C_2: P_T(T) = \frac{1}{3}$$

what is ^{the} prior that yields

$$P(\theta=1 | HHTH) = P(\theta=2 | HHTH)$$

$$\text{let } P(\theta=1) = x$$

$$\text{then } P(\theta=1, HHTH) = x \cdot \left(\frac{1}{3}\right)^3 \frac{2}{3}$$

$$P(\theta=2, HHTH) = (1-x) \left(\frac{2}{3}\right)^3 \frac{1}{3}$$

$$P(HHTH) = x \left(\frac{1}{3}\right)^3 \cdot \frac{2}{3} + (1-x) \left(\frac{2}{3}\right)^3 \frac{1}{3}$$

$$\text{want } P(\theta=1 | HHTH) = P(\theta=2 | HHTH)$$

$$\Leftrightarrow P(\theta=1, HHTH) = P(\theta=2, HHTH)$$

$$\therefore x \left(\frac{1}{3}\right)^3 \frac{2}{3} = (1-x) \left(\frac{2}{3}\right)^3 \frac{1}{3}$$

$$\Rightarrow 2x = (1-x) 8 \quad \Leftrightarrow \frac{x}{1-x} = 4$$

$$\text{or } x = \frac{4}{5}$$

Examples

Let us consider 3 die.

$$P(D=1) = \frac{1}{4}, \quad P(D=2) = \frac{1}{3}, \quad P(D=3) = \frac{5}{12}$$

Prob (when $D=1$, prob of $\{1, \dots, 6\}$ are $\{\frac{1}{3}, 0, \frac{1}{3}, 0, \frac{1}{3}, 0\}$)

when $D=2$, prob of $\{1, \dots, 6\}$ are $\{\frac{1}{8}, \frac{1}{8}, \frac{1}{8}, \frac{1}{8}, \frac{1}{4}, \frac{1}{4}\}$

when $D=3$, prob of $\{1, 2, \dots, 6\}$ are $\{\frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}\}$

updated prior after seeing (8)

3 2 6 2 4

$$P(\theta=1 | 3,2,6,2,4) = 0$$

$$P(\theta=2 | 3,2,6,2,4) \propto \frac{1}{3} \times \left(\frac{1}{8}\right)^4 \times \frac{1}{4} = \frac{\frac{1}{3 \times 8^4 \times 4}}{\frac{1}{3 \times 8^4 \times 4} + \frac{5}{6^6 \times 2}}$$

$$P(\theta=3 | 3,2,6,2,4) \propto \frac{5}{12} \times \left(\frac{1}{6}\right)^5$$

$$\rightarrow = \frac{\frac{1}{3072}}{\frac{1}{3072} + \frac{5}{27 \times 16}} = \frac{243}{883}$$

new

0

$$\frac{243}{883}$$

$$\frac{640}{883}$$

$$\frac{1}{4}$$

$$\frac{1}{3}$$

$$\frac{5}{12}$$

Maximum Likelihood Estimator

$$\operatorname{argmax}_{\theta} f_{X|\theta}(x|\theta)$$

VS

MAP Estimator

$$\operatorname{argmax}_{\theta} f_{\theta|X}(\theta|x)$$