1. The joint PDF of X and Y is

$$f_{X,Y}(x,y) = \begin{cases} C(x+y+1)y, & \text{if } 0 \le x \le 2, 0 \le y \le 2, \\ 0, & \text{otherwise.} \end{cases}$$

Find (a) the value of C and (b) The conditional PDF  $f_{Y|X}(y|x)$ .

## Solution:

(a) The PDF must integrate to one, so

$$1 = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f_{X,Y}(x,y) dx dy = \int_{0}^{2} \int_{0}^{2} C(x+y+1)y dx dy = \frac{40}{3}C.$$

Therefore  $C = \frac{3}{40}$ .

(b)  $f_X(x) = \int_{-\infty}^{+\infty} f_{X,Y}(x,y) dy = \int_0^2 C(x+y+1)y dy = C(2x+\frac{14}{3})$ . Using the convolution formula,

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{(x+y+1)y}{2x+\frac{14}{3}}$$

- 2. Alice and Bob agree to meet. Alice's arrival time A is uniform between 12:00 and 12:45 and Bob's arrival time B is uniform between 12:15 and 13:00. Let E be the event "Alice and Bob arrive within 30 minutes of one another".
  - (a) What is P(E) assuming A and B are independent?
  - (b) If you don't know the joint PDF of A and B, how large can P(E) be?
  - (c) (Optional) If you don't know the joint PDF of A and B, how small can P(E) be?

## Solution:

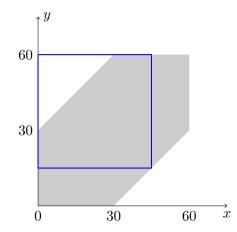
(a) We can model A as a Uniform (0, 45) random variable and B as an independent Uniform (15, 60) random variable. Their marginal PMFs are:

$$f_A(x) = \begin{cases} 1/45 & \text{if } 0 \le x < 45\\ 0 & \text{otherwise} \end{cases}$$
$$f_B(y) = \begin{cases} 1/45 & \text{if } 15 \le y < 60\\ 0 & \text{otherwise} \end{cases}$$

Because A and B are independent, their joint PMF is

$$f(x,y) = f_A(x)f_B(y) = \begin{cases} 1/2025 & \text{if } 0 \le x < 45 \text{ and } 15 \le y < 60\\ 0 & \text{otherwise} \end{cases}$$

The event E is |B - A| < 30. We want to calculate  $P(E) = \int_E f(x, y) dx dy$ .



To understand this expression look at the above picture. The blue square consists of the points (x, y) where f(x, y) is nonzero (it equals 1/2025). The grey strip are the points x, y such that  $|x - y| \le 30$ . The intersection of the two consists of the blue square minus the triangle T with vertices (0, 30), (30, 30), (0, 60) so

$$\mathbf{P}(E) = \int_{E} f(x, y) dx dy = 1 - \int_{T} \frac{1}{2025} dx dy = 1 - \frac{\operatorname{area}(T)}{2025} = 1 - \frac{1}{2025} \times 450 = \frac{7}{9}.$$

- (b) If B = A + 15, then the marginal PMFs of A and B are as in part (a), but P(E) = P(|B A| < 30) = P(15 < 30) = 1.
- (c) Let  $E_A$  be the event 15 < A < 30 and  $E_B$  be the event 15 < B < 30. By the axioms of probability,

$$P(E) \ge P(E_A \cap E_B) = P(E_A) + P(E_B) - P(E_A \cup E_B) \ge P(E_A) + P(E_B) - 1.$$

As  $P(E_A^c) = P(E_B^c) = 1/3$ , we can conclude that P(E) must be at least 1/3. To argue that P(E) cannot in general be smaller we describe a probability model in which  $P(E) \le 1/3$ . One way to do this is by specifying the marginal PDF of B given A as follows: B = A + 30 when 0 < A < 30 and B = 60 - A otherwise. It is easy to check that when A is Uniform(0, 45) then B is Uniform(0, 60), yet  $P(|B - A| < 30) = P(A \ge 30) = 1/3$ .

- 3. Raindrops hit the ground at a rate of 1 per second. An observatory has a raindrop sensing equipment. A signal is received by the computer with a maximum delay of 1 second after sensing a raindrop, with all delays equally likely. Find
  - (a) The joint PDF of the time T of the first raindrop and the time S of the signal reception.
  - (b) The marginal PDF of S.
  - (c) The conditional PDF of T given S.

## Solution:

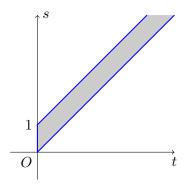
(a) S is a Uniform(T, T+1) random variable, where T is an Exponential(1) random variable. The joint PDF is  $f_{S,T}(s,t) = f_T(t)f_{S|T}(s|t)$ . We are given that

$$f_T(t) = \begin{cases} e^{-t} & \text{if } t \ge 0\\ 0 & \text{otherwise} \end{cases}, \qquad f_{S|T}(s|t) = \begin{cases} 1 & \text{if } t \le s \le t+1\\ 0 & \text{otherwise} \end{cases}$$
  
Therefore  $f_{S,T}(s,t) = f_T(t)f_{S|T}(s|t) = \begin{cases} e^{-t} & \text{if } t \ge 0, t \le s \le t+1\\ 0 & \text{otherwise} \end{cases}$ 

Alternative Solution: We can write S = T + D where the delay D is a Uniform(0,1) random variable that is independent of T. S and T take values s and t whenever T and D take values t and s - t, respectively, so by independence

$$f_{S,T}(s,t) = f_T(t)f_D(s-t) = \begin{cases} e^{-t} & \text{if } t \ge 0, \ 0 \le s-t \le 1\\ 0 & \text{otherwise} \end{cases}$$

(b) The marginal PDF of S is  $f_S(s) = \int_0^{+\infty} f_{S,T}(s,t) dt$ . Let R be the region greyed out in the following plot in which  $f_{S,T}(s,t)$  takes nonzero value  $e^{-t}$ .



Thus the integrand equals  $e^{-t}$  with bounds 0 to s when s is between 0 and 1, and s to s + 1 when s is larger than 1. This gives

$$f_S(s) = \int_0^s e^{-t} dt = 1 - e^{-s}, \quad \text{when } 0 \le s \le 1$$
  
$$f_S(s) = \int_{s-1}^s e^{-t} dt = e^{-s}(e-1), \quad \text{when } s > 1.$$

Alternative Solution: As T and D are independent we can calculate the PDF of S using the convolution formula:

If 
$$s \ge 1$$
:  
 $f_S(s) = \int_{-\infty}^{+\infty} f_D(x) f_T(s-x) dx = \int_0^1 e^{-(s-x)} dx = e^{-s}(e-1),$   
If  $0 \le s \le 1$ :  
 $f_S(s) = \int_{-\infty}^{+\infty} f_D(x) f_T(s-x) dx = \int_0^s e^{-(s-x)} dx = 1 - e^{-s}.$ 

(c) The conditional PDF is given by

$$f_{T|S}(t|s) = \frac{f_{S,T}(s,t)}{f_S(s)} = \begin{cases} \frac{e^{-t}}{1-e^{-s}} & \text{if } 0 \le s \le 1, \ 0 \le t \le s\\ \frac{e^{-t}}{e^{-s}(e-1)} & \text{if } s > 1, \ s-1 \le t \le s\\ 0 & \text{otherwise} \end{cases}$$

- 4. Here is a way to solve Buffon's needle problem without calculus. Recall that an  $\ell$  inch needle is dropped at random onto a lined sheet, where the lines are one inch apart.
  - (a) Let A be the number of lines that the needle hits. Let B be the number of times that a polygon of perimeter  $\ell$  hits a line. Show that E[A] = E[B]. (Hint: Use linearity of expectation.)
  - (b) Assume that  $\ell < \pi$ . Calculate the expected number of times that a circle of perimeter  $\ell$  hits a line.
  - (c) Assume that  $\ell < 1$ . Use part (a) and (b) to derive a formula for the probability that the needle hits a line. (**Hint:** The number of hits is a Bernoulli random variable.)

## Solution:

(a) Suppose the polygon has n edges of length  $a_1, a_2, \ldots, a_n$ . Break up the needle into segments of lengths  $a_1, a_2, \ldots, a_n$ . Let  $A_i$  and  $B_i$  be the number of lines hit by the *i*-th segment of the needle and the *i*-th edge of the polygon, respectively. Then

$$A = A_1 + \dots + A_n$$
 and  $B = B_1 + \dots + B_n$ .

By linearity of expectation

$$\mathbf{E}[A] = \mathbf{E}[A_1] + \dots + \mathbf{E}[A_n]$$
 and  $\mathbf{E}[B] = \mathbf{E}[B_1] + \dots + \mathbf{E}[B_n].$ 

Since the *i*-th edge of the polygon and the *i*-th segment of the needle are identical,  $E[A_i] = E[B_i]$ . It follows that E[A] = E[B].

(b) Let C be the number of times a circle intersects a line. We calculate the PMF of C. Let d be the line segment representing the diameter of the circle that is perpendicular to the lines on the sheet. Since  $\ell < \pi$ , the length of d is less than 1. The circle hits a line twice if d crosses a line, once if d touches one of the lines, and zero times if d does not intersect any of the lines. The probability that d crosses a line is exactly the length of d, namely  $\ell/\pi$ , and the probability that d touches a line is zero. Summarizing, the PMF of C is

$$\frac{c}{P(C=c)} \frac{0}{1-\ell/\pi} \frac{1}{0} \frac{2}{\ell/\pi}$$

Therefore  $E[C] = 2\ell/\pi$ .

(c) If we view the circle as a polygon with infinitely many sides, putting together part (a) and (b) we get that  $E[A] = E[C] = 2\ell/\pi$ . Since  $\ell < 1$ , the number of times the needle intersects a line is a 0/1 valued random variable, so E[A] = P(A = 1) = P(the needle hits a line). Therefore the probability the needle hits a line is exactly  $2\ell/\pi$ .